

ON ERGODIC SEQUENCES OF MEASURES

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ABSTRACT. Let Z be the group of integers and \bar{Z} its Bohr compactification. A sequence of probability measures $\{\mu_n, n = 1, 2, \dots\}$ defined on Z is said to be ergodic provided μ_n converges weakly to $\bar{\mu}$, the Haar measure on \bar{Z} . Let $A_n \subset Z, n = 1, 2, \dots$, and define μ_n by $\mu_n(B) = |A_n \cap B|/|A_n|$ where $|B|$ is the cardinality of B . Then it is easy to show that if $|A_n \cap A_n + k|/|A_n| \rightarrow 1$ for every $k \in Z$, then μ_n is ergodic. Let $0 \leq p_k \leq 1$. In this paper we construct (random) sequences $\{\mu_n\}$ which are ergodic, and such that $\lim(|A_n \cap A_n + k|/|A_n|) = p_k$, for every $k \in Z$.

1. Introduction. Let G be a locally compact abelian (l.c.a.) group. Let $\{\mu_n, n = 1, 2, \dots\}$ be a sequence of probability measures defined on the Borel sets of G . We shall say that such a sequence is ergodic provided μ_n converges weakly to $\bar{\mu}$, where $\bar{\mu}$ is Haar measure on the Bohr compactification of G . The reason for this terminology is that ergodicity of such a sequence is necessary and sufficient for the generalized mean ergodic theorem to hold: let $\{U_g, g \in G\}$ be any strongly continuous unitary representation G on a Hilbert space H . We say that the generalized mean ergodic theorem holds with respect to the sequence $\{\mu_n\}$ provided $\lim_{n \rightarrow \infty} \int_G U_g f \mu_n(dg) = P f$ strongly, for every $f \in H$, where P is the projection of H on the space $\{f \mid U_g f = f, g \in G\}$.

As mentioned above and shown in [1], the generalized mean ergodic theorem holds for a sequence $\{\mu_n\}$ if and only if the sequence is ergodic. Let $\{A_n\}$ be a sequence of Borel subsets of G , with $\mu(A_n) < \infty$, where μ is Haar measure on G . For $g \in G$ let $A_n g$ be A_n translated by g . Define the probability measures μ_n by

$$\mu_n(B) = \mu(A_n \cap B) / \mu(A_n),$$

for B a Borel set.

It is easy to show that if $\lim_{n \rightarrow \infty} (\mu(A_n \cap A_n g) / \mu(A_n)) = 1$ for every $g \in G$ then μ_n is an ergodic sequence. Such sequences will exist if and

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only if G is σ -compact. (See, e.g., [2].) That the above condition is not necessary was shown in [1].

Now let $G = Z$, the group of integers, and let $\{k_j, j = 1, 2, \dots\}$ be a sequence of positive integers. Let $A_n = \{k_1, \dots, k_n\}$ and define μ_n as above by $\mu_n(B) = |A_n \cap B|/|A_n|$, when $|A|$ is the cardinality of A . Clearly each μ_n may be thought of as a measure on \bar{Z} , the Bohr compactification of Z , and it follows from the Levy continuity theorem that μ_n is ergodic if and only if

$$(1) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{j=1}^n e^{2\pi i k_j \alpha} = 0 \quad \text{for } 0 < \alpha < 1.$$

From the criterion mentioned above it follows that such sequences will have this property provided $|A_n \cap A_n + k|/|A_n| \rightarrow 1$ for every integer k , where $A_n + k$ is A_n translated by k . In a personal communication to one of the authors, Niederreiter [3] proved that given p , with $0 \leq p \leq 1$, and a positive integer k , there exists (in fact, he constructed it) a sequence $\{k_j\}$ such that

$$\lim_{n \rightarrow \infty} \frac{|A_n \cap A_n + k|}{|A_n|} = p,$$

and such that the corresponding measures $\{\mu_n\}$ are ergodic. In this note we prove that given p_k with $0 \leq p_k \leq 1$, there exist (uncountably many) random sequences $\{k_j\}$ such that

- (a) the corresponding sequences of measures are ergodic, and
- (b) for every integer $k \neq 0$ we have

$$\lim_{n \rightarrow \infty} \frac{|A_n \cap A_n + k|}{|A_n|} = p_k.$$

In fact we show that (a) and (b) are true on a set of probability one.

