

GENERALIZATION OF A PROBABILITY LIMIT THEOREM OF CRAMÉR

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1. **Introduction.** Let $\{X_k\}$, $k=1, 2, \dots$, be mutually independent random variables, that is to say independent real functions in a space in which a probability measure is defined. Let $V_k(x) = \text{Prob} \{X_k \leq x\}$ be the distribution function of X_k and suppose that the second moment

$$(1.1) \quad \sigma_k^2 = \int_{-\infty}^{+\infty} x^2 dV_k(x)$$

exists; without loss of generality we may then assume that

$$(1.2) \quad \int_{-\infty}^{+\infty} x dV_k(x) = 0.$$

Let

$$(1.3) \quad S_n = X_1 + \dots + X_n$$

and

$$(1.4) \quad s_n^2 = \sigma_1^2 + \dots + \sigma_n^2.$$

Finally, let $F_n(x) = \text{Prob} \{S_n \leq x\}$ be the distribution function of S_n . The independence of the X_n implies that

$$(1.5) \quad F_1(x) = V_1(x), \quad F_{n+1}(x) = \int_{-\infty}^{+\infty} F_n(x-y) dV_{n+1}(y).$$

The central limit theorem asserts that, under very general conditions,

$$(1.6) \quad F_n(xs_n) \rightarrow \Phi(x) = (2\pi)^{-1/2} \int_{-\infty}^x e^{-(1/2)y^2} dy,$$

uniformly in every finite interval. If, however, with $n \rightarrow \infty$ also $x \rightarrow \infty$ it follows only that $F_n(xs_n) \rightarrow 1$, but it is not even clear whether $1 - F(xs_n)$ is of the same order of magnitude as $1 - \Phi(x)$. In the Bernoulli case with

$$V_k(x) \equiv V(x) = \begin{cases} 0 & \text{for } x < -p < 0, \\ 1-p & \text{for } -p \leq x < 1-p, \\ p & \text{for } 1-p \leq x, \end{cases}$$

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the behavior of $F_n(xs_n)$ when $x=x(n)\rightarrow\infty$ has been the object of detailed investigations by A. Khintchine [5]⁽¹⁾, N. Smirnoff [7], P. Lévy [6] and M. Fréchet [4]. A more general case has recently been treated by H. Cramér [2]; he considers the case of "equal components" where all $V_k(x)$ are identical, $V_k(x)\equiv V(x)$, and it is moreover supposed that the integral

$$(1.7) \quad \int_{-\infty}^{+\infty} e^{ax} dV(x)$$

exists for all $|a| < a_0 \neq 0$. Nothing seems to be known in the case of unequal components. The problem as such has been stated by P. Lévy [6, p. 289].

The purpose of the present paper is to generalize Cramér's result to the case of unequal components. This generalization seems of interest in itself, but it is also to serve as the basis for the establishment of the general form of the so-called law of the iterated logarithm.

We shall consider only the case where $|X_n|$ has an upper bound of the form $o(s_n)$. Since $s_n \rightarrow \infty$, this restriction is not very strong. In some cases it is possible to free oneself from this restriction, provided the "tails" $\int_{|x| > \epsilon s_n} dV_n(x)$ are sufficiently small. However, in many applications the principle of equivalent sequences of random variables will automatically reduce the consideration to the case of bounded variables. Our proof is based on a straightforward adaptation of Cramér's proof, although we shall avoid the use of characteristic functions. Cramér's proof, in turn, rests on a transformation first used by F. Esscher [3] in connection with an insurance problem. In the special case of equal components our condition $|X_n| = o(s_n)$ is more restrictive than Cramér's condition that the integral (1.7) should exist. On the other hand, our results will be slightly sharper than Cramér's, since it will be possible for us to free ourselves of the logarithmic term in the remainder without introducing a new hypothesis.

