

# SCATTERING FROM RANDOM LINEAR ARRAYS WITH CLOSEST APPROACH\*

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1. Probability distributions for the echo returned from a random linear scattering array are of some importance in several different situations, especially in connection with the use of radar and sonar. In the simplest case  $n$  identical isotropic point scatterers are placed at random on the interval  $[0, L]$ , the source of radiation is at the point  $(-s, 0)$  which is sufficiently far from the origin so that the ratio  $(s + x_1)/(s + x_2)$  is sufficiently close to 1 for any two points  $x_1$  and  $x_2$  on  $[0, L]$ , and the length  $L$  is an integral multiple of the wave-length  $\lambda$ . The problem of finding the probability density for the components of the scattered signal reduces then to finding the joint probability density  $W_n(X, Y)$  for the components of the vector

$$\phi = \left( \sum_{i=1}^n \cos(2\pi x_i/\lambda), \sum_{i=1}^n \sin(2\pi x_i/\lambda) \right),$$

where each variable  $x_i$  is taken from the uniform rectangular probability distribution:  $\Pr(x_i \leq x) = x/L, 0 \leq x \leq L$ . The probability density for the amplitude of the scattered signal is obtained in terms of the probability density  $W_n(R)$  for the quantity  $R = (X^2 + Y^2)^{1/2}$ . It follows that aside from some normalization factor the problem is formally equivalent to the classical isotropic plane random walk.

If the scatterers are spheres of radius  $r$ , it may still be convenient to regard them as points, but with the additional restriction that no two points are closer than  $2r$ . This assumption also fits the case where each scatterer has a radius of repulsion, within which no other scatterer may enter. We shall consider here the corresponding scattering problem: to find the probabilities  $W_n(X, Y)$  and  $W_n(R)$  under the restriction that the  $n$  point scatterers on  $[0, L]$  are not allowed to be closer to each other than  $a, 0 \leq a \leq L/(n - 1)$ , and are otherwise at random. Our approach will be mainly geometrical, and we shall begin by finding the sample space of the configurations.

2. We start with a well known problem in elementary geometrical probability: on the segment  $[0, L]$   $n$  points are taken at random, given a number  $a, 0 \leq a \leq L/(n - 1)$ , what is the probability  $P(n, a, L)$  that no two points are closer than  $a$ ? If  $n = 2$ , then the sample space of pairs of points  $(x_1, x_2), 0 \leq x_1, x_2 \leq L$ , is the square of side  $L$ ; let  $D$  be its diagonal through the origin, and draw the two lines parallel to  $D$  at the distance  $2^{-1/2} a$  from it. The hexagonal subset of the square, contained between these two lines, is then the sample space of the forbidden configurations  $(x_1, x_2)$  with  $|x_1 - x_2| \leq a$ ; the remainder of the square consists of two triangles which can be moved together to form a square of side length  $L - a$ . Hence, by the randomness assumption,  $P(2, a, L) = (L - a)^2/L^2 = (1 - a/L)^2$ .

The general case can be handled in the same way. In the  $n$ -dimensional Euclidean space  $E^n$  we consider a Cartesian coordinate system with the  $n$  axes  $X_1, \dots, X_n$ . Let  $H$  be the hypercube

$$H = \{(x_1, \dots, x_n) : 0 \leq x_i \leq L, i = 1, \dots, n\};$$

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$H$  is the sample space of all  $n$ -tuples of points on the segment  $[0, L]$ . Let  $I_i$  be the interval  $[0, L]$  on the  $X_i$ -axis, and in the two-dimensional square face  $S_{ii} = I_i \times I_i$  let  $D_{ii}$  be the diagonal through the origin. Let  $B_{ii}$  be the hexagonal subset of  $S_{ii}$ , consisting of all points no further from  $D_{ii}$  than  $2^{-1/2}a$ . Let  $F_{ii}$  be the Cartesian product of  $B_{ii}$  and all  $I_k, k \neq i, j$ . That is,

$$F_{ii} = \{(x_1, \dots, x_n) : (x_i, x_j) \in B_{ii}, 0 \leq x_k \leq L, k \neq i, j\}.$$

