ON THE FREQUENCY OF SMALL FRACTIONAL PARTS IN CERTAIN REAL SEQUENCES

BY

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1. Introduction. Let $X_1, X_2, \dots$ be a sequence of independent random variables, each uniformly distributed on $[0, 1/2]$. If f is an arbitrary function from the positive integers to $[0, 1/2]$, the equation

$$(1) \quad \Pr \{X_k < f(k)\} = 2f(k)$$

holds, and it is a consequence of the Borel-Cantelli lemmas [3] that the probability that the inequality $X_k < f(k)$ is satisfied for infinitely many k is zero or one, according as the series

$$(2) \quad \sum_{k=1}^{\infty} f(k)$$

is convergent or divergent. While it is well known that no such general assertion can be made when the X_k are dependent, Khinchin [6] has found a direct analogue in an important case. His theorem is usually stated in measure-theoretic language: the inequality $|kx - p| < f(k)$ has infinitely many integral solutions k, p for almost all x or almost no x , according as (2) diverges or converges. We may, however, consider x as a random variable uniformly distributed over some interval, and define the quantity U_k ($k=1, 2, \dots$) as the distance $\langle kx \rangle$ between kx and the nearest integer to kx . Then the U_k form a sequence of dependent random variables uniformly distributed on $[0, 1/2]$; Khinchin's theorem shows that the nature of the dependence is not such as to affect the finiteness of the number of solutions of the inequality $U_k < f(k)$.

From a probabilistic standpoint the Borel-Cantelli lemmas yield very crude information about a sequence of random variables, and it is of some interest to know whether the U_k also resemble the X_k in their finer structure. We consider here the case in which (2) diverges, so that there are almost surely infinitely many solutions of $|kx - p| < f(k)$, and investigate in §§2-3 the number T_n of such solutions with $k \leq n$. The result is not quite what would be expected from the case of independent variables. For if we put Y_k equal to 1 or 0 according as the inequality $X_k < f(k)$ does or does not hold, then $S_n = Y_1 + \dots + Y_n$ is the number of $k \leq n$ such that $X_k < f(k)$. Since

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$$E(Y_k) = 1 \cdot 2f(k) + 0 \cdot (1 - 2f(k)) = 2f(k),$$

$$\text{Var } Y_k = E(Y_k^2) - E^2(Y_k) = 2f(k) - 4f^2(k),$$

$$E(S_n) = 2 \sum_{k=1}^n f(k),$$

$$\text{Var } S_n = 2 \sum_{k=1}^n f(k) - 4 \sum_{k=1}^n f^2(k),$$

we deduce from the central limit theorem that if $\sum_1^\infty f^2(k)$ converges, then

$$(3) \quad \lim_{n \rightarrow \infty} \Pr \left\{ S_n < 2 \sum_{k=1}^n f(k) + \omega \left(2 \sum_{k=1}^n f(k) \right)^{1/2} \right\} = \phi(\omega),$$

where

$$\phi(\omega) = \frac{1}{(2\pi)^{1/2}} \int_{-\infty}^{\omega} e^{-u^2/2} du$$

is the normal distribution function.

The law of the iterated logarithm yields the closely related result that

$$\Pr \left\{ \limsup_{n \rightarrow \infty} \left| \frac{S_n - 2 \sum_{k=1}^n f(k)}{4 \left(\sum_{k=1}^n f(k) \log \log \sum_{k=1}^n f(k) \right)^{1/2}} \right| = 1 \right\} = 1$$

and so in particular

$$(4) \quad \Pr \left\{ S_n \sim 2 \sum_{k=1}^n f(k) \right\} = 1.$$

Theorem 1 exhibits the result corresponding to (3) for T_n ; it differs from (3) in that the coefficient 2 is replaced by $12\pi^{-2}$.

In §§4-6 we consider the much less strongly dependent sequence $\langle r_1 r_2 \cdots r_k x \rangle$, where $r_1, r_2, \cdots$ is a fixed increasing sequence of positive integers, and show that here the situation is again as described in (3) and (4).

2. A lemma. Let f be a function with the following properties:

$$(5) \quad f(x) \text{ is positive and decreasing for } x \geq 0;$$

$$(6) \quad f(x) = O(x^{-1}) \text{ and } f'(x) = O(x^{-2}) \text{ as } x \rightarrow \infty;$$

$$(7) \quad \sum_{k=1}^{\infty} f(k) = \infty.$$

We shall need some further properties of f , which we collect in the following lemma.