2. Results. Throughout what follows we shall suppose that the $\{X_k\}$ are subject to the conditions (1.1) and (1.2). Moreover, we shall suppose that (cf. (1.4))

$$(2.1) \quad |X_k| < \lambda_n s_n, \quad \text{for } k = 1, \dots, n.$$

In most applications the $\{X_k\}$ will form an infinite sequence, and $\lambda_n \rightarrow 0$. Condition (2.1) can then be replaced by $|X_n| < \lambda_n s_n$, provided that the sequence λ_n is non-increasing. It should be noted, however, that Theorem 1 is strictly a theorem on a finite sequence of random variables.

In §§3-5 we shall prove

THEOREM 1. *If*

$$(2.2) \quad 0 < \lambda_n x < (3 - 5^{1/2})/4$$

⁽¹⁾ The numbers in brackets refer to the bibliography at the end of the paper.

then

$$(2.3) \quad 1 - F_n(x s_n) = e^{-(1/2)x^2 Q_n(x)} \left[\{1 - \Phi(x)\} + \theta \lambda_n e^{-(1/2)x^2} \right].$$

Here θ has an upper bound independent of the sequence $\{X_k\}$ and of n ; furthermore

$$(2.4) \quad Q_n(x) = \sum_{\nu=1}^{\infty} q_{n,\nu} x^\nu$$

where the coefficient $q_{n,\nu}$ depends only on the ν first moments of $V_1(x), \dots, V_n(x)$. If, more particularly,

$$(2.5) \quad 0 < \lambda_n x < 1/12$$

then

$$(2.6) \quad |\theta| < 9$$

and

$$(2.7) \quad |q_{n,\nu}| < (1/7)(12\lambda_n)^\nu.$$

Moreover, for any $0 < i < j \leq n$ we have

$$(2.8) \quad |Q_j(x) - Q_i(x)| \leq (1/2)(s_j^2 - s_i^2)/s_n^2.$$

Before describing the construction of $q_{n,\nu}$ in general we proceed to describe a few implications of Theorem 1. First let us observe that by applying the theorem to the sequence $\{-X_k\}$, we obtain a corresponding theorem for negative x . Since however $Q_n(x)$ depends also on moments of odd order the behavior of $F_n(x s_n)$ for $x \rightarrow -\infty$ will, in general, be different.

Since

$$(2.9) \quad 1 - \Phi(x) = (2\pi)^{-1/2} x^{-1} e^{-(1/2)x^2} (1 - \vartheta/x^2)$$

with $0 < \vartheta < 1$ we can write (2.3) in the form

$$(2.10) \quad 1 - F_n(x s_n) = (2\pi)^{-1/2} x^{-1} e^{-(1/2)x^2(1+Q_n(x))} [1 - \vartheta/x^2 + (2\pi)^{1/2} \theta \lambda_n x],$$

which shows the order of magnitude of the remainder as compared with that of the main term.

To appreciate the meaning of the theorem consider the representative case of a uniformly bounded sequence $\{X_k\}$ with $s_n^2 \sim n$. Then $\lambda_n = O(n^{-1/2})$. If, then, $x = o(n^{1/6})$ it is seen that $Q_n(x) \rightarrow 0$ and $\lambda_n x_n = o(n^{-1/3})$; thus $1 - F_n(x s_n)$ will be of the order of magnitude of $1 - \Phi(x)$. However, if x is of an order of magnitude between $n^{1/6}$ and $n^{1/4}$, the order of magnitude of $1 - F_n(x s_n)$ will be that of $\{1 - \Phi(x)\} e^{-(1/2)q_{n,1}x^2}$; for an x between $n^{1/4}$ and $n^{3/10}$ the term $(1/2)q_{n,2}x^4$ will be significant also, whereas $\sum_{\nu=5}^{\infty} q_{n,\nu} x^{\nu+2} \sim 0$, and so on. Still

$$Q_n(x) = o(1) \quad \text{for } x = o(n^{1/2}),$$

so that for $x = o(n^{1/2})$ at least $\log(1 - F_n(xs_n)) \sim \log(1 - \phi(x))$. Now if $x \sim n^{1/2}$, the behavior is quite different. Then $(x^2/2)Q_n(x) \sim \text{const. } x^2$, and $\log(1 - F(xs_n))$ will be asymptotically equivalent to $(\alpha/2)x^2$ with $\alpha \neq 1$; in other words $\log(1 - F_n(xs_n))$ will no longer be asymptotically equivalent to $\log(1 - \phi(x))$. A similar change in the asymptotic behavior of $1 - F_n(xs_n)$ can be observed in the general case and is of fundamental importance in connection with the law of the iterated logarithm.