Then  $F_{ii}$  is the sample space of those  $n$ -tuples  $(x_1, \dots, x_n)$  for which  $|x_i - x_j| \leq a$ . Hence the sample space of the allowed configurations, that is, of all  $n$ -tuples  $(x_1, \dots, x_n)$  such that  $|x_i - x_j| \geq a$  for all  $i$  and  $j, i \neq j$ , is the set

$$H - \bigcup_{1 \leq i < j \leq n} F_{ij}.$$

Therefore by the randomness assumption

$$P(n, a, L) = \text{vol} (H - \bigcup_{1 \leq i < j \leq n} F_{ij}) / \text{vol} H. \tag{1}$$

When the  $\binom{n}{2}$  sets  $F_{ij}$  are removed from  $H$ , the remainder  $H - \bigcup_{1 \leq i < j \leq n} F_{ij}$  consists of  $n!$  congruent simplexes. These can be assembled by a sequence of translations so as to form an  $n$ -dimensional hypercube  $H'$  of edge-length  $L - (n - 1)a$ . The required translations are as follows: for  $i = 1, \dots, n$  let  $\xi_i$  denote the vector in  $E^n$  whose  $i$ -th coordinate is  $-a$  and all others are 0, now translate by  $\xi_i$  (by  $\xi_i$ ) every one of the  $n!/2$  simplexes, which lies on the same side of  $F_{ij}$  as the positive half of the  $X_i$ -axis (the  $X_j$ -axis), and carry out this operation for all  $i$  and  $j, 1 \leq i < j \leq n$ . We have now by (1)  $P(n, a, L) = \text{vol} H' / \text{vol} H$ , and therefore

$$P(n, a, L) = [1 - (n - 1)a/L]^n, \tag{2}$$

which is the well known solution of the problem, [1], [2]. An equivalent way of expressing it is the following: let  $n$  points  $x_1, \dots, x_n$  be placed at random on the interval  $[0, L]$ , let  $u = \min_{i < j} |x_i - x_j|$  be the nearest approach of any two of the  $n$  points, and let  $P_n(u, L)$  be the probability density for  $u$ ; then

$$P(n, a, L) = \int_a^{L/(n-1)} P_n(u, L) du,$$

and so by (2) we obtain

$$\begin{aligned} P_n(u, L) &= [n(n - 1)/L][1 - (n - 1)u/L]^{n-1}, & 0 \leq u \leq L/(n - 1), \\ &= 0 & L/(n - 1) < u. \end{aligned} \tag{3}$$

The  $k$ -th moment of  $u$  is

$$\mu_{n,k} = \int_0^{L/(n-1)} u^k P_n(u, L) du = \left[ \binom{n+k}{k} \right]^{-1} [L/(n-1)]^k,$$

the mean is  $L/(n^2 - 1)$  and the standard deviation is  $[L/(n^2 - 1)][n/(n + 2)]^{1/2}$ . It may be observed that if  $n \rightarrow \infty$  and  $a \rightarrow 0$  so that  $n^2 a \rightarrow \alpha, 0 < \alpha < \infty$ , then

$$\lim P(n, a, L) = P(\alpha, L) = \exp(-\alpha/L);$$

and if  $n \rightarrow \infty$  and  $L \rightarrow \infty$  so that  $L/n^2 \rightarrow \beta$ ,  $0 < \beta < \infty$ , then