As mentioned above, what must be shown is that (1) holds for all α . For the kind of sequences we construct it was shown by Robbins [4] that this holds for *each* α on a set of probability one, and the problem is to construct a single set of probability one such that the limiting relation holds for all α simultaneously. We also consider this problem for the group of reals.

2. The main result. Let X_1, X_2, \dots be a sequence of independent, identically distributed random variables with characteristic function

$$\phi(\alpha) = E e^{i\alpha X_1}.$$

Let $S_n = X_1 + \dots + X_n$ and

$$T_n(\alpha) = \frac{1}{n} \sum_{k=1}^n e^{i\alpha S_k}.$$

Theorem. *If $\phi(\alpha) \neq 1$ for $\beta \leq \alpha \leq \gamma$ and if $E(X_1)$ is finite, then $\sup_{\beta \leq \alpha \leq \gamma} |T_n(\alpha)| \rightarrow 0$ almost surely as $n \rightarrow \infty$.*

Proof. It is shown in the Lemma below that $\sup_{\beta \leq \alpha \leq \gamma} E|T_n(\alpha)|^4 = O(n^{-2})$. Now let k be integer valued and set

$$A_n = \left[\max_{\beta \leq k/n^{9/7} \leq \gamma} |T_n(k/n^{9/7})| \leq 1/n^{1/7} \right].$$

Then, using Boole's and Chebychev's inequalities and the Lemma,

$$\begin{aligned} P(A_n^c) &\leq \sum_{k: \beta \leq k/n^{9/7} \leq \gamma} P[|T_n(k/n^{9/7})|^2 \geq 1/n^{2/7}] \\ &= \sum_{k: \beta \leq k/n^{9/7} \leq \gamma} O(n^{-10/7}) = O(n^{-1/7}). \end{aligned}$$

Also

$$\sup_{\alpha} \left| \frac{d}{d\alpha} T_n(\alpha) \right| \leq \frac{1}{n} \sum_{k=1}^n |S_k|$$

and, letting $\mu = E|X_j|$,

$$E|S_k| \leq \sum_{j=1}^k E|X_j| = k\mu.$$

Let

$$B_n = \left[\sup_{\alpha} \left| \frac{d}{d\alpha} T_n(\alpha) \right| \leq n^{8/7} \right].$$

By Markov's inequality

$$P(B_n^c) \leq \frac{1}{n^{8/7}} E\left(\frac{1}{n} \sum_{k=1}^n |S_k|\right) \leq \frac{O(n^2)}{n^{15/7}} = O(n^{-1/7}).$$

On $A_n B_n$ we have

$$\sup_{\beta \leq \alpha \leq \gamma} |T_n(\alpha)| \leq \frac{1}{n^{1/7}} + \frac{n^{8/7}}{2n^{9/7}} = O(n^{-1/7}),$$

while $P(A_n B_n)^c = O(n^{-1/7})$ from the above estimates.

The proof is completed by using the Borel strong law of large numbers argument: for the subsequence $\{n^8\}$,

$$\sum_n P(A_{n^8} B_{n^8})^c = \sum O(n^{-8/7}) < \infty$$

and so $\sup_{\beta \leq \alpha \leq \gamma} |T_{n^8}(\alpha)| \rightarrow 0$ almost surely.