For use in the same connection we note also the following simple corollary to Theorem 1.

THEOREM 2. *Let $\alpha > 0$ be an arbitrary constant and let $\max(2, \alpha) < x < 1/(100\lambda_n)$ and $\lambda_n < 1/100$. Then there exist two constants $0 < a_1 < a_2 < 1$, depending only on α , such that*

$$(2.11) \quad a_1 < \frac{1 - F((x + \alpha/x)s_n)}{1 - F(xs_n)} < a_2.$$

Proof. The theorem is an easy consequence of (2.10). The term in brackets lies obviously between $1/2$ and $3/2$. From (2.7) we conclude furthermore that

$$\begin{aligned} |q_{n,\nu}(x + \alpha/x)^{\nu+2} - q_{n,\nu}x^{\nu+2}| &\leq |q_{n,\nu}|(\nu + 2)(x + \alpha/x)^{\nu+1}\alpha/x \\ &\leq \frac{3\alpha(\nu + 2)}{14} \cdot \left(\frac{24}{100}\right)^\nu, \end{aligned}$$

so that $\alpha \leq |(x + \alpha/x)^2(1 + Q_n(x + \alpha/x)) - x^2(1 + Q_n(x))| \leq 3\alpha$.

Choosing, say, $a_2 = 3e^{-\alpha/2}$ and $a_1 = (1/5)e^{-3\alpha/2}$ the theorem follows.

Theorem 2 is essentially equivalent to the following more elegant theorem which has been discovered by Khintchine [5] in the case of the binomial and generalized by Cramér [2] to the case of equal components.

THEOREM 3. *If $x \rightarrow \infty$, $x = o(1/\lambda_n)$, then*

$$\lim_{n \rightarrow \infty} \frac{1 - F_n((x + \alpha/x)s_n)}{1 - F_n(xs_n)} = \lim_{n \rightarrow \infty} \frac{1 - \Phi(x + \alpha/x)}{1 - \Phi(x)} = e^{-\alpha}.$$

The proof is similar to the preceding one.

We proceed, finally, to describe the *construction of the function $Q_n(x)$* . Let⁽²⁾

$$(2.12) \quad b_{k,\nu} = \int_{-\infty}^{+\infty} x^\nu dV_k(x)$$

and define quantities $\gamma_{k,\nu}$ (Thiele's semi-invariants of $V_k(x)$) by

$$(2.13) \quad \log \int_{-\infty}^{+\infty} e^{hy} dV_k(y) = \sum_{\nu=2}^{\infty} \gamma_{k,\nu} \frac{h^\nu}{\nu!};$$

⁽²⁾ Note that here as in the following the integral actually extends only over the interval $|x| < \lambda_n s_n$.

(we have in particular

$$(2.14) \quad \begin{aligned} \gamma_{k,2} &= b_{k,2}, & \gamma_{k,3} &= b_{k,3}, & \gamma_{k,4} &= b_{k,4} - 3b_{k,2}^2, \\ \gamma_{k,5} &= b_{k,5} - 10b_{k,2}b_{k,3}, \dots \end{aligned}$$

while $\gamma_{k,1} = 0$ by (1.2)). We put

$$(2.15) \quad \Gamma_{n,\nu} = \sum_{k=2}^n \gamma_{k,\nu}.$$

Then

$$(2.16) \quad \frac{x^2}{2} + \frac{x^2}{2} Q_n(x) = \sum_{\nu=2}^{\infty} \Gamma_{n,\nu} \frac{\nu-1}{\nu!} h^\nu$$

where $h = h(x)$ is defined as the inverse function of

$$(2.17) \quad x = \frac{1}{s_n} \sum_{\nu=2}^{\infty} \Gamma_{n,\nu} \frac{h^{\nu-1}}{(\nu-1)!}.$$