$$\lim P(n, a, L) = P(\beta, a) = \exp(-a/\beta), \lim \mu_{n,k} = k! \beta^k.$$

We observe further that if  $a \ll L$  then (2) may be written as

$$P(n, a, L) = 1 - 2\binom{n}{2}a/L + O(a^2/L). \tag{4}$$

We consider next the analogous problem for  $k$ -tuple clumping:  $n$  points are taken at random on the interval  $[0, L]$ , an integer  $k$  and a number  $a$  are given, where  $2 \leq k \leq n$  and  $0 < a < L$ , what is the probability  $P(k, n, a, L)$  that no  $k$  points lie on an interval of length  $a$ ? Equivalently, if a segment of length  $a$  is dragged along  $[0, L]$ ,  $P(k, n, a, L)$  is the probability that this segment never covers more than  $k - 1$  points at a time. Again, if  $n$  events are to occur at random times during the time interval  $[0, T]$   $P(k, n, \tau, T)$  is the probability that no  $k$  events occur within a time interval of length  $\tau$ . This problem is more difficult than the previous one, but by a similar geometrical reasoning it can be shown that if  $a \ll L$  then

$$P(k, n, a, L) = 1 - k\binom{n}{k}(a/L)^{k-1} + O(a^k/L^k). \tag{5}$$

An outline of the proof follows. Let  $H$  be the same hypercube as before, and let the sets  $I_1, \dots, I_n$  be as previously defined. Let  $i_1, i_2, \dots, i_k$  be  $k$  integers, such that

$1 \leq i_1 < i_2 < \dots < i_k$ , then  $S_{i_1 i_2 \dots i_k} = I_{i_1} \times I_{i_2} \times \dots \times I_{i_k}$  is one of the  $\binom{n}{k}$   $k$ -

dimensional faces of  $H$ , which contain the origin. Let  $D_{i_1 i_2 \dots i_k}$  be the longest diagonal of the  $k$ -dimensional hypercube  $S_{i_1 i_2 \dots i_k}$ , passing through the origin. Let  $K(p)$  denote the  $k$ -dimensional hypercube of edge-length  $a$ , whose center is the point  $p$ , and whose edges are parallel to the coordinate axes  $X_{i_1}, X_{i_2}, \dots, X_{i_k}$ . In analogy to the previously considered sets  $B_{ij}$ , define

$$B_{i_1 i_2 \dots i_k} = (S_{i_1 i_2 \dots i_k}) \cap \left( \bigcup_{p \in D_{i_1 i_2 \dots i_k}} K(p) \right),$$

and let  $F_{i_1 i_2 \dots i_k}$  be the Cartesian product of  $B_{i_1 i_2 \dots i_k}$  with all the  $I_j$ 's,  $j \neq i_1, i_2, \dots, i_k$ . That is,

$$F_{i_1 i_2 \dots i_k} = \{(x_1, x_2, \dots, x_n) : (x_{i_1}, x_{i_2}, \dots, x_{i_k}) \in B_{i_1 i_2 \dots i_k}, 0 \leq x_j \leq L, j \neq i_1, i_2, \dots, i_k\}.$$

Then  $F_{i_1 i_2 \dots i_k}$  is the sample space of all  $n$ -tuples  $(x_1, \dots, x_n)$ , for which the points corresponding to  $x_{i_1}, x_{i_2}, \dots, x_{i_k}$  are in the forbidden configuration, that is, are covered by a segment of length  $\leq a$ . Hence by the randomness assumption

$$P(k, n, a, L) = \text{vol}(H - \bigcup F_{i_1 i_2 \dots i_k}) / \text{vol } H,$$

where the union is taken over all the  $\binom{n}{k}$  selections of the integers  $i_1, i_2, \dots, i_k$ . By using the inclusion-exclusion principle, [3], and some elementary volume estimates, it may be shown that

$$\text{vol}(H - \bigcup F_{i_1 i_2 \dots i_k}) = \text{vol } H - \sum \text{vol } F_{i_1 i_2 \dots i_k} + O(a^k/L^k).$$

Since there are  $\binom{n}{k}$  sets  $F_{i_1, \dots, i_k}$ , and since the volume of each one is  $L^{n-k}[kLa^{k-1} + O(a^k)]$ ,

(5) follows at once.

