

CATALYTIC BRANCHING RANDOM WALK AND QUEUEING SYSTEMS WITH RANDOM NUMBER OF INDEPENDENT SERVERS

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ABSTRACT. A continuous time branching random walk on the lattice \mathbf{Z} in which particles may produce children only at the origin is considered. Assuming that the underlying random walk is symmetric and the offspring reproduction law is critical, we find the asymptotic behavior of the survival probability of the process at time t as $t \rightarrow \infty$ and the probability that the number of particles at the origin at time t is positive. We also prove a Yaglom type conditional limit theorem for the total number of particles existing at time t . A relation between the model considered and a queueing system with a random number of independently operating servers is discussed.

1. INTRODUCTION

We consider the following modification of a standard branching random walk on \mathbf{Z} . Let a population of particles evolves as follows. The population is initiated at time $t = 0$ by a single particle. Being outside of the origin the particle performs a continuous time random walk on \mathbf{Z} with the infinitesimal transition matrix $A = |a(x, y)|_{x, y \in \mathbf{Z}}$, $a(0, 0) < 0$, until the moment when it hits the origin. At the origin it spends an exponentially distributed time with parameter 1 and then either jumps to a point $y \neq 0$ with probability $-(1 - \alpha)a(0, y)a^{-1}(0, 0) \stackrel{\text{def}}{=} (1 - \alpha)\pi_y$, or dies with probability α producing just before its death a random number of children ξ in accordance with the offspring generating function

$$(1) \quad f(s) \stackrel{\text{def}}{=} \mathbf{E} s^\xi = \sum_{k=0}^{\infty} f_k s^k.$$

The new particles behave independently and stochastically in the same way as the parent particle.

Now we describe our basic hypotheses on the characteristics of the process:

Hypothesis (I). The underlying random walk is symmetric, homogeneous, and irreducible on \mathbf{Z} : $a(x, y) = a(y, x)$, $a(x, y) = a(0, y - x) \stackrel{\text{def}}{=} a(y - x)$ with $a(x) \geq 0$, $x \neq 0$, $a(0) < 0$ and $\sum_{x \in \mathbf{Z}} a(x) = 0$; in addition

$$(2) \quad b^2 \stackrel{\text{def}}{=} \sum_{x \in \mathbf{Z}} x^2 a(x) < \infty.$$

Hypothesis (II). The offspring generating function $f(s)$ is critical ($f'(1) = 1$) and $\sigma^2 \stackrel{\text{def}}{=} f''(1) \in (0, \infty)$.

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Let $\zeta(t)$ denote the number of particles in the process at time t at the origin, $\mu(t)$ the number of particles in the process at time t outside the origin, and let

$$\eta(t) = \zeta(t) + \mu(t)$$

be the total number of particles in the process at time t . The goal of the present paper is to study the asymptotic behavior of the probabilities $Q(t) = \mathbb{P}(\mu(t) > 0)$ and $q(t) = \mathbb{P}(\zeta(t) > 0)$ and to establish a Yaglom type conditional limit theorem for $\eta(t)$.

Let

$$K \stackrel{\text{def}}{=} \frac{2^{3/4}}{\sigma\pi^{1/4}} \sqrt{\frac{b(1-\alpha)}{\alpha}}.$$

Our main results are contained in the following two theorems.

Theorem 1. *Let Hypotheses (I) and (II) hold. Then*

$$(3) \quad Q(t) = \mathbb{P}(\mu(t) > 0) \sim Kt^{-1/4}, \quad t \rightarrow \infty,$$

and for any $s \in [0, 1]$,

$$\lim_{t \rightarrow \infty} \mathbb{E} \left[s^{\mu(t)} \mid \eta(t) > 0 \right] = 1 - \sqrt{1-s}.$$

Theorem 2. *Under the conditions of Theorem 1,*

$$q(t) = \mathbb{P}(\zeta(t) > 0) \sim \frac{K^2\pi}{\sqrt{t \ln t}}, \quad t \rightarrow \infty.$$

The following assertion is an easy consequence of the previous theorems.

Corollary 3. *Under the conditions of Theorem 1,*

$$\mathbb{P}(\eta(t) > 0) \sim Kt^{-1/4}, \quad t \rightarrow \infty,$$

and for any $s \in [0, 1]$

$$\lim_{t \rightarrow \infty} \mathbb{E} \left[s^{\eta(t)} \mid \eta(t) > 0 \right] = 1 - \sqrt{1-s}.$$

This model was first considered in [1, 2, 4, 5] in a more general setting, namely, for the case where particles walk on the d -dimensional lattice \mathbf{Z}^d . In these papers, the authors deduce basic equations for the probability generating functions of the random variables $\eta(x; t)$, the total number of particles in the population at time t , and $\eta(x, y; t)$, the number of particles at point $y \in \mathbf{Z}^d$ at time t , given that the population is initiated at time zero by a single particle located at point $x \in \mathbf{Z}^d$. Besides, those papers contain asymptotic representations for $\mathbb{E} \eta^k(x; t)$ and $\mathbb{E} \eta^k(x, y; t)$, $k = 1, 2, \dots$, in all dimensions. A superprocess version of this model, called the catalytic superprocess, is considered in [6, 8, 9].

These results can be applied in a natural way to the following queueing model.

Denote by $M(\lambda)$ a random variable having the exponential distribution with parameter λ .

Our model of the queueing system corresponds to a catalytic branching random walk on the lattice of integers with symmetric ± 1 random walk (or on the half-line $y \in \mathbf{Z}^+$).

This system consists of a countable number of servers. Each server can be in one of the four states: busy, idle, broken or locked.

I. For each busy server:

- 1) the arrival process is Poisson with $\lambda = 1/2$ and is independent of the arrival processes for the remaining servers (busy or idle);
- 2) service times are $M(1/2)$;
- 3) all customers arriving at a particular busy server stay in the queue to this server.

