

A Probabilistic Approach to a Differential-Difference Equation Arising in Analytic Number Theory

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Abstract. The differential-difference equation

$$\begin{aligned} tv'(t) + v(t - 1) &= 0, & t > 1, \\ v(t) &= 0, & t < 0, \\ v(t) &= \text{constant}, & 0 \leq t \leq 1, \end{aligned}$$

can be solved by the Monte-Carlo method, for the initial condition $v(t) = e^{-t}$, $0 \leq t \leq 1$, where the $v(t)$ represent the probability density of a random variable:

$$t = \lim_{n \rightarrow \infty} \sum_{i=1}^n \prod_{j=1}^i x_j,$$

where the x_j are independent and uniformly distributed on $(0, 1)$.

I. Introduction. The function $\psi(x, y)$ is equal to the number of integers less than or equal to x and free of prime factors greater than y . Chowla and Vijayaraghavan, Ramaswami, Buchstab and de Bruijn have shown that [1]:

$$\lim_{y \rightarrow \infty} \frac{\psi(y^t, y)}{y^t} = v(t),$$

where $v(t)$ is a function satisfying

$$\begin{aligned} tv'(t) + v(t - 1) &= 0, & t > 1, \\ v(t) &= 0, & t < 0, \\ v(t) &= 1, & 0 \leq t \leq 1. \end{aligned}$$

Many authors have studied the limits and asymptotic behaviour of this equation [2]; Norton gives an exhaustive bibliography [3]. Highly accurate numerical results were obtained by Dickman, Bellman, Van de Lune ([4], [5], [6]).

The differential-difference equation solution by the Monte-Carlo method does not claim to be as accurate as these previous calculations but only shows a probabilistic aspect of this equation.

II. Stochastic Model. Let u_n be the random variable: $u_n = x_1 + x_1x_2 + \dots + x_1x_2 \dots x_n$, where x_i are independent random variables uniformly distributed on $(0, 1)$.

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It may be deduced from the distribution of a product of x_i variables that if $n \rightarrow \infty$, u_n converges in probability to a limit.

LEMMA. Assume that $v(t)$ is a function continuous on $0 < t < \infty$ satisfying the following equation:

$$(1) \quad \begin{aligned} tv'(t) + v(t-1) &= 0, & t > 1, \\ v(t) &= 0, & t < 0, \\ v(t) &= C, & 0 \leq t \leq 1. \end{aligned}$$

This function is identical to $f(t)$: the probability density of a random variable:

$$t = \lim_{n \rightarrow \infty} \sum_{i=1}^n \prod_{j=1}^i x_j,$$

where x_i are independent random variables uniformly distributed on $(0, 1)$ if the constant C equals $e^{-\gamma}$, γ being the Euler constant.

*Proof.** Introduce

$$t_a = \sum_{i=1}^{\infty} \prod_{j=1}^i x_j \quad \text{and} \quad t_b = \sum_{i=2}^{\infty} \prod_{j=2}^i x_j;$$

t_a and t_b have the same probability distribution and $t_a = x_1(1 + t_b)$, t_b and x_1 are independent.

Let $F(t)$ be the distribution function of t_a :

$$F(t) = \text{Pr}[t_a \leq t];$$

of course, if $t < 0$, then $F(t) = 0$.

If $t > 0$, we have

$$\begin{aligned} F(t) &= \text{Pr}[t_a \leq t] = \text{Pr}[x_1(t_b + 1) \leq t] \\ &= \sum \text{Pr}[t_b + 1 \leq t/x] \text{Pr}[x \leq x_1 \leq x + dx] \\ &= \sum F\left(\frac{t}{x} - 1\right) \text{Pr}[x \leq x_1 \leq x + dx] = \int_0^1 F\left(\frac{t}{x} - 1\right) dx. \end{aligned}$$

Put $(t/x) - 1 = s$, then

$$F(t) = t \int_{t-1}^{\infty} \frac{F(s)}{(s+1)^2} ds.$$

If $0 \leq t \leq 1$, then

$$F(t) = t \int_0^{\infty} \frac{F(s) ds}{(s+1)^2} = C \cdot t,$$

where C is a constant. Hence, $f(t) = F'(t) = C$ for $0 \leq t \leq 1$.

If $t > 1$, by differentiating once, we get

$$f(t) = (F(t) - F(t-1))/t \geq 0;$$

by differentiating again, we find $tf'(t) = -f(t-1)$, $t > 1$.

* I am indebted to J. J. A. M. Brands for the correction of my initial proof.