Now for any m , there exists an n such that $n^8 \leq m < (n + 1)^8$, and

$$\begin{aligned} |(|T_m(\alpha)| - |T_n(\alpha)|)| &\leq |T_m(\alpha) - T_n(\alpha)| \\ &= \left| \frac{1}{m} \sum_{k=n^8+1}^m e^{i\alpha S_k} - \left(\frac{1}{n^8} - \frac{1}{m}\right) \sum_{k=1}^{n^8} e^{i\alpha S_k} \right| \\ &\leq 2 \frac{m - n^8}{m} = O\left(\frac{1}{m^{1/8}}\right) \end{aligned}$$

uniformly in α . The Theorem follows.

In the Lemma below we have occasion to use the relation valid for any complex a and b with $|a| \leq 1$ and $a \neq 1$:

$$(*) \quad \left| \sum_{j=1}^{\nu} a^j b \right| = \left| \frac{a - a^{\nu+1}}{1 - a} b \right| \leq \frac{2|b|}{|1 - a|} = \frac{2|b|}{|1 - \bar{a}|}.$$

We also use without further comment the fact that, since the X_k 's are independent and identically distributed, for any $k > j$,

$$E e^{i\alpha(S_k - S_j)} = \phi(\alpha)^{k-j}$$

and the fact that $\phi(-\alpha) = \overline{\phi(\alpha)}$ and $|\phi(\alpha)| \leq 1$.

Lemma. *If $\phi(\alpha) \neq 1$ for $\beta \leq \alpha \leq \gamma$ then*

$$\sup_{\beta \leq \alpha \leq \gamma} E|T_n(\alpha)|^4 = O(1/n^2).$$

Proof. First observe that

$$\begin{aligned} E|T_n(\alpha)|^4 &= E(T_n(\alpha)^2 \overline{T_n(\alpha)^2}) \\ &= \frac{1}{n^4} \sum_{j,k,l,m=1}^n E e^{i\alpha(S_j + S_k - S_l - S_m)} \\ &= \frac{1}{n^4} \left(\sum_1 + \sum_2 + \sum_3 + \sum_4 \right) \end{aligned}$$

where

$$\sum_{\nu}(\alpha) = \sum_{j,k,l,m=1; |\{j,k,l,m\}|=\nu}^n E e^{i\alpha(S_j + S_k - S_l - S_m)}.$$

The modulus of each term in the sum for \sum_{ν} is at most one, so $|\sum_1| = O(n)$ and $|\sum_2| = O(n^2)$.

It is best to break \sum_3 into two sums, \sum_3' and \sum_3'' , where \sum_3' consists of those terms of \sum_3 which for $j = k$ or $l = m$ and \sum_3'' consists of those

terms of Σ_3 for which $|\{j, k\} \cap \{l, m\}| = 1$. Then

$$\begin{aligned} \left| \Sigma_3'(\alpha) \right| &= 2 \left| \operatorname{Re} \left(\sum_{j \neq k \neq l} E e^{i\alpha(2S_j - S_k - S_l)} \right) \right| \\ &= 4 \left| \operatorname{Re} \left(\sum_{j > k > l} \{ E e^{i\alpha(2S_j - S_k - S_l)} + E e^{i\alpha(2S_k - S_j - S_l)} + E e^{i\alpha(2S_l - S_j - S_k)} \} \right) \right| \\ &= 4 \left| \operatorname{Re} \left(\sum_{j > k > l} \{ \phi(2\alpha)^{j-k} \phi(\alpha)^{k-l} + \overline{\phi(\alpha)}^{j-k} \phi(\alpha)^{k-l} + \overline{\phi(\alpha)}^{j-k} \overline{\phi(2\alpha)}^{k-l} \} \right) \right| \\ &\leq \frac{8}{|1 - \phi(\alpha)|} \left(\left| \sum_{j > k} \phi(2\alpha)^{j-k} \right| + \left| \sum_{j > k} \overline{\phi(\alpha)}^{j-k} \right| + \left| \sum_{k > l} \overline{\phi(2\alpha)}^{k-l} \right| \right) \\ &= O(n^2)/|1 - \phi(\alpha)|. \end{aligned}$$

where the inequality uses (*).