II. Each idle server:

- 1) is waiting for the first customer for a time η , which is an $M(\alpha)$ random variable;
- 2) the arrival process of customers is Poisson with parameter $1 - \alpha$;
- 3) if the arrival time τ_1 of the first customer does not exceed η , the server starts to serve the customer with service time $M(1/2)$ and perform the service of this and subsequent customers according to the service discipline mentioned until it is idle again;
- 4) if $\tau_1 > \eta$, then just after the deadline moment for its waiting time the server becomes broken and simultaneously ξ new servers are unlocked, that is, they are in the operational state and idle from this moment on (note that $\xi = 0$ with probability f_0).

III. At the initial moment $t = 0$ we have one idle server and the remaining servers are locked.

It is not difficult to check that if, in terms of the catalytic branching random walk described above, the negative values on the lattice \mathbf{Z} are identified with the positive ones and each particle on \mathbf{Z} is associated with a server, then the total amount of particles is the number of servers in the system, while the coordinates of the particles on \mathbf{Z} (or, to be more precise, on \mathbf{Z}^+) are the queue lengths at the corresponding servers.

Thus, Theorems 1 and 2 and Corollary 3 deal with the total number of busy servers and the total number of servers in the operational state at the moment t and with probability that there are idle servers in the system at this moment.

2. AUXILIARY RESULTS

We temporarily forget that our random walk has a point of catalysis and consider an ordinary random walk on \mathbf{Z} satisfying Hypothesis (I).

Let $p(t; x, y)$ be the transition probability of the random walk, that is, the probability that starting at time $t = 0$ at point x the particle is located at point y at time t . Under the conditions of Hypothesis (I) $p(t; x, y) = p(t; 0, y - x) \stackrel{\text{def}}{=} p(t; y - x)$ and the backward Kolmogorov equations for $p(t; x, y)$ are as follows:

$$(4) \quad \frac{\partial p(t; x)}{\partial t} = \sum_{y \in \mathbf{Z}} a(y) p(t; y - x), \quad p(0; x) = \delta_0(x),$$

where $\delta_0(x) = 1$ if $x = 0$, and $\delta_0(x) = 0$ otherwise.

Lemma 4. *Let Hypothesis (I) be valid. Then the function*

$$\phi(\theta) \stackrel{\text{def}}{=} \sum_{x \in \mathbf{Z}} a(x) e^{i\theta x}, \quad \theta \in [-\pi, \pi],$$

is twice differentiable, real-valued, and nonpositive. It has a unique maximum on $[-\pi, \pi]$ at the point $\theta_0 = 0$, and there exists a constant $\gamma > 0$ such that $\phi(\theta) \leq -\gamma\theta^2$ in a neighborhood of zero.

Proof. Since $a(0) = -\sum_{x \neq 0} a(x)$ and the random walk is symmetric, we have

$$\phi(\theta) = a(0) + \frac{1}{2} \sum_{x \neq 0} a(x) (e^{i\theta x} + e^{-i\theta x}) = \sum_{x \neq 0} a(x) (\cos \theta x - 1) \leq 0.$$

The remaining part of the lemma follows from (2) and the irreducibility of the random walk. □

For brevity we use the notation $p(t) \stackrel{\text{def}}{=} p(t; 0, 0)$.

Lemma 5. *Let Hypothesis (I) be valid. Then $p(t)$ and $-p'(t)$ are monotone decreasing in t and admit the representations*

$$(5) \quad p(t) \sim \gamma_1 t^{-1/2}, \quad t \rightarrow \infty,$$

$$(6) \quad -p'(t) \sim \frac{\gamma_1}{2} t^{-3/2}, \quad t \rightarrow \infty,$$

where

$$(7) \quad \gamma_1 \stackrel{\text{def}}{=} \frac{1}{b\sqrt{2\pi}}.$$

Proof. Let $\tilde{p}(t; \theta) \stackrel{\text{def}}{=} \sum_{x \in \mathbf{Z}} p(t; x) e^{i\theta x}$. Then (4) can be rewritten as follows:

$$\frac{\partial \tilde{p}(t; \theta)}{\partial t} = \phi(\theta) \tilde{p}(t; \theta), \quad \tilde{p}(0; \theta) = 1.$$

Hence $\tilde{p}(t; \theta) = e^{t\phi(\theta)}$ and therefore

$$p(t, x) = \frac{1}{2\pi} \int_{-\pi}^{\pi} e^{t\phi(\theta) - ix\theta} d\theta.$$

Thus

$$p(t) = \frac{1}{2\pi} \int_{-\pi}^{\pi} e^{t\phi(\theta)} d\theta, \quad p'(t) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \phi(\theta) e^{t\phi(\theta)} d\theta,$$

$$p''(t) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \phi^2(\theta) e^{t\phi(\theta)} d\theta.$$

Since $\phi(\theta) \leq 0$, $\theta \in [-\pi, \pi]$, the desired monotonicity of $p(t)$ and $-p'(t)$ follows. To complete the proof it remains to note that in view of Lemma 4 we can apply the Laplace method to the integral

$$p(t) = \frac{1}{2\pi} \int_{-\pi}^{\pi} e^{t\phi(\theta)} d\theta$$

to get (5) and (6) (see, for instance, [10], Chapter 3, Section 7, Theorem 7.1). \square

Now we assume that the random walk starts at the origin at time $t = 0$. Let τ_1 be the time spent by the particle at the origin until it leaves the origin, and let τ_2 be the time spent by this particle outside the origin until its first return to the origin. Set $G_1(t) \stackrel{\text{def}}{=} \mathbf{P}(\tau_1 \leq t) = 1 - e^{-t}$ and $G_2(t) \stackrel{\text{def}}{=} \mathbf{P}(\tau_2 \leq t)$.