A different generalization of the first problem is the following: let  $A = (a_{ij})$  be an  $n \times n$  real matrix with  $a_{i\sigma} = a_{ji} > 0$  for  $1 \leq i < j \leq n$ , and  $a_{ii} = 0, i = 1, \dots, n$ ,  $n$  labelled points  $x_1, \dots, x_n$  are placed at random on the segment  $[0, L]$ ; what is the probability  $P(n, A, L)$  that  $|x_i - x_j| \geq a_{ij}$  for all  $i$  and  $j$ ? Here it may be shown, [4], that if the matrix  $A$  satisfies the triangular condition

$$a_{ij} + a_{jk} \geq a_{ik}, \quad 1 \leq i, j, k \leq n,$$

then

$$P(n, A, L) = 1/n! \sum_{\sigma \in G} [\max(0, 1 - 1/L \sum_{k=1}^{n-1} a_{\sigma(k)\sigma(k+1)})]^n,$$

where  $G$  is the symmetric group on  $n$  elements, whose members are

$$\sigma = \begin{pmatrix} 1 & 2 & \dots & n \\ \sigma(1)\sigma(2) & \dots & \sigma(n) \end{pmatrix}.$$

If  $a_{ij} \ll L$  for all  $i$  and  $j$ , then in analogy to (4) and (5) we have

$$P(n, A, L) = 1 - 2/L \sum_{i < j} a_{ij} + O[\max a_{ij}/L]^2.$$

3. We take up now the scattering problem mentioned at the end of section 1. For simplicity it will be assumed throughout that  $\lambda = 2\pi$ . If the assumption concerning the closest approach of two scatterers were absent, then by the standard methods, for instance, by that in [5], we could have shown that

$$W_n(X, Y) = (2\pi)^{-2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp[-i(\rho_1 X + \rho_2 Y)] A_n(\rho_1, \rho_2) d\rho_1 d\rho_2, \quad (6)$$

where

$$A_n(\rho_1, \rho_2) = 1/L^n \int_0^L \dots \int_0^L \exp i \left( \rho_1 \sum_1^n \cos x_k + \rho_2 \sum_1^n \sin x_k \right) dx_1 \dots dx_n. \quad (7)$$

We notice that the region of integration in (7) is the hypercube  $H$  of section 2. Recalling that the sample space of the forbidden configurations  $(x_1, \dots, x_n)$ , with  $|x_i - x_j| \leq a$  for some  $i$  and  $j$ , is the set  $F = \bigcup_{i < j} F_{ij}$ , and the sample space of the allowed configurations is therefore the residual set  $R = H - F$  of the hypercube  $H$ , we have the following: under the assumption that no two  $x$ 's are closer than  $a$ , and otherwise the  $x$ 's are at random, the probability  $W_n(X, Y)$  is still given by (6) but in the integral for  $A_n(\rho_1, \rho_2)$  the region of integration is  $R$ , not  $H$ . That is,

$$W_n(X, Y) = (2\pi)^{-2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp[-i(\rho_1 X + \rho_2 Y)] B_n(\rho_1, \rho_2) d\rho_1 d\rho_2, \quad (8)$$

where

$$B_n(\rho_1, \rho_2) = (\text{vol } R)^{-1} \int_R \dots \int_R \exp i \left( \rho_1 \sum_1^n \cos x_k + \rho_2 \sum_1^n \sin x_k \right) dx_1 \dots dx_n. \quad (9)$$

Most of our work from now on will be concerned with estimating  $B_n(\rho_1, \rho_2)$ . We shall assume that  $a$  is small, and we shall derive a formula of the type

$$B_n(\rho_1, \rho_2) = B_0(\rho_1, \rho_2) + aB_1(\rho_1, \rho_2) + O(a^2); \tag{10}$$

$B_0$  and  $B_1$  will be found, and since the final aim is finding the Fourier transform (8) of  $B_n$ , it will be necessary to examine the dependence of the  $O(a^2)$  term in (10) on  $\rho_1$  and  $\rho_2$ . This will complicate somewhat the derivation of the succeeding estimates.

