

RANDOM DISTRIBUTION FUNCTIONS¹

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1. Introduction. A *random distribution function* F is a measurable map from a probability space $(\Omega, \mathfrak{F}, Q)$ to the space Δ of distribution functions on the closed unit interval I , where Δ is endowed with its natural Borel σ -field, that is, the smallest σ -field containing the customary weak* topology. It determines a *prior* probability measure $P = QF^{-1}$ in the space Δ . Of course, F is essentially the same as the stochastic process $\{F_t, 0 \leq t \leq 1\}$ on $(\Omega, \mathfrak{F}, Q)$, where $F_t(\omega) = F(\omega)(t)$. Therefore, this note can be thought of as dealing with a certain class of random distribution functions, or a class of stochastic processes, or a class of prior probabilities.

Which class? Practically any *base probability* μ on the Borel subsets of the unit square S determines a random distribution function F and so a prior probability P_μ in Δ , which will be described somewhat informally in §2, by explaining how to select a value of F , i.e., a distribution function F , at random. §§3, 4 and 5 describe some properties of P_μ . Proofs will be given elsewhere. For ease of exposition, we assume that μ concentrates on, that is, assigns probability 1 to, the interior of S .

2. The construction. To select a value F of F at random, begin by selecting a point (x, y) from the interior of S according to μ . The horizontal and vertical lines through (x, y) divide S into four rectangles; consider the closed lower left rectangle L and the upper right one R . The unique (affine) transformation of the form $(u, v) \rightarrow (\alpha u + \beta, \gamma v + \delta)$, α and γ positive, which maps S onto L carries μ into a probability μ_L concentrated on L . The probability μ_R is defined in a similar way. Now select a point (x_L, y_L) at random from the interior of L according to μ_L , and a point (x_R, y_R) at random from the interior of R according to μ_R . As before, (x_L, y_L) determines four subrectangles of L , and (x_R, y_R) determines four subrectangles of R . Consider the lower left subrectangle LL in L , the upper right subrectangle RL in L , and the two analogous subrectangles LR and RR in R . The

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construction may be continued by selecting one point at random from each of these four rectangles, according to the appropriate affine image of μ , and so on. This procedure yields a nested decreasing sequence of closed sets, each being a finite union of closed rectangles: namely, $S, L \cup R, LL \cup RL \cup LR \cup RR$, and so on. The intersection of these closed sets is a nonempty closed set which, with probability 1, is the graph of a distribution function. This function is taken as the random value F of F .

With probability 1, F is continuous and strictly monotone, so that P_μ concentrates on the continuous strictly monotone distribution functions. Unless μ concentrates on the main diagonal, F is almost certainly purely singular, so that P_μ concentrates on the purely singular distribution functions.

Three interesting choices for μ [cited below as examples (1), (2), (3)] are: (1) the uniform distribution on the vertical line segment $x = 1/2, 0 \leq y \leq 1$; (2) the uniform distribution on the horizontal line segment $0 \leq x \leq 1, y = 1/2$; (3) the uniform distribution on S .

3. The average distribution function. A probability P in Δ determines as usual an *average distribution function* F_P according to the relation

$$F_P(x) = \int_{G \in \Delta} G(x) dP(G).$$

Consider the mapping T_μ of Δ into Δ defined by

$$\begin{aligned} (T_\mu F)x &= \int_0^1 \int_x^1 \beta F\left(\frac{x}{\alpha}\right) \mu(d\alpha, d\beta) \\ &+ \int_0^1 \int_0^x \left[\beta + (1 - \beta)F\left(\frac{x - \alpha}{1 - \alpha}\right) \right] \mu(d\alpha, d\beta). \end{aligned}$$

The average F_{P_μ} , or F_μ for short, satisfies the functional equation $T_\mu F = F$. Since T_μ is a uniformly strict contraction of the complete metric space Δ in the sup norm, T_μ has a unique fixed point, and if $G \in \Delta, (T_\mu)^n G \rightarrow F_\mu$ as $n \rightarrow \infty$.

In example (1) of §2, $F_\mu(x) = x, 0 \leq x \leq 1$; while in examples (2) and (3), $F_\mu(x) = 2\pi^{-1} \sin^{-1} x^{1/2}$. Surprisingly, therefore, the base probabilities μ of examples (1) and (2) yield different priors P_μ . It follows easily that the base probability of example (3) produces a third distinct prior.

To generalize example (1) slightly, if μ concentrates on the vertical line segment $x = r, 0 \leq y \leq 1$, and has mean (r, w) , the equation $T_\mu F = F$

takes the form

$$F(x) = wF\left(\frac{x}{r}\right), \quad 0 \leq x \leq r,$$

$$= w + (1-w)F\left(\frac{x-r}{1-r}\right), \quad r_1^* \leq x \leq 1$$

which, as shown in Chapter 6 of [2], has the unique solution

$$F(x) = Q_w[Q_r^{-1}(x)],$$

where the coin-tossing distribution function Q_w may be defined as follows. Let $\{\epsilon_j, 1 \leq j < \infty\}$ be independent random variables with the common distribution $P(\epsilon_j=0)=w$, $P(\epsilon_j=1)=1-w$; then Q_w is the distribution function of $\sum_{j=1}^{\infty} \epsilon_j 2^{-j}$. Since Q_r is strictly monotone on I , its inverse function Q_r^{-1} is also a distribution function on I .