In the sequel the following convention is used: for a function $g(t)$, $t \geq 0$, we set

$$\hat{g}(\lambda) \stackrel{\text{def}}{=} \int_0^{\infty} e^{-\lambda t} g(t) dt;$$

if, in addition, $g(t)$ is nonnegative and monotone increasing, then

$$\check{g}(\lambda) \stackrel{\text{def}}{=} \int_0^{\infty} e^{-\lambda t} dg(t).$$

Lemma 6. *Let Hypothesis (I) be valid. Then*

$$1 - G_2(t) \sim \frac{b\sqrt{2}}{\sqrt{\pi t}}, \quad t \rightarrow \infty.$$

Proof. It is not difficult to check that

$$p(t) = 1 - G_1(t) + \int_0^t p(t-u) d(G_1 * G_2(u)),$$

where $*$ denotes the convolution. Using this relation we obtain

$$\hat{p}(\lambda) = \widehat{1 - G_1}(\lambda) + \hat{p}(\lambda) \check{G}_1(\lambda) \check{G}_2(\lambda)$$

or

$$(8) \quad \hat{p}(\lambda) (1 - \check{G}_1(\lambda)\check{G}_2(\lambda)) = \widehat{1 - G_1}(\lambda).$$

Recalling that $\check{G}_1(\lambda) = (1 + \lambda)^{-1}$ we get

$$\hat{p}(\lambda) \left(1 - \frac{1}{1 + \lambda} \check{G}_2(\lambda)\right) = \frac{1}{1 + \lambda} \quad \text{and} \quad (\widehat{1 - G_2})(\lambda) = \frac{1 - \lambda\hat{p}(\lambda)}{\lambda\hat{p}(\lambda)}.$$

Since

$$(9) \quad \hat{p}(\lambda) \sim \frac{\gamma_1 \sqrt{\pi}}{\sqrt{\lambda}}, \quad \lambda \rightarrow +0,$$

in view of Lemma 5 and a Tauberian theorem (Theorem 2 in [7], Chapter XIII, §5), we have

$$(\widehat{1 - G_2})(\lambda) \sim \frac{1}{\lambda\hat{p}(\lambda)} \sim \frac{1}{\gamma_1 \sqrt{\pi} \sqrt{\lambda}}, \quad \lambda \rightarrow +0.$$

Applying the Tauberian theorem once again we get

$$1 - G_2(t) \sim \frac{1}{\gamma_1 \pi \sqrt{t}} = \frac{b\sqrt{2}}{\sqrt{\pi t}}, \quad t \rightarrow \infty$$

(recall the definition of γ_1 in (7)). □

Since $p(0) = 1$, we can regard $G(t) \stackrel{\text{def}}{=} 1 - p(t)$ as a distribution function. For $\alpha \in (0, 1)$ let $U_\alpha(t) = \sum_{k=0}^\infty \alpha^k G^{*k}(t)$ and let $P(t)$ be the solution of the renewal equation

$$(10) \quad P(t) = 1 - G_1(t) + \int_0^t P(t - u) dG_3(u)$$

with

$$(11) \quad G_3(t) \stackrel{\text{def}}{=} \alpha G_1(t) + (1 - \alpha) G_1 * G_2(t).$$

It is clear that

$$P(t) = (1 - G_1(\cdot)) * U(t)$$

where

$$(12) \quad U(t) \stackrel{\text{def}}{=} \sum_{k=0}^\infty G_3^{*k}(t).$$

Lemma 7. *We have*

$$(13) \quad P(t) = \frac{1}{\alpha} - \frac{1 - \alpha}{\alpha} U_\alpha(t).$$

Proof. Applying the Laplace transform to both sides of (10) we get

$$\hat{P}(\lambda) = \frac{(\widehat{1 - G_1})(\lambda)}{\alpha (1 - \check{G}_1(\lambda)) + (1 - \alpha) (1 - \check{G}_1(\lambda)\check{G}_2(\lambda))}$$

or

$$\hat{P}(\lambda) = \frac{\hat{p}(\lambda)(\widehat{1 - G_1})(\lambda)}{\alpha (1 - \check{G}_1(\lambda)) \hat{p}(\lambda) + (1 - \alpha)(\widehat{1 - G_1})(\lambda)}$$

in view of (8). Hence, observing that $1 - \check{G}_1(\lambda) = \lambda(\widehat{1 - G_1})(\lambda)$ we deduce

$$(14) \quad \begin{aligned} \hat{P}(\lambda) &= \frac{\hat{p}(\lambda)(\widehat{1 - G_1})(\lambda)}{\alpha \lambda (\widehat{1 - G_1})(\lambda) \hat{p}(\lambda) + (1 - \alpha)(\widehat{1 - G_1})(\lambda)} \\ &= \frac{\hat{p}(\lambda)}{\alpha \lambda \hat{p}(\lambda) + (1 - \alpha)} = \frac{\hat{p}(\lambda)}{1 - \alpha(1 - \lambda\hat{p}(\lambda))} = \frac{\hat{p}(\lambda)}{1 - \alpha\check{G}(\lambda)}. \end{aligned}$$

On the other hand,

$$(15) \quad \int_0^\infty e^{-\lambda t} \left(\frac{1}{\alpha} - \frac{1-\alpha}{\alpha} U_\alpha(t) \right) dt = \frac{1}{\lambda\alpha} - \frac{1-\alpha}{\alpha} \frac{1}{1-\alpha\check{G}(\lambda)}$$

$$= \frac{1-\check{G}(\lambda)}{\lambda(1-\alpha\check{G}(\lambda))} = \frac{\hat{p}(\lambda)}{1-\alpha\check{G}(\lambda)}.$$

Combining (14) and (15) we prove (13). \square

Lemma 8. *$P(t)$ is a monotone decreasing function and*

$$(16) \quad P(t) \sim \frac{1}{1-\alpha} p(t) \sim \frac{\gamma_1}{1-\alpha} t^{-1/2}, \quad t \rightarrow \infty.$$

In addition, there exists a constant $C > 0$ such that

$$(17) \quad |P'(t)| \leq C(t+1)^{-3/2}$$

for all $t \geq 0$.