4. Let  $S_1, \dots, S_N$  be a finite number of sets in the Euclidean space  $E^n$ , let  $S$  be their union, and let  $f$  be an integrable function defined on  $S$ . Then

$$\int_S f dV = \sum_{i=1}^N \int_{S_i} f dV - \sum_{1 \leq i < j \leq N} \int_{S_i \cap S_j} fg_{ij} dV, \tag{11}$$

where  $g_{ij}$  is a function defined on  $S_i \cap S_j$  and  $0 \leq g_{ij} \leq 1$ , for each  $i$  and  $j$ . To prove (11), we consider the set of  $2^N$  atoms, corresponding to the sets  $S_i$ . These atoms are defined as follows: let  $\tilde{S}_i = S - S_i$ , and let  $K(S_i)$  stand for either one of the sets  $S_i$  and  $\tilde{S}_i$ ; then the  $2^N$  atoms are all the  $2^N$  sets of the form

$$K(S_1) \cap K(S_2) \cap \dots \cap K(S_N).$$

An atom which is not the empty set, is said to be of order  $k$  if it lies in exactly  $k$  different sets  $S_i$ . Define now  $g_{ij}$  as follows:  $g_{ij}(x) = 2/k$  if  $x \in A$  and  $A$  is an atom of order  $k$ .

Since such an atom lies in exactly  $\binom{k}{2}$  sets  $S_i \cap S_j$  the corresponding contribution in (11) from the first sum is  $k \int_A f dV$ , and the contribution from the second sum is  $(k - 1) \int_A f dV$ ; since  $S$  is the union of all the  $2^N$  atoms, the equation (11) balances.

It may be mentioned that a formula more general than (11) is

$$\begin{aligned} \int_S f dV = \sum_{i=1}^N \int_{S_i} f dV - \sum_{1 \leq i_1 < i_2 \leq N} \int_{S_{i_1} \cap S_{i_2}} f dV + \sum_{1 \leq i_1 < i_2 < i_3 \leq N} \int_{S_{i_1} \cap S_{i_2} \cap S_{i_3}} f dV - \dots \\ + (-1)^{r+1} \sum_{1 \leq i_1 < i_2 < \dots < i_r \leq N} \int_{S_{i_1} \cap \dots \cap S_{i_r}} fg_{i_1 i_2 \dots i_r} dV, \end{aligned}$$

where  $g = g_{i_1 i_2 \dots i_r}$  is defined on the set  $S_{i_1} \cap S_{i_2} \cap \dots \cap S_{i_r}$ , and  $0 \leq g \leq 1$ . This formula can be proved in the same way as (11); it would be the proper starting point if, instead of deriving (10), we were interested in deriving the superior approximation

$$B_n(\rho_1, \rho_2) = B_0(\rho_1, \rho_2) + aB_1(\rho_1, \rho_2) + \dots + a^r B_r(\rho_1, \rho_2) + O(a^{r+1}).$$

We proceed to apply (11) in order to derive (10). Recall that

$$R = H - F = H - \bigcup_{1 \leq i < j \leq n} F_{ij},$$

and let  $T$  stand for the integrand in (9). Then

$$\int_R T dV = \int_H T dV - \int_F T dV; \tag{12}$$

we now apply (11) with  $N = \binom{n}{2}$  and with the sets  $F_{ij}$  as the  $S_i$ 's. We obtain then

$$\int_F T dV = \sum_{1 \leq i < j \leq n} \int_{F_{ij}} T dV - \sum \int_{F_{ij} \cap F_{rs}} Tg_{ijrs} dV, \tag{13}$$

where the second summation is over all the distinct

$$\binom{n}{2} = (n^4 - 2n^3 - n^2 + 2n)/8 \tag{14}$$

pairs  $(i, j), (r, s)$  with  $1 \leq i < j \leq n$  and  $1 \leq r < s \leq n$ .