For the first two coefficients we obtain

$$(2.18) \quad q_{n,1} = \frac{1}{3s_n^3} \sum_{k=1}^n b_{k,3},$$

$$(2.19) \quad q_{n,2} = \frac{1}{12s_n^4} \sum_{k=1}^n b_{k,4} - \frac{1}{4s_n^4} \sum_{k=1}^n b_{k,2}^2 - \frac{1}{4s_n^6} \left(\sum_{k=1}^n b_{k,3} \right)^2.$$

3. A transformation. Following Esscher and Cramér we shall introduce a transformation which will reduce $\{V_k(x)\}$ to a new sequence of distribution functions $\{\bar{V}_k(x)\}$ for which the central limit theorem will become applicable. For this purpose we introduce a real parameter $h > 0$ which will later on be specialized as a function of x and n .

We put

$$(3.1) \quad p_k = \int_{-\infty}^{+\infty} e^{hx} dV_k(x),$$

and

$$(3.2) \quad \bar{V}_k(x) = \frac{1}{p_k} \int_{-\infty}^x e^{hy} dV_k(y).$$

Then $\bar{V}_k(x)$ is again a distribution function. Moreover, $\bar{V}_k(x)$ obviously varies only within a finite interval. Hence the moments

$$(3.3) \quad \bar{m}_k = \int_{-\infty}^{+\infty} x d\bar{V}_k(x)$$

and

$$(3.4) \quad \bar{\sigma}_k^2 = \int_{-\infty}^{+\infty} (x - \bar{m}_k)^2 d\bar{V}_k(x) = \int_{-\infty}^{+\infty} x^2 d\bar{V}_k(x) - \bar{m}_k^2$$

exist. In a manner analogous to (1.4) and (1.5) we define

$$(3.5) \quad \bar{M}_n = \bar{m}_1 + \cdots + \bar{m}_n,$$

$$(3.6) \quad \bar{s}_n^2 = \bar{\sigma}_1^2 + \cdots + \bar{\sigma}_n^2,$$

and

$$(3.7) \quad \begin{aligned} \bar{F}_1(x) &= \bar{V}_1(x), \\ \bar{F}_{n+1}(x) &= \int_{-\infty}^{+\infty} \bar{F}_n(x-y) d\bar{V}_{n+1}(y). \end{aligned}$$

Now by (3.2)

$$(3.8) \quad \int_{-\infty}^x e^{-hy} d\bar{V}_k(y) = \frac{1}{p_k} \int_{-\infty}^x dV_k(y) = \frac{1}{p_k} V_k(x).$$

Hence it is seen that

$$(3.9) \quad F_n(x) = p_1 \cdots p_n \int_{-\infty}^x e^{-hy} d\bar{F}_n(y)$$

holds for $n=1$; using (1.5) and (3.7) it is seen that (3.9) holds generally. Since $F_n(x)$ is a distribution function we have, letting $x \rightarrow \infty$,

$$(3.10) \quad 1 - F_n(x) = p_1 \cdots p_n \int_x^{\infty} e^{-hy} d\bar{F}_n(y).$$

Now it is our purpose to apply the central limit theorem to $\bar{F}_n(x)$. It will therefore be convenient to write (3.10) in the form

$$(3.11) \quad 1 - F_n(x) = p_1 \cdots p_n \int_{(x-\bar{M}_n)/\bar{s}_n}^{\infty} e^{-h(\bar{M}_n+y\bar{s}_n)} d\bar{F}_n(\bar{M}_n+y\bar{s}_n).$$

This formula will be used for $x = \bar{M}_n$ only. Introducing the quantity

$$(3.12) \quad P_n = \sum_{k=1}^n \log p_k$$

we deduce from (3.11) finally

$$(3.13) \quad 1 - F_n(\bar{M}_n) = e^{P_n - h\bar{M}_n} \int_0^{\infty} e^{-hy\bar{s}_n} d\bar{F}_n(\bar{M}_n + y\bar{s}_n).$$