Proof. The monotonicity of $P(t)$ is a trivial consequence of (13). To establish the remaining statements of the lemma observe that $1 - G(t) = p(t) \sim \gamma_1 t^{-1/2}$ as $t \rightarrow \infty$. It is shown in ([3], Chapter IV, Section 4) that under this condition

$$U_\alpha(t) = \frac{1}{1-\alpha} - \frac{\alpha}{(1-\alpha)^2} p(t) + o(p(t)), \quad t \rightarrow \infty.$$

This equality and the previous lemma complete the proof of (16). To prove inequality (17) we need more delicate arguments. First observe that

$$(18) \quad P'(t) = -\frac{1-\alpha}{\alpha} U'_\alpha(t) = -\frac{1-\alpha}{\alpha} \sum_{k=1}^{\infty} \alpha^k g^{*k}(t), \quad t > 0,$$

where

$$g^{*k}(t) \stackrel{\text{def}}{=} (G^{*k}(t))' = \int_0^t g^{*(k-1)}(t-u)g(u) du$$

is the k th convolution of the density $g(t)$ of $G(t)$. Note that $g(t) = -p'(t)$ and therefore $g(t)$ is monotone decreasing by Lemma 5. Consequently, for any $k > 1$

$$g^{*k}(t) \leq g\left(\frac{t}{k}\right) \int_{tk^{-1}}^t g^{*(k-1)}(t-u) du + g^{*(k-1)}\left(t\frac{k-1}{k}\right) \int_0^{tk^{-1}} g(u) du$$

$$\leq g\left(\frac{t}{k}\right) + g^{*(k-1)}\left(t\frac{k-1}{k}\right) \leq 2g\left(\frac{t}{k}\right) + g^{*(k-2)}\left(t\frac{k-2}{k}\right) \leq kg\left(\frac{t}{k}\right).$$

Using Lemma 5 once more we obtain that

$$g(t) = -p'(t) \sim \frac{\gamma_1}{2} t^{-3/2}, \quad t \rightarrow \infty,$$

whence $g(t)$ is bounded by (4) with $x = 0$. As a result we get that there exists a constant $c > 0$ such that

$$g(t) \leq \frac{c}{(t+1)^{3/2}}$$

for all $t \geq 0$. Hence it follows that

$$g^{*k}(t) \leq kg\left(\frac{t}{k}\right) \leq \frac{ck^{5/2}}{(t+k)^{3/2}} \leq ck^{5/2}(t+1)^{-3/2}, \quad t \geq 0.$$

Substituting this estimate into (18) we get

$$|P'(t)| \leq c(t+1)^{-3/2} \sum_{k=1}^{\infty} k^{5/2} \alpha^k \stackrel{\text{def}}{=} C(t+1)^{-3/2},$$

since $\sum_{k=1}^{\infty} k^{5/2} \alpha^k < \infty$ in view of $\alpha < 1$. This completes the proof of (17). \square

3. BRANCHING RANDOM WALK AND BELLMAN–HARRIS PROCESSES

In this section we prove Theorem 1 by introducing an auxiliary Bellman–Harris branching process with two types of particles.

A Bellman–Harris branching process with two types of particles can be described as follows. It is initiated by a single particle of type i , $i = 1, 2$. This particle has a random life time with a distribution function $G_i(t)$. When dying this particle produces children according to an offspring generating function $f_i(s_1, s_2)$. The new particles of type $j = 1, 2$ evolve independently with the life time distribution $G_j(t)$ and offspring generating function $f_j(s_1, s_2)$. Let $M = \left\| \frac{\partial f_i}{\partial s_j}(1, 1) \right\|_{i,j=1,2}$ be the mean matrix of the process. The process is called critical indecomposable if the maximum (in the absolute value) eigenvalue of M (the Perron root of M) equals 1 and there is an integer n such that all entries of M^n are positive.

Let $v = (v_1, v_2)$ and $u = (u_1, u_2)$ be the left and right positive eigenvectors corresponding to the Perron root of M and such that $u_1 + u_2 = 1$ and $v_1 u_1 + v_2 u_2 = 1$. Denote by $Z_i(t)$, $i = 1, 2$, the number of particles of type i in this process at time t . Set

$$F_i(t; s_1, s_2) = \mathbf{E} s_1^{Z_1(t)} s_2^{Z_2(t)}, \quad i = 1, 2,$$

for the probability generating functions of the number of particles of both types given that the process is initiated at time zero by a single particle of type i .

Now we are ready to recall a result from [12] concerning the critical Bellman–Harris branching processes adapted to the case of two types of particles.

Theorem 9 (see [12]). *Let*

$$x - v_1(1 - f_1(1 - u_1 x, 1 - u_2 x)) - v_2(1 - f_2(1 - u_1 x, 1 - u_2 x)) \sim B^2 x^2$$

as $x \rightarrow +0$ with $B > 0$. If there exist $\beta \in (0, 1)$ and nonnegative constants c_1 and c_2 with $c_1 + c_2 > 0$ such that $1 - G_i(t) = c_i t^{-\beta} + o(t^{-\beta})$ as $t \rightarrow \infty$, then

$$(19) \quad 1 - F_i(t; s_1, s_2) \sim u_i B^{-1} \sqrt{v_1 c_1 (1 - s_1) + v_2 c_2 (1 - s_2)} t^{-\beta/2}, \quad t \rightarrow \infty,$$

and

$$\begin{aligned} & \lim_{t \rightarrow \infty} \mathbf{E} \left[s_1^{Z_1(t)} s_2^{Z_2(t)} \mid Z_1(t) + Z_2(t) > 0; Z_1(0) = 1, Z_2(0) = 0 \right] \\ &= 1 - \sqrt{\frac{v_1 c_1 (1 - s_1) + v_2 c_2 (1 - s_2)}{v_1 c_1 + v_2 c_2}} \end{aligned}$$

for all fixed $s_1, s_2 \in [0, 1)$.

We apply this theorem to the following critical Bellman–Harris process with two types of particles. A particle of the first type has the life time distribution

$$G_1(t) = \mathbf{P}(\tau_1 \leq t) = 1 - e^{-t}, \quad t \geq 0;$$

when dying the particle produces the offspring of two types in accordance with the probability generating function $f_1(s_1, s_2) = \alpha f(s_1) + (1 - \alpha) s_2$, that is, it produces with probability αf_k exactly k particles of the first type and with probability $1 - \alpha$ exactly one particle of the second type (recall the definition of $f(s)$ in (1)). The life time distribution of a particle of the second type is $G_2(t) = \mathbf{P}(\tau_2 \leq t)$ (that is, it coincides in distribution with the time spent outside the origin by the parent particle of the catalytic branching random walk under investigation until the first return to the origin, provided that the initial particle is located at point 0 at time $t = 0$ and it does not produce children during its first stay at 0). When dying a particle of the second type produces the offspring in accordance with the probability generating function $f_2(s_1, s_2) = s_1$, that is, it produces

exactly one particle of the first type and nothing else. It is not difficult to understand that we have $(Z_1(t), Z_2(t)) \stackrel{\text{distr}}{=} (\zeta(t), \mu(t))$ for the process constructed in this way.

