

A STRONG LAW OF LARGE NUMBERS FOR MARTINGALES¹

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ABSTRACT. We derive a moment inequality for the Skorohod representation theorem and apply it to obtain a strong law of large numbers for martingales.

1. Introduction. In the early sixties, Skorohod proved the following result; see Skorohod [5, p. 163], Hall and Heyde [4, p. 269].

THEOREM 1. *Let $\{S_n = \sum_{j=1}^n X_j, n \geq 1\}$ be a zero mean martingale. There exists a probability space supporting a standard Brownian motion $\{B(t)\}$ and a sequence of nonnegative random variables $\{\tau_n, n \geq 1\}$ such that*

- (i) $\{B(\sum_{j=1}^n \tau_j)\}$ has the same joint distribution as $\{S_n\}$,
- (ii) $E(\tau_n^r) \leq C_r E(|X_n|^{2r})$, C_r is a constant depending on r , $r \geq 1$.

In this paper, we shall derive an inequality more general than Theorem 1(ii) and apply this inequality to prove a strong law of large numbers for martingales.

2. A moment inequality. We shall adopt the following representation method which is slightly different from the one given by Freedman [3, p. 68]. Let X be a zero mean random variable with distribution function F , let $f(x) = F^{-1}(x) = \inf\{t: F(t) \geq x\}$, $0 \leq x \leq 1$. Let $g(x)$ be a monotone solution of the function equation $\int_x^{g(x)} f(t) dt = 0$. Now construct a probability space supporting a standard Brownian motion process $\{B(t)\}$ and a random variable U which is independent of $\{B(t)\}$ and uniformly distributed over $[0, 1]$. Conditioning on $U = x$, define $\tau(x) = \inf\{t: B(t) = f(x) \text{ or } f(g(x))\}$. Then the distribution of B_τ is F and $E(\tau) = E(X^2)$. To see this, let $G(u, v)$ be the distribution which assign probability $|u|/(|u|+|v|)$ at v and probability $|v|/(|u|+|v|)$ at u , $uv < 0$. Note that conditioning on $U = x$, B_τ has distribution $G(f(x), f(g(x)))$. Therefore, if ψ is a bounded measurable function, then

$$E\psi(B_\tau) = \int_0^1 [\psi(f(x))|f(g(x))| + \psi(f(g(x)))|f(x)|] (|f(g(x))| + |f(x)|)^{-1} dx.$$

Using the facts that

$$g(g(x)) = x, \quad f(x) = f(g(x))g'(x)$$

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we see

$$\begin{aligned} & \int_0^1 \psi(f(g(x)))|f(x)|(|f(x)| + |f(g(x))|)^{-1} dx \\ &= \int_0^1 \psi(f(x))|f(x)|(|f(x)| + |f(g(x))|)^{-1} dx. \end{aligned}$$

Hence

$$E\psi(B_\tau) = \int_0^1 \psi(f(x)) dx = \int_{-\infty}^{\infty} \psi(x) dF(x).$$

That is, B_τ has distribution F . By the same reason,

$$E(\tau) = \int_0^1 |f(x)f(g(x))| dx = \int_0^1 f^2(x) dx = \int_{-\infty}^{\infty} x^2 dF(x) = E(X^2).$$

Moreover, we shall show

THEOREM 2. *Let $\phi(x)$, $x \geq 0$, be a nonnegative function such that for some a , b , $\frac{1}{2} < a < b < \infty$, $\phi(x)/x^a$ is nondecreasing and $\phi(x)/x^b$ is nonincreasing. Then $E\phi(\tau) \leq CE\phi(X^2)$, where the constant C depends only on a and b .*

PROOF. Following Breiman [1], let $T = \inf\{t: B(t) = u \text{ or } v\}$, $-u \leq 0 \leq v$. Then

$$P(T \geq t) = \frac{4}{\pi} \sum_{n=1}^{\infty} (2n + 1)^{-1} \exp\left[-\frac{\lambda_n^2 t}{2r^2}\right] \sin\left(\frac{\lambda_n u}{r}\right),$$

where $\lambda_n = (2n + 1)$, $r = u + v$. By induction, $\sin(\lambda_n u/r) \leq 4\pi^2(2n + 1)uv/r^2$. Hence,

$$\begin{aligned} E\phi(T) &= \int_0^{\infty} P(T \geq t) d\phi(t) \\ &\leq C_1 \frac{uv}{r^2} \int_0^{\infty} \left(\sum_{n=1}^{\infty} \exp\left[-\frac{\lambda_n^2 t}{2r^2}\right] \right) d\phi(t). \end{aligned}$$