5. We shall derive in this section an estimate of the first sum in (13). By the definition of  $F_{i,j}$  as a product set, and by the multiplicative property of the integrand  $T$ , we have

$$\int_{F_{i,j}} T dV = \left[ \int_0^L \cdots \int_0^L \exp i \left( \rho_1 \sum_1^{n-2} \cos x_k + \rho_2 \sum_1^{n-2} \sin x_k \right) dx_1 \cdots dx_{n-2} \right] \cdot \left[ \int_{B_{i,j}} \exp i[\rho_1(\cos x + \cos y) + \rho_2(\sin x + \sin y)] dx dy \right], \tag{15}$$

where  $B_{i,j}$  is the hexagon in the  $xy$ -plane, with the vertices  $(0, 0), (0, a), (a, 0), (L, L), (L - a, L), (L, L - a)$ . By adding onto  $B_{i,j}$  the two triangles with the vertices  $(0, 0), (a, 0), (0, -a)$  and  $(L, L), (L - a, L), (L, L + a)$  we complete the hexagon to the parallelogram given by  $x - a \leq y \leq x + a, 0 \leq x \leq L$ . Therefore, denoting by  $J$  the second integral on the right hand side in (15), we have

$$J = \int_0^L \int_{x-a}^{x+a} \exp i[\rho_1(\cos x + \cos y) + \rho_2(\sin x + \sin y)] dy dx + c_1 a^2 \tag{16}$$

where  $|c_1| \leq 1$ . We observe that

$$\int_{x-a}^{x+a} f(y) dy = 2af(x) + c_2 M a^2,$$

where  $|c_2| \leq 1$  and  $M = \max |f'(y)|$ . Applying this to (16), with

$$f(y) = \exp i(\rho_1 \cos y + \rho_2 \sin y),$$

we have  $|f'(y)| \leq \rho$ , where

$$\rho = (\rho_1^2 + \rho_2^2)^{1/2}, \tag{17}$$

and therefore

$$J = \int_0^L [2a \exp 2i(\rho_1 \cos x + \rho_2 \sin x) + c_2 \rho a^2 \exp i(\rho_1 \cos x + \rho_2 \sin x)] dx + c_1 a^2.$$

It was assumed that  $L$  is an integral multiple of the wave-length  $\lambda = 2\pi$ , therefore

$$J = 2aLJ_0(2\rho) + c_2 \rho a^2 L J_0(\rho) + c_1 a^2, \tag{18}$$

where  $J_0(x)$  is the Bessel function of the first kind and order 0, and  $\rho$  is given by (17). The  $(n - 2)$ -fold integral in (15) is likewise expressed in terms of the Bessel functions:

$$\int_0^L \cdots \int_0^L \exp i \left( \rho_1 \sum_1^{n-2} \cos x_k + \rho_2 \sum_1^{n-2} \sin x_k \right) dx_1 \cdots dx_{n-2} = [LJ_0(\rho)]^{n-2}.$$

This, together with (18), gives an estimate on the first sum in (13):

$$\sum_{1 \leq i < j \leq n} \int_{F_{i,j}} T dV = \binom{n}{2} \cdot \int_{F_{i,j}} T dV = \binom{n}{2} a L^{n-2} [J_0(\rho)]^{n-2} [2LJ_0(2\rho) + c_2 a \rho L J_0(\rho) + c_1 a]. \tag{19}$$

6. We consider next the terms of the second sum in (13). These are of two kinds: (a)  $N_1$  terms coming from the pairs  $F_{ij}, F_{rs}$  with all four indices distinct, and (b)  $N_2$  terms coming from the pairs  $F_{ij}, F_{rs}$  where two indices coincide, that is,  $i = r$  or  $j = s$  or  $j = r$ .