LEMMA 1. *If f satisfies (5)–(7) and if c and δ are positive constants, then*

- (a)
$$\sum_{k=1}^n f(k) = \int_1^n f(u) du + O(1);$$
- (b)
$$f(k + O(k^{1-\delta})) = f(k) + O(k^{-1-\delta});$$
- (c)
$$\sum_{k=1}^{cn} f(k) = \sum_{k=1}^n f(k) + O(1);$$
- (d)
$$\sum_{k=1}^n cf(ck) = \sum_{k=1}^n f(k) + O(1);$$
- (e)
$$\sum_{k=1}^n f(k) = c \sum_{k=1}^{e^n} \frac{f(c \log k)}{k} + O(1),$$
- (f) *if $a_1, a_2, \dots$ and α are such that*

$$\sum_{k=1}^n a_k \sim n\alpha$$

as $n \rightarrow \infty$, then

$$\sum_{k=1}^n a_k f(k) = \alpha \sum_{k=1}^n f(k) + O(1).$$

Part (a) is trivial, and (b) follows from (6) and the law of the mean. Part (c) follows from the estimate

$$\sum_{k=n}^{cn} f(k) = \sum_{k=n}^{cn} O(k^{-1}) = O(\log cn - \log n) = O(1),$$

and (d) from the fact that

$$\sum_{k=1}^n cf(ck) = \int_1^n cf(cu) du + O(1) = \int_c^{cn} f(t) dt + O(1) = \sum_{k=c}^{cn} f(k) + O(1).$$

The substitution $u = c \log v$ in (a) gives (e). To obtain (f), write

$$\sum_{k=1}^n (a_k - \alpha) f(k) = f(n) \sum_{k=1}^n (a_k - \alpha) + \sum_{k=1}^{n-1} \left(\sum_{l=1}^k (a_l - \alpha) \right) (f(k) - f(k+1))$$

and note that

$$f(n) \sum_{k=1}^n (a_k - \alpha) = O(n^{-1})o(n) = o(1)$$

and

$$\begin{aligned} \sum_{k=1}^{n-1} \left(\sum_{l=1}^k (a_l - \alpha) \right) (f(k) - f(k+1)) &= \sum_{k=1}^{n-1} o(k) (f(k) - f(k+1)) \\ &= O(n) \sum_{k=1}^{n-1} (f(k) - f(k+1)) \\ &= O(nf(n)) = O(1). \end{aligned}$$

We shall use the following notation: $\mathfrak{M}\{A\}$ means the measure of the set of $x \in [0, 1]$ such that A , if A is a sentence, and it means the measure of A if A is a set.

$\text{No}\{m \leq n \mid \dots\}$ means the number of positive integers $m \leq n$ such that $\dots$.

$E_x\{\dots\}$ or $\{x \mid \dots\}$ means the set of $x \in [0, 1]$ such that $\dots$.

3. **The fractional part of mx .** We prove the following theorem:

THEOREM 1. *Suppose that f satisfies conditions (5)–(7) and put*

$$g(x) = f(\log x)/x.$$

Let

$$T_n = T_n(x) = \text{No}\{m \leq n \mid \langle mx \rangle < g(m)\}.$$

Then for fixed ω ,

$$\lim_{n \rightarrow \infty} \mathfrak{M} \left\{ T_n < \frac{12}{\pi^2} \sum_{k=1}^n g(k) + \omega \left(\frac{12}{\pi^2} \sum_{k=1}^n g(k) \right)^{1/2} \right\} = \phi(\omega).$$

If x is a real number with continued fraction expansion

$$x = a_0 + \frac{1}{a_1 + \frac{1}{a_2 + \dots}} = a_0 + \frac{1}{a_1 + \dots} \frac{1}{a_k + \frac{1}{x_{k+1}}}$$

and convergents

$$\frac{p_k}{q_k} = a_0 + \frac{1}{a_1 + \dots} \frac{1}{a_k},$$

then

$$x = \frac{p_k x_{k+1} + p_{k-1}}{q_k x_{k+1} + q_{k-1}}$$

and

$$\left| q_k x - p_k \right| = \frac{1}{q_k x_{k+1} + q_{k-1}}.$$

LEMMA 2. *Put*

$$W_n = \text{No} \left\{ k \leq n \mid |q_k x - p_k| < \frac{f(k)}{q_k} \right\}.$$