From the inequality

$$\begin{aligned} \sum_{n=0}^{\infty} e^{-\lambda_n^2 s} &\leq e^{-8s} \sum_{n=0}^{\infty} e^{-4n^2 \pi^2 s} \\ &\leq e^{-8s} \left(1 + \int_0^{\infty} e^{-4\pi^2 s x^2} dx \right) = \left(1 + \frac{C_2}{\sqrt{s}} \right) e^{-8s} \end{aligned}$$

follows

$$E\phi(T) \leq C_3 \left[\frac{uv}{r^2} \int_0^{\infty} \exp\left[-\frac{4t}{r^2}\right] d\phi(t) + \frac{uv}{r} \int_0^{\infty} \exp\left[-\frac{4t}{r^2}\right] \frac{d\phi(t)}{\sqrt{t}} \right].$$

Write

$$\int_0^{\infty} e^{-st} d\phi(t) = \int_0^{1/s} e^{-st} d\phi(t) + \int_{1/s}^{\infty} e^{-st} d\phi(t),$$

where

$$\begin{aligned} \int_0^{1/s} e^{-st} d\phi(t) &= e^{-1} \phi\left(\frac{1}{s}\right) + s \int_0^{1/s} e^{-st} \phi(t) dt \\ &\leq e^{-1} \phi(s^{-1}) + s^{a+1} \phi(s^{-1}) \int_0^{1/s} t^a e^{-st} dt \\ &\leq C_4 \phi(s^{-1}) \end{aligned}$$

and

$$\begin{aligned} \int_{1/s}^{\infty} e^{-st} d\phi(t) &= -e^{-1}\phi(s^{-1}) + s \int_{1/s}^{\infty} e^{-st}\phi(t) dt \\ &\leq -e^{-1}\phi(s^{-1}) + s^{b+1}\phi(s^{-1}) \int_{1/s}^{\infty} t^b e^{-st} dt \\ &\leq C_5\phi(s^{-1}). \end{aligned}$$

Similarly

$$\begin{aligned} \int_0^{\infty} t^{-1/2} e^{-st} d\phi(t) &= \int_0^{1/s} t^{-1/2} e^{-st} d\phi(t) + \int_{1/s}^{\infty} t^{-1/2} e^{-st} d\phi(t) \\ &\leq C_6\sqrt{s}\phi(s^{-1}), \end{aligned}$$

where $C_1, C_2, C_3, C_4, C_5, C_6$ are constants and C_4, C_5 depend on a , C_6 depends on b . Therefore, one concludes

$$E\phi(T) \leq C \frac{uv}{r^2} \phi\left(\frac{r^2}{4}\right) \leq C \frac{uv}{r^2} \phi\left(\frac{u^2 + v^2}{2}\right),$$

where C is a constant depending on a and b .

Note

$$\phi\left(\frac{1}{2}(u^2 + v^2)\right) \leq \max\{\phi(u^2), \phi(v^2)\} \leq \phi(u^2) + \phi(v^2).$$

Hence,

$$E\phi(T) \leq C \frac{uv}{r^2} (\phi(u^2) + \phi(v^2)).$$

This implies

$$\begin{aligned} E\phi(\tau) &\leq C \int_0^1 \frac{|f(x)f(g(x))|}{(|f(x)| + |f(g(x))|)^2} [\phi(f^2(x)) + \phi(f^2(g(x)))] dx \\ &= C \int_0^1 |f(x)|\phi(f^2(x))(|f(x)| + |f(g(x))|)^{-1} dx \\ &\leq C \int_0^1 \phi(f^2(x)) dx = CE\phi(X^2). \quad \text{Q.E.D.} \end{aligned}$$

3. A strong law. Using the result in Theorem 2, we shall prove a strong law of large numbers for martingales. First, observe a simple fact: if two sequences of random variables $\{X_n\}$ and $\{Y_n\}$ have the joint distribution, C is a constant, then $P(|X_n - C| \geq \epsilon, \text{i.o.}) = P(|Y_n - C| \geq \epsilon, \text{i.o.}) \forall \epsilon > 0$; therefore, $X_n \rightarrow C$, a.s. if and only if $Y_n \rightarrow C$, a.s.

THEOREM 3. Let $S_n = \{\sum_{j=1}^n X_j, n \geq 1\}$ be a zero mean martingale. Let $\phi(x), x \geq 0$, satisfy the assumptions in Theorem 2 and $\phi(x + y) \leq \phi(x) + \phi(y)$, for all $x, y \geq 0$. If for some sequence of positive numbers $\{a_n\}$, we have $a_n \uparrow \infty$ and $\sum_{n=1}^{\infty} E\phi(X_n^2)/\phi(a_n) > \infty$, then $S_n/a_n^p \rightarrow 0$ a.s. for every $p > \frac{1}{2}$.