Now to prove Theorem 1 it suffices to apply Theorem 9 to $(Z_1(t), Z_2(t))$. It follows from the description above that in this case

$$M = \begin{pmatrix} \alpha & 1 - \alpha \\ 1 & 0 \end{pmatrix}, \quad v_1 = \frac{2}{2 - \alpha}, \quad v_2 = \frac{2(1 - \alpha)}{2 - \alpha}, \quad u_1 = u_2 = \frac{1}{2},$$

and therefore

$$\begin{aligned} & x - v_1(1 - f_1(1 - u_1x, 1 - u_2x)) - v_2(1 - f_2(1 - u_1x, 1 - u_2x)) \\ &= x - \frac{2}{2 - \alpha} \left(1 - f_1\left(1 - \frac{x}{2}, 1 - \frac{x}{2}\right)\right) - \frac{2(1 - \alpha)}{2 - \alpha} \left(1 - f_2\left(1 - \frac{x}{2}, 1 - \frac{x}{2}\right)\right) \\ &= x - \frac{2}{2 - \alpha} \left(\alpha \left(1 - f\left(1 - \frac{x}{2}\right)\right) + (1 - \alpha) \frac{x}{2}\right) - \frac{2(1 - \alpha)x}{2 - \alpha} \frac{1}{2} \\ &= \frac{2\alpha}{2 - \alpha} \left(\frac{x}{2} - \left(1 - f\left(1 - \frac{x}{2}\right)\right)\right) \\ &\sim \frac{2\alpha}{2 - \alpha} \frac{f''(1)}{2} \left(\frac{x}{2}\right)^2 = \frac{\alpha}{2 - \alpha} \frac{f''(1)}{4} x^2 = \frac{\alpha}{2 - \alpha} \frac{\sigma^2}{4} x^2 = B^2 x^2 \end{aligned}$$

as $x \rightarrow +0$. Moreover $\beta = 1/2$ and $c_2 = b\sqrt{2/\pi}$ by Lemma 6, while $1 - G_1(t) = o(t^{-1/2})$ as $t \rightarrow \infty$ implies that $c_1 = 0$.

Using these relations in (19) with $s_1 = s_2 = 0$ we get

$$\begin{aligned} \mathbf{P}(\eta(t) > 0) &= \mathbf{P}(\zeta(t) + \mu(t) > 0) = \mathbf{P}(Z_1(t) + Z_2(t) > 0; Z_1(0) = 1, Z_2(0) = 0) \\ &= 1 - F_1(t; 0, 0) \\ &\sim \frac{1}{2} \sqrt{\frac{2 - \alpha}{\alpha}} \frac{2}{\sigma} \sqrt{\frac{2(1 - \alpha)}{2 - \alpha}} b \sqrt{\frac{2}{\pi}} t^{-1/4} = \frac{2^{3/4}}{\sigma \sqrt[3]{\pi}} \sqrt{\frac{b(1 - \alpha)}{\alpha}} t^{-1/4} = K t^{-1/4}, \\ &\quad t \rightarrow \infty, \end{aligned}$$

and

$$\lim_{t \rightarrow \infty} \mathbf{E} \left[s_1^{\zeta(t)} s_2^{\mu(t)} \mid \zeta(t) + \mu(t) > 0 \right] = 1 - \sqrt{\frac{v_1 c_1 (1 - s_1) + v_2 c_2 (1 - s_2)}{v_1 c_1 + v_2 c_2}} = 1 - \sqrt{1 - s_2}.$$

This shows, in particular, that $\mathbf{P}(\eta(t) > 0) \sim \mathbf{P}(\zeta(t) > 0)$ as $t \rightarrow \infty$ and establishes (3), completing the proof of Theorem 1. Thus if the process with a catalyst at the origin survives, then there are “practically” no particles at the origin and, besides, conditioned on the survival up to time t , the limiting distribution of the particles located outside the origin is discrete.

4. PARTICLES AT THE ORIGIN

The main goal of this section is to investigate the asymptotic behavior of the probability that there are particles at the origin at time t as $t \rightarrow \infty$. Recall that under our assumption on $f_i(s_1, s_2)$, $i = 1, 2$,

$$F_1(t; s_1, s_2) = s_1(1 - G_1(t)) + \int_0^t (\alpha f(F_1(t - u; s_1, s_2)) + (1 - \alpha)F_2(t - u; s_1, s_2)) dG_1(u)$$

and

$$F_2(t; s_1, s_2) = s_2(1 - G_2(t)) + \int_0^t F_1(t - u; s_1, s_2) dG_2(u)$$

(see [11], Chapter VIII, §1). Substituting the second of these equalities into the first one we get

$$\begin{aligned} F_1(t; s_1, s_2) &= s_1(1 - G_1(t)) + s_2(1 - \alpha)(1 - G_2(\cdot)) * G_1(t) \\ &\quad + \int_0^t \alpha f(F_1(t - u; s_1, s_2)) dG_1(u) \\ &\quad + \int_0^t (1 - \alpha)F_1(t - u; s_1, s_2) d(G_1 * G_2(u)). \end{aligned}$$

In particular, letting $F(t; s) \stackrel{\text{def}}{=} \mathbf{E} s^{\mu(t)} = F_1(t; s, 1)$ we obtain

$$\begin{aligned} (20) \quad F(t; s) &= s(1 - G_1(t)) + (1 - \alpha)(1 - G_2(\cdot)) * G_1(t) \\ &\quad + \int_0^t \alpha f(F(t - u; s)) dG_1(u) \\ &\quad + \int_0^t (1 - \alpha)F(t - u; s) d(G_1 * G_2(u)). \end{aligned}$$