4. **Some estimates.** Operating for the moment in a purely formal manner we have by (2.13) and (3.1)

$$(4.1) \quad \log p_k = \sum_{\nu=2}^{\infty} \gamma_{k,\nu} \frac{h^\nu}{\nu!}.$$

A simple computation shows that (cf. (3.3) and (3.4))

$$(4.2) \quad \bar{m}_k = \frac{d \log p_k}{dh} = \sum_{\nu=2}^{\infty} \gamma_{k,\nu} \frac{h^{\nu-1}}{(\nu-1)!},$$

$$(4.3) \quad \bar{\sigma}_k^2 = \frac{d^2 \log p_k}{dh^2} = \sum_{\nu=2}^{\infty} \gamma_{k,\nu} \frac{h^{\nu-2}}{(\nu-2)!}.$$

Using (2.16), (3.5), and (3.6) we obtain

$$(4.4) \quad P_n = \sum_{\nu=2}^{\infty} \Gamma_{n,\nu} \frac{h^\nu}{\nu!},$$

$$(4.5) \quad \bar{M}_n = \sum_{\nu=2}^{\infty} \Gamma_{n,\nu} \frac{h^{\nu-1}}{(\nu-1)!},$$

$$(4.6) \quad \bar{S}_n^2 = \sum_{\nu=2}^{\infty} \Gamma_{n,\nu} \frac{h^{\nu-2}}{(\nu-2)!}.$$

We now proceed to find majorating power series for these expressions. It follows from (2.1) and (2.12) that for $k=1, \dots, n$

$$(4.7) \quad |b_{k,\nu}| \leq \sigma_k^2 (\lambda_n s_n)^{\nu-2} \leq (\lambda_n s_n)^\nu.$$

Now by (1.2)

$$(4.8) \quad p_k = 1 + \sum_{\nu=2}^{\infty} b_{k,\nu} \frac{h^\nu}{\nu!}.$$

We shall actually use only values $0 < \lambda_n s_n h < 1/5$, so that p_k remains near 1. It follows from (4.1) and (4.7) that

$$(4.9) \quad \begin{aligned} \sum_{\nu=2}^{\infty} \gamma_{k,\nu} \frac{h^\nu}{\nu!} &< -\log \left\{ 1 - \frac{\sigma_k^2}{\lambda_n^2 s_n^2} \sum_{\nu=2}^{\infty} \frac{(\lambda_n s_n h)^\nu}{\nu!} \right\} \\ &= -\log \left\{ 1 - \frac{\sigma_k^2}{\lambda_n^2 s_n^2} (e^{\lambda_n s_n h} - 1 - \lambda_n s_n h) \right\}, \end{aligned}$$

where the sign $<$ indicates that the power series on the left is majorated by that on the right. Now write for the moment

$$(4.10) \quad -\log \{1 - (e^x - 1 - x)\} = \sum_{\nu=2}^{\infty} \rho_{\nu} \frac{x^{\nu}}{\nu!}$$

so that $\rho_2=1, \rho_3=1, \rho_4=4, \rho_5=11, \rho_{\nu} \geq 0$. Differentiating (4.10) twice we obtain

$$(4.11) \quad \frac{xe^x + 1}{(2 + x - e^x)^2} = \sum_{\nu=2}^{\infty} \rho_{\nu} \frac{x^{\nu-2}}{(\nu-2)!}$$

Substituting here $x=1/2$ we obtain easily that

$$\sum_{\nu=6}^{\infty} \rho_{\nu} \frac{1}{(\nu-2)!} (1/2)^{\nu-2} < 3/10,$$

whence

$$(4.12) \quad 0 \leq \rho_{\nu}/(\nu-2)! \leq (1/2) \cdot 2^{\nu-2} \quad \text{for } \nu \geq 3.$$