PROOF. Let $\{\mathcal{B}(\sum_{j=1}^n \tau_j)\}$ represent $\{S_n\}$ as in Theorem 1. By Theorem 2 and the assumption that $\sum_{n=1}^{\infty} E\phi(X_n^2)/\phi(a_n) < \infty$, we conclude

$$\sum_{n=1}^{\infty} \phi(\tau_n)/\phi(a_n) < \infty \quad \text{a.s.}$$

Applying Kronecker's Lemma we get

$$\frac{1}{\phi(a_n)} \left(\sum_{j=1}^n \phi(\tau_j) \right) \rightarrow 0 \quad \text{a.s.}$$

By the assumption $\phi(x + y) \leq \phi(x) + \phi(y)$, we see

$$\phi \left(\sum_{j=1}^n \tau_j \right) / \phi(a_n) \rightarrow 0 \quad \text{a.s.}$$

Because $\phi(x)/x^a$ is nondecreasing and $\phi(x)/x^b$ is nonincreasing, we have

$$\phi \left(\sum_{j=1}^n \tau_j \right) / \phi(a_n) \geq \min \left\{ \left[\frac{1}{a_n} \left(\sum_{j=1}^n \tau_j \right) \right]^a, \left[\frac{1}{a_n} \left(\sum_{j=1}^n \tau_j \right) \right]^b \right\}.$$

Therefore,

$$\frac{1}{a_n} \left(\sum_{j=1}^n \tau_j \right) \rightarrow 0 \quad \text{a.s.}$$

Without loss of generality, we may assume $S_n \rightarrow \infty$ a.s. which implies $\sum_{j=1}^n \tau_j \rightarrow \infty$ a.s. But the law of the iterated logarithm for Brownian paths tells us

$$B \left(\sum_{j=1}^n \tau_j \right) / \left(\sum_{j=1}^n \tau_j \right)^p \rightarrow 0 \quad \text{a.s. } \forall p > \frac{1}{2}.$$

Hence,

$$S_n/a_n^p \rightarrow 0 \quad \text{a.s. } \forall p > \frac{1}{2}. \quad \text{Q.E.D.}$$

COROLLARY 4. *Let $\{S_n = \sum_{j=1}^n X_j\}$ be a zero mean martingale. If $E(X_n^2) < \infty \forall n$, and $s_n^2 = \sum_{j=1}^n E(X_j^2) \rightarrow \infty$, then $S_n/s_n^p \rightarrow 0$, a.s. $\forall p > 1$.*

PROOF. Choose $\phi(x) = x$, $a_n = s_n^2(\log s_n^2)^{1+\varepsilon} \uparrow \infty$, $\varepsilon > 0$. Then

$$\sum_{n=1}^{\infty} \frac{E\phi(X_n^2)}{\phi(a_n)} = \sum_{n=1}^{\infty} \frac{E(X_n^2)}{s_n^2(\log s_n^2)^{1+\varepsilon}} < \infty$$

by Dini's theorem, and consequently

$$S_n/s_n^{2p}(\log s_n^2)^{p+p\varepsilon} \rightarrow 0 \quad \text{a.s. for every } p > \frac{1}{2}$$

which implies

$$S_n/s_n^p \rightarrow 0 \quad \text{a.s. for every } p > 1. \quad \text{Q.E.D.}$$

COROLLARY 5. *Let $\{S_n = \sum_{j=1}^n X_j\}$ be a zero mean martingale. If*

$$\sum_{n=1}^{\infty} E|X_n|^r/n^r < \infty, \quad 1 < r \leq 2,$$

then $S_n/n^p \rightarrow 0$ a.s. $\forall p > 1$.

PROOF. Choose $\phi(x) = x^{r/2}$, $a_n = n^2 \uparrow \infty$. Apparently, $\phi(x)$ satisfies the required conditions in Theorem 3 and

$$\sum_{n=1}^{\infty} \frac{E\phi(X_n^2)}{\phi(a_n)} = \sum_{n=1}^{\infty} E|X_n|^r/n^r < \infty$$

by assumption. Hence $S_n/n^p \rightarrow 0$ a.s. for every $p > 1$. Q.E.D.

REMARK 6. Chow [2] has shown that if $\sum_{n=1}^{\infty} E(X_n^2)/n^2 < \infty$, then $S_n/n \rightarrow 0$ a.s.

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