Hence

$$q(t; s) = (1 - s)(1 - G_1(t)) + q(\cdot; s) * (\alpha G_1 + (1 - \alpha)G_1 * G_2)(t) - \alpha h(q(\cdot; s)) * G_1(t),$$

where $q(t; s) \stackrel{\text{def}}{=} 1 - F(t; s)$ and $h(x) \stackrel{\text{def}}{=} 1 - x - f(1 - x)$. Let $q(t) \stackrel{\text{def}}{=} q(t; 0)$. Then

$$q(t) = 1 - G_1(t) + q(\cdot) * G_3(t) - \alpha h(q(\cdot)) * G_1(t)$$

in view of (11). Solving this renewal equation with respect to $q(t)$ we see that

$$(21) \quad q(t) = (1 - G_1(\cdot)) * U(t) - \alpha h(q(\cdot)) * G_1 * U(t),$$

where $U(t)$ is the same as in (12). Since $G_1(t) = 1 - e^{-t}$, it follows that

$$(G_1 * U(t))' = (1 - G_1(\cdot)) * U(t) = P(t).$$

This allows us to rewrite (21) as follows:

$$(22) \quad q(t) = P(t) - \alpha \int_0^t h(q(t - u))P(u) du.$$

Now we are ready to start the proof of Theorem 2. We divide the proof into several steps. The first lemma gives an estimate from below for $q(t) = \mathbf{P}(\zeta(t) > 0)$.

Lemma 10. *Under the condition of Theorem 2 there exists a constant $C_1 > 0$ such that*

$$q(t) > \frac{C_1}{\sqrt{t} \ln t}$$

for all sufficiently large t .

Proof. Set $A(t) = \mathbf{E} \zeta(t)$ and $B(t) = \mathbf{E} \zeta(t)(\zeta(t) - 1)$. Differentiating (20) with respect to s at the point $s = 1$ and using $f'(1) = 1$ we get

$$A(t) = 1 - G_1(t) + A * G_3(t) = (1 - G_1(\cdot)) * U(t) = P(t)$$

and

$$\begin{aligned} B(t) &= \alpha A^2 * G_1(t) + B * G_3(t) = \alpha A^2 * G_1 * U(t) \\ &= \alpha \int_0^t P^2(t - u) d(G_1 * U)(u) = \alpha \int_0^t P^2(t - u) P(u) du. \end{aligned}$$

Using (16) we deduce

$$B(t) \sim \frac{\alpha \gamma_1^3}{(1 - \alpha)^3} \frac{\ln t}{\sqrt{t}},$$

whence

$$q(t) = \mathbb{P}(\zeta(t) > 0) \geq \frac{\mathbb{E}^2 \zeta(t)}{\mathbb{E} \zeta^2(t)} \sim \frac{1 - \alpha}{\alpha \gamma_1} \frac{1}{\sqrt{t \ln t}}, \quad t \rightarrow \infty,$$

which proves the lemma. \square

Lemma 11.

$$\alpha \int_0^\infty h(q(u)) du = 1.$$

Proof. Applying the Laplace transform to both sides of (22) we get

$$(23) \quad \hat{q}(\lambda) = \hat{P}(\lambda)(1 - \alpha \hat{h}(\lambda))$$

where we set

$$\hat{h}(\lambda) = \int_0^\infty e^{-\lambda t} h(q(t)) dt.$$

Since $\hat{h}(\lambda)$ is monotone increasing in λ as $\lambda \rightarrow +0$, the limit $\lim_{\lambda \rightarrow +0} \alpha \hat{h}(\lambda) = c_0 \leq 1$ exists and, in particular, $\int_0^\infty h(q(t)) dt < \infty$. Now we show that $c_0 = 1$. Indeed, assuming the contrary we obtain from (9) and (14) that

$$\hat{q}(\lambda) \sim (1 - c_0) \hat{P}(\lambda) \sim \frac{(1 - c_0) \gamma_1 \sqrt{\pi}}{1 - \alpha} \frac{1}{\sqrt{\lambda}}, \quad \lambda \rightarrow +0,$$

whence

$$\int_0^t q(u) du \sim \frac{2(1 - c_0)}{1 - \alpha} \gamma_1 \sqrt{t}, \quad t \rightarrow \infty.$$

Thus

$$\int_t^{2t} q(u) du \sim (\sqrt{2} - 1) \frac{2(1 - c_0)}{1 - \alpha} \gamma_1 \sqrt{t}, \quad t \rightarrow \infty.$$

On the other hand, by Hölder's inequality,

$$\frac{1}{t} \left(\int_t^{2t} q(u) du \right)^2 \leq \int_t^{2t} q^2(u) du \leq \frac{4}{f''(1)} \int_t^{2t} h(q(u)) du$$

for large t , since $h(x) = 1 - x - f(1 - x) \sim f''(1)x^2/2$ as $x \rightarrow +0$.

Thus if $c_0 < 1$, then the left-hand side of the preceding relation tends to a positive constant as $t \rightarrow \infty$, while the right-hand side tends to zero. This contradiction shows that $c_0 = 1$ as desired. \square

Corollary 12. *Under the conditions of Theorem 2*

$$q(t) = o(P(t)), \quad t \rightarrow \infty.$$

Proof. Using (22) and Lemma 11 we have

$$q(t) \leq P(t) \left(1 - \alpha \int_0^t h(q(t - u)) du \right) = \alpha P(t) \int_t^\infty h(q(u)) du = o(P(t)),$$

$t \rightarrow \infty. \quad \square$

The next lemma is one of the crucial steps in the proof of Theorem 2.