TABLE I

n	$\Pr(u_n \leq 1)$ Explicit value	Monte-Carlo value (10^5 - runs)
2	0.69315	0.69416
3	0.61428	0.61622
4	0.58498	0.58350
5	0.57246	0.57356
6	0.56674	0.57016
7	0.56404	0.56303
8	0.56273	0.56290
9	0.56209	0.56381
10	0.56177	0.56030
∞	0.56146	

Let $h(s)$ be the Laplace transform of $f(t)$ [7]:

$$h(s) = (C_0/s) \exp\{-E_1(s)\},$$

where

$$E_1(s) = \int_s^\infty \frac{e^{-z}}{z} dz.$$

TABLE II

t	v(t) Explicit value	t $\Delta t = 0.1$	Monte-Carlo value (20 000 runs)	
			Rough value	Smooth value using REINSCH's (10) program
0	1	0 -0.1	0.96801	
		0.1-0.2	0.99206	
		0.2-0.3	1.03391	
		0.3-0.4	0.96890	
		0.4-0.5	1.01788	
		0.5-0.6	1.01432	
		0.6-0.7	1.02233	
		0.7-0.8	0.99206	
		0.8-0.9	1.01343	
		0.9-1.0	0.95733	
1.1	0.9046898202	1.0-1.1	0.95911	0.9624
1.2	0.8176784432	1.1-1.2	0.91547	0.8874
1.2	0.7376357355	1.2-1.3	0.81484	0.8132
1.4	0.6635277634	1.3-1.4	0.69016	0.7403
1.5	0.5945348919	1.4-1.5	0.58419	0.6693
1.6	0.5299963708	1.5-1.6	0.57974	0.6006
1.7	0.4693717489	1.6-1.7	0.50849	0.5346
1.8	0.4122133351	1.7-1.8	0.43992	0.4719
1.9	0.3581461138	1.8-1.9	0.37492	0.4128
		1.9-2.0	0.31169	0.3578
2.0	0.3068528194	2.0-2.1	0.29031	0.3070
2.1	0.2604057802	2.1-2.2	0.23510	0.2608
2.2	0.2203571379	2.2-2.3	0.17810	0.2193
2.3	0.1857994616	2.3-2.4	0.17098	0.1826
2.4	0.1559912639	2.4-2.5	0.16208	0.1506
2.5	0.1303195618	2.5-2.6	0.11132	0.1231
2.6	0.1082724430	2.6-2.7	0.09172	0.0999
2.7	0.08941856572	2.7-2.8	0.07748	0.0808
2.8	0.07339158076	2.8-2.9	0.05699	0.0653
2.9	0.05987811599	2.9-3.0	0.05076	0.0528
3.0	0.04860838829	3.0-3.1	0.05165	0.0425
3.1	0.03932296954	3.1-3.2	0.04186	0.0333
3.2	0.03170344451	3.2-3.3	0.02583	0.0250
3.3	0.02546472387	3.3-3.4	0.01514	0.0186
3.4	0.02037177906	3.4-3.5	0.01603	0.0145
3.5	0.01622959324	3.5-3.6	0.00980	0.0125
3.6	0.01287543418	3.6-3.7	0.01069	0.0121
3.7	0.01017283782	3.7-3.8	0.01514	0.0120
3.8	0.008006872188	3.8-3.9	0.00801	0.0092
3.9	0.006280373062	3.9-4.0	0.00534	0.0053
4.0	0.004910925648	4.0-4.1	0.00178	0.0018

Assuming that $f(t)$ is a probability $h(0) = \int_0^\infty f(t) dt = 1$, the constant C_0 equals $e^{-\gamma}$, where γ is the Euler constant.

Since $f(t) = C$ as $t = 0$, we obtain the boundary condition: $\lim_{s \rightarrow \infty} sh(s) = C = e^{-\gamma}$.

From $f(t) = C$ as $t = 1$, inverting Laplace transform, it may be deduced again that $f(1) = e^{-\gamma}$, so that

$$\begin{aligned}
 f(t) &= 0, & t < 0, \\
 f(t) &= e^{-\gamma}, & 0 \leq t \leq 1, \\
 f'(t) &= -f(t-1)/t, & t > 1.
 \end{aligned}$$

TABLE II

t	v(t) explicit value (*)	$\Delta t = 0.1$ t	Monte-Carlo value (3.10 ⁵ runs)
4.1	0.38285853 10 ⁻²	4.1-4.2	0.39 10 ⁻²
4.2	0.29754751 10 ⁻²	4.2-4.3	0.35 10 ⁻²
4.3	0.23050507 10 ⁻²	4.3-4.4	0.27 10 ⁻²
4.4	0.17795423 10 ⁻²	4.4-4.5	0.165 10 ⁻²
4.5	0.13701182 10 ⁻²	4.5-4.6	0.135 10 ⁻²
4.6	0.10514453 10 ⁻²	4.6-4.7	0.13 10 ⁻²
4.7	0.80455901 10 ⁻³	4.7-4.8	0.095 10 ⁻²
4.8	0.61395778 10 ⁻³	4.8-4.9	0.065 10 ⁻²
4.9	0.46728046 10 ⁻³	4.9-5.0	0.085 10 ⁻²
5.0	0.35472534 10 ⁻³	5.0-5.1	0.08 10 ⁻²
(Δ) 5.1	0.268580 10 ⁻³		
5.2	0.202822 10 ⁻³		
5.3	0.152768 10 ⁻³		
5.4	0.114775 10 ⁻³		
5.5	0.860192 10 ⁻⁴		
5.6	0.643153 10 ⁻⁴		
5.7	0.479771 10 ⁻⁴		
5.8	0.357089 10 ⁻⁴		
5.9	0.265188 10 ⁻⁴		
6.0	0.196503 10 ⁻⁴		