Comparing (4.9) with (4.10) and noting that $\sigma_k^2 < \lambda_n^2 s_n^2$ we deduce readily that

$$(4.13) \quad |\gamma_{k,\nu}| \leq (\sigma_k^2/(\lambda_n s_n)^2) \rho_{\nu} (\lambda_n s_n)^{\nu}, \quad \nu \geq 3,$$

or by (4.12)

$$(4.14) \quad (1/(\nu-2)!) |\gamma_{k,\nu}| \leq (1/2) \sigma_k^2 (2\lambda_n s_n)^{\nu-2}, \quad \nu \geq 3.$$

Accordingly we have by (2.15)

$$(4.15) \quad (1/(\nu-2)!) |\Gamma_{n,\nu}| \leq (s_n^2/2) (2\lambda_n s_n)^{\nu-2}, \quad \nu \geq 3,$$

whilst $\Gamma_{n,2} = s_n^2$; this allows us to deduce majorating expressions for P_n, \bar{M}_n and \bar{s}_n^2 (cf. (4.4)–(4.6)).

5. Application of the central limit theorem. Using (4.14) and (4.2) it follows that

$$(5.1) \quad |\bar{m}_k| \leq \sigma_k^2 h \left(1 + \frac{1}{2} \frac{\lambda_n s_n h}{1 - 2\lambda_n s_n h} \right) \leq \lambda_n^2 s_n^2 h \left(1 + \frac{1}{2} \frac{\lambda_n s_n h}{1 - 2\lambda_n s_n h} \right);$$

observing that $\Gamma_{n,2} = s_n^2$ it follows similarly from (4.6) that

$$(5.2) \quad |\bar{s}_n^2 - s_n^2| \leq \sum_{\nu=3}^{\infty} |\Gamma_{n,\nu}| \frac{h^{\nu-2}}{(\nu-2)!} \leq s_n^2 \frac{\lambda_n s_n h}{1 - 2\lambda_n s_n h}.$$

Now, by definition, $\bar{V}_k(x)$ ($k=1, \dots, n$) varies only within the interval $|x| < \lambda_n s_n$. Accordingly, if we restrict $\lambda_n s_n h$ to an interval $0 < \lambda_n s_n h < c < 1/2$, $\bar{V}_k(x)$ will not vary outside of an interval centered at \bar{m}_k and of length $O(\lambda_n s_n)$. It follows (A. C. Berry [1, Theorem 4]) that (cf. (1.6))

$$|\bar{F}_n(\bar{M}_n + x\bar{s}_n) - \Phi(x)| = O(\lambda_n).$$

More precisely let us suppose that, say,

$$(5.3) \quad 0 < \lambda_n s_n h < 1/10.$$

Then

$$|s_n^2/s_n^2 - 1| < 1/8, \quad \text{and} \quad |\bar{m}_k| < (17/160)\lambda_n s_n.$$

Therefore for $k=1, \dots, n$

$$\begin{aligned} \int_{|x| > (8/5)\lambda_n \bar{s}_n} d\bar{V}_k(x + \bar{m}_k) &= \frac{1}{p_k} \int_{|x - \bar{m}_k| > (8/5)\lambda_n \bar{s}_n} e^{hx} d\bar{V}_k(x) \\ &\leq \frac{1}{p_k} \int_{|x| > \lambda_n s_n} e^{hx} dV_k(x) = 0. \end{aligned}$$

Accordingly, using Berry's theorem, we have

$$(5.4) \quad \bar{F}_n(\bar{M}_n + x\bar{s}_n) = \Phi(x) + R_n(x)$$

with

$$(5.5) \quad |R_n(x)| < (13/3)\lambda_n.$$

Introducing (5.4) into (3.13) we obtain

$$(5.6) \quad 1 - F_n(\bar{M}_n) = e^{P_n - h\bar{M}_n} \left\{ (2\pi)^{-1/2} \int_0^\infty e^{-h\bar{s}_n y - (1/2)v^2} dy + \int_0^\infty e^{-h\bar{s}_n y} dR_n(y) \right\}.$$

Now, using (5.5) we obtain

$$\left| \int_0^\infty e^{-h\bar{s}_n y} dR_n(y) \right| = \left| -R_n(0) + h\bar{s}_n \int_0^\infty e^{-h\bar{s}_n y} R_n(y) dy \right| \leq (26/3)\lambda_n$$