Lemma 13. *Under the conditions of Theorem 2 the function $L(\lambda) \stackrel{\text{def}}{=} \hat{q}(\lambda)/\hat{P}(\lambda)$ is slowly varying as $\lambda \rightarrow +0$.*

Proof. It follows from (23) that for any $c > 0$ there exists a constant $C_1 > 0$ such that

$$(24) \quad \begin{aligned} L(\lambda) - L(c\lambda) &= \alpha(\hat{h}(\lambda) - \hat{h}(c\lambda)) = \alpha \int_0^\infty e^{-\lambda t} \left(1 - e^{-\lambda(c-1)t}\right) h(q(t)) dt \\ &\leq C_1 \alpha \int_0^\infty e^{-\lambda t} \left(1 - e^{-\lambda(c-1)t}\right) q^2(t) dt. \end{aligned}$$

Our goal is to show that the right-hand side of (24) is $o(L(\lambda))$ as $\lambda \rightarrow +0$. Put

$$\begin{aligned} I_1(\lambda) &\stackrel{\text{def}}{=} \int_0^{1/\lambda} e^{-\lambda t} \left(1 - e^{-\lambda(c-1)t}\right) q^2(t) dt, \\ I_2(\lambda) &\stackrel{\text{def}}{=} \int_{1/\lambda}^\infty e^{-\lambda t} \left(1 - e^{-\lambda(c-1)t}\right) q^2(t) dt. \end{aligned}$$

In view of the inequality $1 - e^{-x} \leq x$, $x \geq 0$, we have

$$\begin{aligned} I_1(\lambda) &\leq \lambda(c-1) \int_0^{1/\lambda} e^{-\lambda t} t q^2(t) dt \leq c\sqrt{\lambda} \int_0^{1/\lambda} e^{-\lambda t} \sqrt{t} q^2(t) dt \\ &\leq c\sqrt{\lambda} \int_0^\infty e^{-\lambda t} \left(\sqrt{t} q(t)\right) q(t) dt. \end{aligned}$$

Note that

$$\int_0^\infty e^{-\lambda t} \left(\sqrt{t} q(t)\right) q(t) dt = o\left(\int_0^\infty e^{-\lambda t} q(t) dt\right), \quad \lambda \rightarrow +0.$$

Indeed, by Lemma 10

$$\int_0^\infty q(t) dt = \infty$$

and by Corollary 12

$$\sqrt{t} q(t) = o\left(\sqrt{t} P(t)\right) = o(1), \quad t \rightarrow \infty.$$

Hence the desired relation follows. Using the equivalence

$$(25) \quad \hat{P}(\lambda) \sim \frac{1}{1-\alpha} \frac{\gamma_1 \sqrt{\pi}}{\sqrt{\lambda}}, \quad \lambda \rightarrow +0,$$

we get

$$I_1(\lambda) = o\left(\sqrt{\lambda} \int_0^\infty e^{-\lambda t} q(t) dt\right) = o\left(\frac{\hat{q}(\lambda)}{\hat{P}(\lambda)}\right) = o(L(\lambda)), \quad \lambda \rightarrow +0.$$

On the other hand,

$$\begin{aligned} I_2(\lambda) &\leq \int_{1/\lambda}^\infty e^{-\lambda t} q^2(t) dt \leq \sup_{u \geq \lambda^{-1}} q(u) \int_{1/\lambda}^\infty e^{-\lambda t} q(t) dt \\ &= o\left(P(1/\lambda) \int_0^\infty e^{-\lambda t} q(t) dt\right) = o\left(\sqrt{\lambda} \hat{q}(\lambda)\right) = o(L(\lambda)), \quad \lambda \rightarrow +0. \end{aligned}$$

Therefore

$$\frac{L(\lambda c)}{L(\lambda)} = 1 + \frac{L(\lambda c) - L(\lambda)}{L(\lambda)} \rightarrow 1, \quad \lambda \rightarrow +0,$$

for all $c > 1$ and this is just the definition of a slowly varying function (see [7], Chapter VIII, §8). \square

Corollary 14. *The integral*

$$\int_t^\infty h(q(u)) du$$

is a slowly varying function as $t \rightarrow \infty$.

Proof. Evidently,

$$L(\lambda) = 1 - \alpha \int_0^\infty e^{-\lambda t} h(q(t)) dt = \lambda \alpha \int_0^\infty e^{-\lambda t} \left(\int_t^\infty h(q(u)) du \right) dt,$$

whence

$$\alpha \int_0^\infty e^{-\lambda t} \left(\int_t^\infty h(q(u)) du \right) dt = \frac{L(\lambda)}{\lambda}$$

is a regularly varying function with parameter 1 as $\lambda \rightarrow +0$. Since

$$\int_t^\infty h(q(u)) du$$

is monotone decreasing in t , Theorem 4 in [7], Chapter XIII, §5 yields

$$\alpha \int_t^\infty h(q(u)) du \sim L(t^{-1}), \quad t \rightarrow \infty,$$

as required. □

Corollary 15. *We have*

$$\int_0^t q(u) du \sim \frac{2\gamma_1}{1-\alpha} \sqrt{t} L(t^{-1})$$

as $t \rightarrow \infty$.

Proof. By (23) and (25),

$$\hat{q}(\lambda) = \hat{P}(\lambda) (1 - \alpha \hat{h}(\lambda)) = \hat{P}(\lambda) L(\lambda) \sim \frac{1}{1-\alpha} \frac{\gamma_1 \sqrt{\pi}}{\sqrt{\lambda}} L(\lambda), \quad \lambda \rightarrow +0.$$

Hence the corollary follows by applying a Tauberian theorem ([7], Chapter XIII, §5). □

Lemma 16. *As $t \rightarrow \infty$,*

$$J(t) \stackrel{\text{def}}{=} \int_0^t h(q(u))(P(t-u) - P(t)) du = o(P(t)L(t^{-1})).$$

Proof. Fix a number $\varepsilon \in (0, 1/2)$ and split $J(t)$ into three integrals:

$$J(t) = J_1(t) + J_2(t) + J_3(t),$$

where

$$\begin{aligned} 0 \leq J_1(t) &\stackrel{\text{def}}{=} \int_{t(1-\varepsilon)}^t h(q(u))(P(t-u) - P(t)) du, \\ 0 \leq J_2(t) &\stackrel{\text{def}}{=} \int_{t\varepsilon}^{t(1-\varepsilon)} h(q(u))(P(t-u) - P(t)) du, \\ 0 \leq J_3(t) &\stackrel{\text{def}}{=} \int_0^{t\varepsilon} h(q(u))(P(t-u) - P(t)) du. \end{aligned}$$