For a term  $K$  of the first kind we have, by the mean value theorem for integrals,

$$K = \int_{F_{ij} \cap F_{rs}} T g_{ijrs} dV = c_3 \int_{F_{ij} \cap F_{rs}} T dV$$

where  $|c_3| \leq 1$ , since  $0 \leq g_{ijrs} \leq 1$ . Now, by the definition of the sets  $F_{ij}$  and by the multiplicative property of the integrand  $T$ , we have

$$K = c_3 [LJ_0(\rho)]^{n-4} J^2,$$

with  $J$  given by (18). Therefore

$$K = c_3 a^2 [LJ_0(\rho)]^{n-4} [2LJ_0(2\rho) + c_2 \rho a LJ_0(\rho) + c_1 a]^2. \tag{20}$$

A typical term of the second kind is

$$\int_{F_{ij} \cap F_{is}} T g_{ijis} dV;$$

here by a method similar to that one used in section 5, we derive the estimate

$$\int_{F_{ij} \cap F_{is}} T g_{ijis} dV = c_4 [LJ_0(\rho)]^{n-3} [a(\rho + a)]^2, \tag{21}$$

where  $c_4$  is a constant.

Putting together the results of this section and the previous one, in particular, by (19), (20) and (21), we get from (12)

$$\begin{aligned} \int_R T dV &= \int_H T dV - \int_F T dV = [LJ_0(\rho)]^n - \binom{n}{2} a [LJ_0(\rho)]^{n-2} [2LJ_0(2\rho) \\ &+ c_2 a \rho LJ_0(\rho) + c_1 a] + N_1 c_3 a^2 [LJ_0(\rho)]^{n-4} [2LJ_0(2\rho) + c_2 a \rho LJ_0(\rho) + c_1 a]^2 \end{aligned} \tag{22}$$

Actually, we do not need every term explicitly, and we write (22) as

$$\int_R T dV = [LJ_0(\rho)]^n - n(n-1) a L^{n-1} J_0^{n-2}(\rho) J_0(2\rho) + ca^2 E, \tag{23}$$

where  $c$  is a constant, and the error term  $E$  is a finite sum of expressions like

$$\rho^j J_0^{n-k}(\rho) J_0^p(2\rho), \quad 0 \leq j \leq 2, \quad 0 \leq k \leq 4, \quad 0 \leq p \leq 2.$$

Hence (9) becomes

$$B_n(\rho_1, \rho_2) = [L - (n-1)a/L]^{-n} [L^n J_0^n(\rho) - n(n-1) a L^{n-1} J_0^{n-2}(\rho) J_0(2\rho) + ca^2 E],$$

and expanding the negative power, we get

$$B_n(\rho_1, \rho_2) = J_0^n(\rho) + n(n-1)(a/L) J_0^{n-2}(\rho) [J_0^2(\rho) - J_0(2\rho)] + c'(a/L)^2 E', \tag{24}$$

where  $c'$  is a constant and  $E'$  is of the same type as  $E$ . This is the expression of the type (10) that we have been looking for.

7. In taking the inverse Fourier transform (8) of  $B_n(\rho_1, \rho_2)$  we observe that the contribution from the term  $c'(a/L)^2 E'$  is  $O(a^2/L^2)$ , since for large  $\rho$  we have  $J_0^n(\rho) =$

$O(\rho^{-m/2})$ , and since each term in  $E'$  involves a sufficiently high power of  $J_0(\rho)$  to compensate for the presence of the powers of  $\rho$  (a tacit assumption is made here that  $n$  is sufficiently large). Hence by (8)

$$W_n(X, Y) = (2\pi)^{-2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp [-i(\rho_1 X + \rho_2 Y)] J_0^n(\rho) d\rho_1 d\rho_2 + (2\pi)^{-2}(a/L)n(n-1) \cdot \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp [-i(\rho_1 X + \rho_2 Y)] J_0^{n-2}(\rho) [J_0^2(\rho) - J_0(2\rho)] d\rho_1 d\rho_2 + O(a^2/L^2). \tag{25}$$