Similarly, using (*) in the last equality,

$$\begin{aligned} \left| \Sigma_3''(\alpha) \right| &= 4 \left| \sum_{j \neq k \neq l} E e^{i\alpha(S_j - S_k)} \right| \\ &= 8(n-2) \left| \operatorname{Re} \left(\sum_{j > k} \phi(\alpha)^{j-k} \right) \right| = O(n^2)/|1 - \phi(\alpha)|. \end{aligned}$$

To estimate Σ_4 we will first write it in terms of ordered summation indices, $j > k > l > m$. There are then six types of terms according to the position of the two positive signs among $\pm S_j \pm S_k \pm S_l \pm S_m$. These can be coalesced into three types of terms by adding conjugates, to give

$$\begin{aligned} \left| \Sigma_4(\alpha) \right| &= 8 \left| \operatorname{Re} \left(\sum_{j > k > l > m} \{ E e^{i\alpha(S_j + S_k - S_l - S_m)} + E e^{i\alpha(S_j - S_k + S_l - S_m)} \right. \right. \\ &\quad \left. \left. + E e^{i\alpha(S_j - S_k - S_l + S_m)} \} \right) \right| \\ &= 8 \left| \operatorname{Re} \left(\sum_{j > k > l > m} \{ \phi(\alpha)^{j-k} \phi(2\alpha)^{k-l} \phi(\alpha)^{l-m} + \phi(\alpha)^{j-k+l-m} \right. \right. \\ &\quad \left. \left. + \phi(\alpha)^{j-k} \overline{\phi(\alpha)}^{l-m} \} \right) \right| \\ &\leq \frac{16}{|1 - \phi(\alpha)|} \left\{ \left| \sum_{k > l > m} \phi(2\alpha)^{k-l} \phi(\alpha)^{l-m} \right| + 2n \left| \sum_{l > m} \phi(\alpha)^{l-m} \right| \right\} \\ &= O(n^2)/|1 - \phi(\alpha)|^2. \end{aligned}$$

Here (*) is used twice, at the inequality and the last equality.

Combining these estimates,

$$\sup_{\beta \leq \alpha \leq \gamma} E|T_n(\alpha)|^4 = \frac{\alpha(n^2)}{n^4} \sup_{\beta \leq \alpha \leq \gamma} \frac{1}{|1 - \phi(\alpha)|^2}.$$

The hypothesis of the Lemma, together with the continuity of ϕ , imply that the supremum is finite, and the assertion is proved.

Let $\mathcal{L}(d)$, for $d > 0$, denote the lattice $\{0, \pm d, \pm 2d, \dots\}$. We will say that X is an $\mathcal{L}(d)$ lattice variable if $P[X \in \mathcal{L}(d)] = 1$ but there is no $d' > d$ such that $P[X \in \mathcal{L}(d')] = 1$. It is well known that X is an $\mathcal{L}(d)$ lattice variable if and only if $\phi(\alpha) \neq 1$ for $0 < \alpha < 2\pi/d$ and $\phi(2\pi/d) = 1$.

Corollary 1. *If X_1 is an $\mathcal{L}(1)$ lattice variable, then there exists a null set N not depending on α such that, except on N , $T_n(\alpha) \rightarrow 0$ as $n \rightarrow \infty$ for every $\alpha \neq 0 \pmod{2\pi}$.*

Proof. In this case the $T_n(\alpha)$ as well as $\phi(\alpha)$ are periodic of period 2π . By the Theorem, for any $k > 0$, $\sup_{1/k \leq \alpha \leq 2\pi - 1/k} |T_n(\alpha)| \rightarrow 0$ as $n \rightarrow \infty$ except on a null set N_k .