Since $P(t)$ is monotone,

$$J_1(t) \leq \int_{t(1-\varepsilon)}^t h(q(u))P(t-u) du = \int_0^{t\varepsilon} h(q(t-u))P(u) du.$$

Recall that

$$q(t) \leq P(t) \int_t^\infty h(q(u)) du \leq cP(t)L(t^{-1}), \quad c > 0.$$

Properties of slowly varying functions and the inequality $h(x) \leq \sigma^2 x^2$ for sufficiently small $x > 0$ imply by Lemma 8 that

$$\begin{aligned} J_1(t) &\leq \sup_{v \in [t(1-\varepsilon), t]} h(q(v)) \int_0^{t\varepsilon} P(u) du \\ (26) \quad &\leq c_1 P^2(t) L^2(t^{-1}) \int_0^{t\varepsilon} P(u) du \leq c_2 P^2(t) L^2(t^{-1}) \sqrt{\varepsilon t} \\ &\leq c_3 P(t) L^2(t^{-1}) \sqrt{\varepsilon} = o(P(t)L(t^{-1})), \quad t \rightarrow \infty, \end{aligned}$$

since $L(t^{-1}) \rightarrow 0$ as $t \rightarrow \infty$. Using Corollary 14 and properties of slowly varying functions we see that

$$\begin{aligned} J_2(t) &\leq P(t\varepsilon) \int_{t\varepsilon}^{t(1-\varepsilon)} h(q(u)) du \\ (27) \quad &= O\left(P(t\varepsilon) \left(L((\varepsilon t)^{-1}) - L((1-\varepsilon)t^{-1})\right)\right) \\ &= o(P(t)L(t^{-1})), \quad t \rightarrow \infty. \end{aligned}$$

By the mean value theorem, Lemma 8, and the Tauberian theorems,

$$\begin{aligned} J_3(t) &\leq \sup_{v \in [t(1-\varepsilon), t]} |P'(v)| \int_0^{t\varepsilon} u h(q(u)) du \leq ct^{-3/2} \int_0^{t\varepsilon} u q^2(u) du \\ (28) \quad &\leq c_1 t^{-3/2} \int_0^{t\varepsilon} u P^2(u) L^2(u^{-1}) du \leq c_2 t^{-3/2} \int_0^{t\varepsilon} L^2(u^{-1}) du \\ &\leq c_3 t^{-3/2} t\varepsilon L^2((\varepsilon t)^{-1}) = o(P(t)L(t^{-1})), \quad t \rightarrow \infty. \end{aligned}$$

Combining (26)–(28) we complete the proof of Lemma 16. □

Corollary 17. *Under the conditions of Theorem 2*

$$(29) \quad q(t) = P(t) \frac{\alpha \sigma^2}{2} \int_t^\infty q^2(u) du (1 + \varepsilon(t))$$

where $\varepsilon(t) \rightarrow \infty$ as $t \rightarrow \infty$.

Proof. Indeed, by (22), Lemma 16, and the equivalence $h(x) \sim \sigma^2 x^2/2$ as $x \rightarrow +0$, we have

$$\begin{aligned} q(t) &= P(t) - \alpha \int_0^t h(q(u)) P(t-u) du = \alpha P(t) \int_t^\infty h(q(u)) du - \alpha J(t) \\ &= \alpha P(t) \int_t^\infty h(q(u)) du (1 + o(1)) = P(t) \frac{\alpha \sigma^2}{2} \int_t^\infty q^2(u) du (1 + \varepsilon(t)). \quad \square \end{aligned}$$

Proof of Theorem 2. Set

$$y(t) = \int_t^\infty q^2(u) du.$$

Then (29) can be rewritten as follows:

$$\sqrt{-y'(t)} = P(t) \frac{\alpha \sigma^2}{2} y(t) (1 + \varepsilon(t))$$

or

$$\begin{aligned}
 (30) \quad -y'(t) &= \left(P(t) \frac{\alpha\sigma^2}{2} y(t) \right)^2 (1 + \varepsilon_1(t)) \\
 &= \frac{1}{t+1} \left(\frac{\gamma_1}{1-\alpha} \right)^2 \left(\frac{\alpha\sigma^2}{2} \right)^2 y^2(t) (1 + \varepsilon_2(t)) = \frac{K_1}{t+1} y^2(t) (1 + \varepsilon_2(t))
 \end{aligned}$$

in view of (16) where

$$K_1 \stackrel{\text{def}}{=} \left(\frac{\gamma_1}{1-\alpha} \right)^2 \left(\frac{\alpha\sigma^2}{2} \right)^2$$

and $\varepsilon_i(t)$, $i = 1, 2, \dots$, denote functions vanishing as $t \rightarrow \infty$. Solving (30) under the initial condition

$$y(0) = C_0^{-1} \stackrel{\text{def}}{=} \int_0^\infty q^2(u) du$$

we obtain

$$y(t) = \left(C_0 + K_1 \int_0^t \frac{(1 + \varepsilon_2(u))}{u+1} du \right)^{-1} \sim \frac{1}{K_1 \ln t}, \quad t \rightarrow \infty.$$

Substituting this equivalence into (29) we see that

$$\begin{aligned}
 q(t) &\sim P(t) \frac{\alpha\sigma^2}{2} \frac{1}{K_1 \ln t} = \frac{2}{\alpha\sigma^2} \left(\frac{1-\alpha}{\gamma_1} \right)^2 \frac{P(t)}{\ln t} \\
 &\sim \frac{2}{\alpha\sigma^2} \frac{1-\alpha}{\gamma_1} \frac{1}{\sqrt{t \ln t}} = \frac{2(1-\alpha)}{\alpha\sigma^2} b\sqrt{2\pi} \frac{1}{\sqrt{t \ln t}} = K^2 \pi \frac{1}{\sqrt{t \ln t}}
 \end{aligned}$$

as $t \rightarrow \infty$. This relation proves Theorem 2. \square

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