(which result can be derived also by means of the second mean value theorem). The last inequality together with (5.6) shows that

$$(5.7) \quad 1 - F_n(\bar{M}_n) = e^{P_n - h\bar{M}_n} \{ e^{(1/2)(h\bar{s}_n)^2} (1 - \Phi(h\bar{s}_n)) + \theta_1 \lambda_n \},$$

where

$$(5.8) \quad |\theta_1| < 26/3.$$

We now have to determine h so that

$$(5.9) \quad \bar{M}_n \equiv \sum_{\nu=2}^\infty \Gamma_{n,\nu} \frac{h^{\nu-1}}{(\nu-1)!} = x s_n.$$

We write the inverse series formally,

$$(5.10) \quad h = \frac{1}{s_n} \sum_{\nu=1}^\infty c_{n,\nu} x^\nu = \frac{1}{s_n} \left(x - \frac{x^2}{2s_n^3} \sum_{k=1}^n b_{k,3} + \dots \right).$$

Using Cauchy's principle for majorating series and the estimate (4.15) it is

readily seen that the formal inversion of (5.9) is possible at least for all values of x for which the inversion of

$$(5.11) \quad x = s_n h - \frac{1}{4} \sum_{\nu=3}^{\infty} (2\lambda_n)^{\nu-2} (s_n h)^{\nu-1}$$

is possible. The substitution

$$(5.12) \quad \lambda_r x = u, \quad \lambda_n s_n h = z$$

reduces (5.11) to

$$(5.13) \quad u = z - \frac{1}{8} \sum_{\nu=2}^{\infty} (2z)^{\nu} = z - \frac{1}{2} \frac{z^2}{1 - z^2};$$

the inverse function

$$(5.14) \quad z = \frac{1 + 2u - |1 - 6u + 4u^2|^{1/2}}{5} = u + \sum_{\nu=2}^{\infty} c_{\nu}^* u^{\nu}$$

is regular in the interval (2.2). Using an inductive argument we readily deduce from the series-expansion in (5.13) that

$$\left. \frac{d^{\nu} z}{d u^{\nu}} \right|_{u=0} \geq 0$$

for all ν . Hence $c_{\nu}^* \geq 0$; we have in particular $c_2^* = 1/2$ and $c_3^* = 3/2$. For $u = 1/6$ we obtain $z = 1/5$ and hence

$$\sum_{\nu=4}^{\infty} c_{\nu}^* 6^{-\nu} = 1/80;$$

accordingly we have for $\nu \geq 2$

$$0 \leq c_{\nu}^* \leq (1/12)6^{\nu-1}.$$

Returning to the original variables x and h it follows that for $\nu \geq 2$

$$(5.15) \quad |c_{n,\nu}| \leq (1/12)(6\lambda_n)^{\nu-1}$$

and

$$(5.16) \quad s_n h - x < \frac{x}{12} \frac{6\lambda_n x}{1 - 6\lambda_n x}.$$

We note in particular that the restriction of x to the interval (2.5) implies that

$$(5.17) \quad \lambda_n s_n h \leq 13/144 < 1/10$$

so that (5.3) holds; this justifies the numerical estimate of §4.

For $u = (3 - 5^{1/2})/4$ we obtain from (5.14) $z = (1/2)(1 - 5^{-1/2}) < 1/2$. Hence

to values of x in the interval (2.2) there correspond values of $h < 1/(2\lambda_n s_n)$; by (4.15) these values are within the interval of convergence of the power series

$$(5.18) \quad h\bar{M}_n - P_n = \sum_{\nu=2}^{\infty} \Gamma_{n,\nu} \frac{\nu - 1}{\nu!} h^\nu.$$