We introduce the polar coordinates:

$$X = R \cos \varphi, \quad Y = R \sin \varphi, \quad \rho_1 = \rho \cos \alpha, \quad \rho_2 = \rho \sin \alpha,$$

and (25) becomes

$$W_n(R) = W_n(R \cos \varphi, R \sin \varphi) = (2\pi)^{-2} \cdot \int_0^{\infty} \int_0^{2\pi} \exp [-i\rho R \cos(\varphi - \alpha)] J_0^n(\rho) \rho d\alpha d\rho + (2\pi)^{-2}(a/L)n(n-1) \cdot \int_0^{\infty} \int_0^{2\pi} \exp [-i\rho R \cos(\varphi - \alpha)] J_0^{n-2}(\rho) [J_0^2(\rho) - J_0(2\rho)] \rho d\alpha d\rho + O(a^2/L^2);$$

the integration over  $\alpha$  can be carried out, and we get finally

$$W_n(R) = (2\pi)^{-1} \int_0^{\infty} J_0(\rho R) J_0^n(\rho) \rho d\rho + n(n-1)(a/2\pi L) \cdot \int_0^{\infty} J_0(\rho R) J_0^{n-2}(\rho) [J_0^2(\rho) - J_0(2\rho)] \rho d\rho + O(a^2/L^2). \tag{26}$$

For  $a = 0$  this reduces, as it should, to the classical solution of the plane isotropic random walk, obtained by Kluyver, [6].

8. In conclusion, we shall formulate a general problem, which includes as special cases all those considered here and many more besides, for instance, the case of the so called hard-sphere gas model in the statistical mechanics. Let  $n$  and  $N$  be two positive integers and let  $B$  and  $P$  be two sets in the Euclidean space  $E^n$ .  $P$  is assumed to have a center of symmetry  $x$ , and it may be thought of as being much smaller than  $B$ .  $N$  points  $x_1, \dots, x_N$  are taken at random in  $B$ , and  $P(x_i)$  is the set obtained by translating  $P$  so that the center of symmetry is at  $x_i$ . The sample space of the points  $x_1, \dots, x_N$ , or of the  $N$ -tuple  $(x_1, \dots, x_N)$ , is the  $N$ -fold Cartesian product  $\{(x_1, \dots, x_N): x_i \in B, i = 1, \dots, N\}$ . We call an  $N$ -tuple a forbidden configuration if  $x_i \in P(x_j)$  for some two members  $x_i, x_j$  of the  $N$ -tuple, otherwise it is called an allowed configuration. The sample space  $A$  of all allowed configurations is then

$$A = \{(x_1, \dots, x_N): x_i \notin P(x_j), 1 \leq i < j \leq N\}.$$

Let  $f(x)$  be an integrable function defined over  $B$ ; our problem is that of evaluating the integral

$$I = \int_A \dots \int_A f(x_1) \dots f(x_N) dx_1 \dots dx_N,$$

where  $dx_i$  is the volume element in  $E^n$ . The problem may be further extended by taking



integrands of the type  $F(x_1, \dots, x_N)$ , and also by letting the points  $x_1, \dots, x_N$  be distributed in  $B$  according to some probability law. In the latter case the integral  $I$  depends on that law, and we may ask, for instance, for its expected value.

While a general solution of this problem does not appear feasible, a possible approach might be to introduce a 'basic ratio', such as, for instance, the ratio of the diameters, widths or volumes, of the sets  $P$  and  $B$ , and to find the coefficients in the expansion of  $I$  in the powers of that ratio.

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