Evidently we can take $N = \bigcup_{k=1}^{\infty} N_k$ to be the set specified in the corollary.

We will say that X is a *nonlattice variable* if X is not an $\mathcal{L}(d)$ lattice variable for any $d > 0$ and if $P[X \neq 0] > 0$. This is the case if and only if $\phi(\alpha) \neq 1$ for any $\alpha \neq 0$. Then in the same way as for the first corollary we have the following result.

Corollary 2. *If X_1 is a nonlattice variable, then there exists a null set N not depending on α such that, except on N , $T_n(\alpha) \rightarrow 0$ as $n \rightarrow \infty$ for every $\alpha \neq 0$.*

Let

$$r_n(\delta) = \frac{|\{S_1, \dots, S_n\} \cap \{S_1 + \delta, \dots, S_n + \delta\}|}{|\{S_1, \dots, S_n\}|}.$$

We will consider only the case that $X_1 > 0$ so $|\{S_1, \dots, S_n\}| = n$.

Lattice case. Let X_1 be an $\mathcal{L}(1)$ lattice variable. Then

$$r_n(1) = \frac{1}{n} \sum_{k=2}^n \chi_{[X_k=1]} \xrightarrow{\text{a.s.}} P[X_1 = 1],$$

as $n \rightarrow \infty$.

Continuous case. Let X_1 be a nonlattice variable. Then

$$\begin{aligned} r_n(\delta) &= \frac{1}{n} \sum_{k=2}^n \chi_{[X_k=\delta \text{ or } X_k+X_{k+1}=\delta \text{ or } \dots \text{ or } X_k+\dots+X_n=\delta]} \\ &= \frac{1}{n} \sum_{j=0}^{n-2} \sum_{k=2}^{n-j} \chi_{A_{k,j}} \end{aligned}$$

where

$$A_{k,j} = [X_k + \dots + X_{k+j} = \delta].$$

Let $A_k = \bigcup_{j=0}^{\infty} A_{k,j}$ and let $p_j = P(A_{1,j})$ and $p = P(A_1)$. Then by the ergodic theorem

$$\frac{1}{n} \sum_{k=1}^n \chi_{A_{k,j}} \xrightarrow{\text{a.s.}} p_j, \quad \frac{1}{n} \sum_{k=1}^n \chi_{A_k} \xrightarrow{\text{a.s.}} p$$

as $n \rightarrow \infty$. But we have $r_n(\delta) \leq n^{-1} \sum_{k=2}^n \chi_{A_k}$ so $\limsup r_n(\delta) \leq p$. On the other hand, by Fatou's lemma,

$$\liminf r_n(\delta) \geq \sum_{j=0}^{\infty} \liminf \left(\frac{1}{n} \sum_{k=2}^{n-j} \chi_{A_{k,j}} \right) = \sum_{j=0}^{\infty} p_j = p$$

since $A_1 = \bigcup_j A_{1,j}$ and the sets in the union are disjoint. Thus $r_n(\delta) \xrightarrow{\text{a.s.}} p$, as $n \rightarrow \infty$. Clearly $p = 0$ except for at most a countable set of δ values.

3. Concluding remarks. Now let X_1, X_2, \dots be $\mathcal{L}(1)$ random variables. Then it follows from the results in §2 that if we define $k_j = X_1 + \dots + X_j$ the corresponding sequences of measures satisfy conditions (a) and (b) of §1 with probability one.

The results also apply to the case when $G = R$, the additive group of real numbers. For if X_1, X_2, \dots are positive nonlattice random variables we can apply Corollary 2 to show that $\lim_n T_n(\alpha) = 0$ for all $\alpha \neq 0$ with probability one. Thus if $k_j = X_1 + \dots + X_j$, and if $Y(t), t \geq 0$ is a stationary stochastic process we see that the mean ergodic theorem applies to the averages $n^{-1} \sum_{j=1}^n Y(k_j)$ provided the process is second order.

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