Hence we can write

$$(5.19) \quad h\bar{M}_n - P_n = \frac{x^2}{2} + \frac{x^2}{2} Q_n(x) = \frac{x^2}{2} + \frac{x^2}{2} \sum_{\nu=1}^{\infty} q_{n,\nu} x^\nu,$$

where the last series converges at least within (2.2). Now define a new power series

$$(5.20) \quad Q_n^*(x) = \sum_{\nu=2}^{\infty} q_{n,\nu}^* x^\nu$$

by substituting

$$(5.21) \quad s_n^* h^* = x + \frac{x}{12} \frac{6\lambda_n x}{1 - 6\lambda_n x}$$

into

$$(5.22) \quad \frac{x^2}{2} + \frac{x^2}{2} Q_n^*(x) = \frac{(s_n^* h^*)^2}{2} + \frac{s_n^{*2} h^{*2}}{6} \sum_{\nu=3}^{\infty} (2\lambda_n s_n^* h^*)^{\nu-2}.$$

It has been shown that the series (5.21) and (5.22) are dominating series for $s_n h$ and for $h\bar{M}_n - P_n$, respectively. Hence (5.22) dominates the series (5.19) so that $|q_{n,\nu}| \leq q_{n,\nu}^*$. However, an easy computation shows that $1 + Q_n^*(1/12\lambda_n) < 1.26$ and that $q_{n,1}^* = 5\lambda_n/3$; it follows that $q_{n,\nu}^* < (1/7)(12\lambda_n)^\nu$, which proves that $Q_n(x)$ has the properties stated in Theorem 1.

Now by (5.7), (5.9) and (5.19)

$$(5.23) \quad 1 - F_n(x s_n) = e^{(1/2)x^2(1+Q_n(x))} \{ e^{(1/2)(h\bar{s}_n)^2} (1 - \Phi(h\bar{s}_n)) + \theta_1 \lambda_n \}.$$

It remains to show that here $h\bar{s}_n$ can be replaced by $x = \bar{M}_n/s_n$. Comparing the power series for $h\bar{s}_n$ and \bar{M}_n/s_n it is seen that the terms of first and second order in h are the same; hence $h\bar{s}_n = x + O(x^3 \lambda_n^2)$; more precisely we have

$$(5.24) \quad |h\bar{s}_n - x| < 2.45 \lambda_n^2 x^3,$$

provided only that (2.5) holds. Now let for the moment

$$(5.25) \quad f(t) = e^{(1/2)t^2} \{ 1 - \Phi(t) \}.$$

Then

$$f'(t) = -(2\pi)^{-1/2} + te^{(1/2)t^2} \{1 - \Phi(t)\},$$

and a well known asymptotic expansion shows that

$$0 \leq -f'(t) \leq 1/(2\pi)^{1/2}t^2.$$

Accordingly by (5.24)

$$|f(h\bar{s}_n) - f(x)| < \frac{2.45}{(2\pi)^{1/2}} \frac{\lambda_n^2 x^3}{(x - 2.6\lambda_n^2 x^3)^2} \leq \lambda_n^2 x.$$

Thus we have by (5.23) and (5.25)

$$1 - F_n(xS_n) = e^{-(1/2)x^2(1+Q_n(x))} \{f(x) + \theta_1\lambda_n + \vartheta\lambda_n\},$$

where $|\vartheta| < \lambda_n x$ and $|\theta_1| < 26/3$ (cf. (5.8)).

This proves the main part of Theorem 1. It remains only to prove (2.8). Now it follows from (2.15) and (2.16) that

$$x^2(Q_j(x) - Q_i(x)) = \sum_{\nu=2}^{\infty} \frac{h^\nu}{(\nu-2)! \nu} \sum_{k=i+1}^j \gamma_{k,\nu}.$$

Accordingly, using (4.14),

$$\begin{aligned} x^2 |Q_j(x) - Q_i(x)| &\leq \frac{h^2}{4} \sum_{\nu=2}^{\infty} 2(\lambda_n s_n h)^{\nu-2} \sum_{k=i+1}^j \sigma_k^2 \\ &= \frac{h^2}{4} (s_j^2 - s_i^2) \frac{1}{1 - 2\lambda_n s_n h}. \end{aligned}$$

This implies (2.8) since $s_n h/x < 13/12$ and $\lambda_n s_n h < 1/10$ by (5.16) and (5.17), respectively.

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