(*) Calculated by 4th order TAYLOR's expansion
 (Δ) Calculated by 5th order TAYLOR's expansion

DICKMAN result Monte-Carlo value

$$\int_0^{\infty} \frac{v(t)}{(1+t)^2} dt = 0.62433 = 0.6238$$

III. Numerical Calculations. For $t \leq 4$, the solution of Eq. (1) is obtained by explicit expression (see Appendix); for $t > 4$, it is impossible to express the solution by means of known functions. This explicit expression can thus be used for the well-known equation of the statistic theory of damage [8].

$$\begin{aligned} tu'(t) &= u(t - 1), & t > 1, \\ u(t) &= 0, & t < 0, \\ u(t) &= 1, & 0 \leq t \leq 1. \end{aligned}$$

For $t \leq 4$, the function $v(t)$ can be calculated with an accuracy depending solely on the polylogarithms which are used in its expression [9]. The random variable u_n is very easy to simulate by means of the pseudo-random numbers of Lehmer's method.

It can be seen in Section II that the u_n distributions achieve rapid convergence as n increases.

For the calculations, n is chosen so that we cannot discriminate between the distributions of u_n and u_{n-1} because the statistical fluctuations of the pseudo-random numbers are greater than the discrepancy between them.

IV. Results. Table I gives an illustration of Section II; notice that we get the Euler constant simulated by $-\text{Log}[\text{Pr}[u_n \leq 1]]$, $n \rightarrow \infty$.

Table II represents the calculation of the function $v(t)$ explicitly and by simulation. Results are smoothed by the spline method [10]. Polylogarithms can be calculated by means of Chebyshev's polynomial expansion [11], [12]; Kölbig gives an excellent algorithm for the dilogarithm's calculation [13].

V. Conclusion. The main purpose of this paper is to test the ability of the Monte-Carlo method to resolve differential-difference equations, and, using a classical example, to justify further studies in the field of the statistical theory of damage and neutron transport problems [14] which involve the same mathematical data.

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Appendix.

$v(t)$ explicit behaviour,

$$v(t) = 1 - \text{Log } t, \quad 1 \leq t \leq 2,$$

$$v(t) = 1 - \text{Log } t + \left[\frac{1}{2} \text{Log}^2 t + L_2(1/t) + L_2(-1)\right], \quad 2 \leq t \leq 3,$$

$$\begin{aligned} v(t) = & 1 - \text{Log } t + \left[\frac{1}{2} \text{Log}^2 t + L_2(1/t) + L_2(-1)\right] \\ & - \left\{ \frac{1}{4} \left[L_3\left(\frac{1}{4}\right) - L_3\left(\frac{1}{(t-1)^2}\right) \right] - \frac{1}{3} (\text{Log}^3(t-1) - \text{Log}^3 2) \right. \\ & + \frac{1}{2} (\text{Log}^2(t-1) \text{Log } t - \text{Log}^2 2 \text{Log } 3) + L_2\left(\frac{1}{t-1}\right) \text{Log } \frac{t}{t-1} \\ & - L_2\left(\frac{1}{2}\right) \text{Log} \left(\frac{3}{2}\right) - L_2\left(-\frac{1}{t-1}\right) \text{Log}(t-2) + L_2(-1) \text{Log } \frac{t}{3} \\ & \left. + \left\{ \underbrace{\left[\text{Log } \frac{1}{2} - \text{Log } \frac{t-2}{t-1} \right] + \left[\frac{1}{2} - \frac{1}{t-1} \right]}_{V_1} \right\} \right. \\ & - \frac{1}{2^2} \left[\underbrace{V_1 + \frac{1}{2} \left(\frac{1}{2^2} - \frac{1}{(t-1)^2} \right)}_{V_2} \right] \\ & \left. + \dots + \frac{(-1)^{p+1}}{(p+1)^2} \left[V_p + \frac{1}{p+1} \left(\frac{1}{2^{p+1}} - \frac{1}{(t-1)^{p+1}} \right) \right] + \dots \right\}, \quad 3 \leq t \leq 4. \end{aligned}$$

By means of Newton's method, the explicit expression permits easy calculation of the roots t_k

$$v(t_k) = \frac{1}{k}, \quad k = 4, 5, \dots, 203.$$

For example, the roots

$$t_4 = 2.1245966, \quad t_5 = 2.2571089$$

are used by Davenport and Erdős [15].

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