The mapping T_μ is the usual operator on probabilities associated with a discrete time Markov process having I for state space and the following transition mechanism: when at $x \in I$, select (α, β) at random from S according to μ and move to αx with probability β , or to $x + \alpha(1-x)$ with probability $1-\beta$.

4. The uniqueness problem. In examples (1), (2), and (3), distinct base probabilities μ_1 and μ_2 lead to distinct priors P_{μ_1} and P_{μ_2} . On the other hand, if μ_1 and μ_2 are distinct but concentrated on the main diagonal of S , then P_{μ_1} and P_{μ_2} coincide, each assigning probability 1 to the distribution function λ , $\lambda(x)=x$, $0 \leq x \leq 1$. We have found no other exceptions to the conjecture that $\mu_1 \neq \mu_2$ implies $P_{\mu_1} \neq P_{\mu_2}$. This implication does hold when μ_1 and μ_2 are both concentrated on the same vertical line segment, say, $x=1/2$, $0 \leq y \leq 1$. As before, write $(1/2, w_i)$ for the mean of μ_i . Then $F_{\mu_i} = Q_{w_i}$, and for $w_1 \neq w_2$, it is well known from the strong law of large numbers that Q_{w_1} and Q_{w_2} are mutually singular. It follows easily that P_{μ_1} and P_{μ_2} are not only different but even mutually singular in the following strong sense. There exist two disjoint Borel subsets B_1 and B_2 of I (e.g., B_i may be taken as the set of binary irrationals whose binary expansion has w_i for limiting relative frequency of 0's), such that P_{μ_i} is concentrated on the collection C_i of distribution functions, where $F \in C_i$ if and only if the probability in I determined by F concentrates on B_i . Obviously, C_1 and C_2 are disjoint Borel subsets of Δ . If $w_1 = w_2$ but $\mu_1 \neq \mu_2$, such B_i do not exist; but P_{μ_1} and P_{μ_2} are still mutually singular in a fairly strong sense. Namely, there are disjoint Borel subsets C_1 and C_2 of Δ , such that P_{μ_i} concentrates on C_i , and having the further property: $F_i \in C_i$ implies that F_1 and F_2 are mutually singular.

5. **Consistency.** Let I^∞ be the space of sequences $\{x_j\}$, $x_j \in I$, $j=1, 2, \dots$, and let $\sigma(I^\infty)$ be its product σ -field. Let $\xi_n(s)$ be the n th coordinate of $s \in I^\infty$. If $\sigma(\Delta)$ denotes the Borel σ -field in Δ , a probability P on $(\Delta, \sigma(\Delta))$ determines a probability \bar{P} on $(\Delta \times I^\infty, \sigma(\Delta) \times \sigma(I^\infty))$ by the relation

$$\bar{P}\{A \times [s \mid \xi_j(s) \in A_j, 1 \leq j \leq n]\} = \int_{F \in A} \prod_{j=1}^n |F|(A_j) dP(F)$$

for $A \in \sigma(\Delta)$, A_j Borel in I ; where $|F|$ denotes the measure in I determined by F . Let P^* be a map from all n -tuples $\{x_j, 1 \leq j \leq n\}$ of elements of I to probabilities on $(\Delta, \sigma(\Delta))$, so that $P^*(\xi_1(s), \dots, \xi_n(s))$, as a function of s , is a version of the conditional distribution of F under \bar{P} , given $\{\xi_j, 1 \leq j \leq n\}$. In other words, $P^*(\xi_1(s), \dots, \xi_n(s))$ is "the" posterior distribution of F given $\{\xi_j(s), 1 \leq j \leq n\}$.

If $G \in \Delta$, let $|G|^\infty$ denote the unique probability on $(I^\infty, \sigma(I^\infty))$ under which the $\{\xi_n\}$ are independent with common distribution function G . Since Δ is compact metrizable, the space of probabilities on $(\Delta, \sigma(\Delta))$ has a weak* topology, as part of the dual of the space of continuous functions on Δ . Write Δ_0 for the set of all $G \in \Delta$ satisfying the following condition: for $|G|^\infty$ -almost all $s \in I^\infty$, $P^*(\xi_1(s), \dots, \xi_n(s))$ converges to point mass at G , in the weak* topology, as $n \rightarrow \infty$. Then $\Delta_0 \in \sigma(\Delta)$, and, as noted by Doob in [1], the forward martingale convergence theorem implies $P(\Delta_0) = 1$. But there is strong evidence that for most P , Δ_0 is only of the first category [3, §5]. Here is a result in the other direction. If the base probability μ concentrates on a vertical line segment $x=r$, $0 \leq y \leq 1$, and assigns positive mass to every nondegenerate subinterval of that segment, then there exists at least one choice of P_μ^* for which $\Delta_0 = \Delta$; which, in the usual terminology, says that Bayes' estimates constructed from P_μ are consistent.

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