Then

$$\lim_{n \rightarrow \infty} \mathfrak{M} \left\{ W_n < \frac{1}{\log 2} \sum_{k=1}^n f(k) + \omega \left(\frac{1}{\log 2} \sum_{k=1}^n f(k) \right)^{1/2} \right\} = \phi(\omega).$$

We take x as a random variable uniformly distributed on $[0, 1]$, and use Pr_k , E_k and Var_k to denote conditional probability, expectation and variance when $a_0, \dots, a_k$ are given. We suppose throughout this section that f satisfies conditions (5)–(7), and we put $\alpha_k = f(k)(1 + q_{k-1}/q_k)$ and

$$V_k = \begin{cases} 1 - \alpha_k & \text{if } |q_k x - p_k| < \frac{f(k)}{q_k}, \\ -\alpha_k & \text{otherwise.} \end{cases}$$

Then

$$\begin{aligned} \text{Pr}_k \{ V_k = 1 - \alpha_k \} &= \text{Pr}_k \left\{ \frac{1}{(q_k x_{k+1} + q_{k-1})} < \frac{f(k)}{q_k} \right\} \\ &= \text{Pr}_k \left\{ x_{k+1} > \frac{1}{f(k)} - \frac{q_{k-1}}{q_k} \right\} \\ &= \text{Pr}_k \left\{ x \in \left[\frac{p_k(1/f(k) - q_{k-1}/q_k) + p_{k-1}}{q_k(1/f(k) - q_{k-1}/q_k) + q_{k-1}}, \frac{p_k}{q_k} \right] \right\} \\ &= \frac{\left| \frac{p_k q_k / f(k) \pm 1}{q_k^2 / f(k)} - \frac{p_k}{q_k} \right|}{\left| \frac{p_k + p_{k-1}}{q_k + q_{k-1}} - \frac{p_k}{q_k} \right|} \\ &= f(k) \left(1 + \frac{q_{k-1}}{q_k} \right) = \alpha_k. \end{aligned}$$

Hence

$$\begin{aligned} E_k(V_k) &= (1 - \alpha_k)\alpha_k + (-\alpha_k)(1 - \alpha_k) = 0, \\ (8) \quad \mu_k^2 &= E_k(V_k^2) = f(k) \left(1 + \frac{q_{k-1}}{q_k} \right) + O(f^2(k)). \end{aligned}$$

P. Lévy [9; 10, p. 321] has shown that

$$\Pr \left\{ \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \left(1 + \frac{q_{k-1}}{q_k} \right) \sim \frac{1}{\log 2} \right\} = 1,$$

and it follows from (f) of Lemma 1 that for almost all x ,

$$(9) \quad \sum_{k=1}^n f(k) \left(1 + \frac{q_{k-1}}{q_k} \right) = \frac{1}{\log 2} \sum_{k=1}^n f(k) + O(1).$$

Combining (8) and (9), we see that for almost all x ,

$$(10) \quad \mu_1^2 + \cdots + \mu_n^2 = \frac{1}{\log 2} \sum_{k=1}^n f(k) + O(1).$$

We now use a form of the central limit theorem for dependent variables due to Lévy [10, p. 246] (and later extended by J. L. Doob [2, p. 383] as a theorem on martingales):

LEMMA 3. *Let $Z_1, Z_2, \dots$ be a sequence of bounded random variables, and let E_{n-1} denote conditional expectation for given $Z_1, \dots, Z_{n-1}$. Suppose that $E_{n-1}(Z_n) = 0$ for $n \geq 2$, and put*

$$\mu_n^2 = E_{n-1}(Z_n^2) = \text{Var}_{n-1}(Z_n).$$

For $t > 0$, determine $N = N(t)$ so that

$$\mu_1^2 + \cdots + \mu_N^2 \sim t,$$

and put

$$S(t) = Z_1 + \cdots + Z_N.$$

Then if

$$\Pr \left\{ \sum_{n=1}^{\infty} \mu_n^2 < \infty \right\} = 0,$$

we have

$$\lim_{t \rightarrow \infty} \Pr \left\{ \frac{S(t)}{t^{1/2}} < \omega \right\} = \phi(\omega).$$

If $Z_k = V_k$, it follows from (10) that aside from a set of measure 0, the functions $N(t)$ corresponding to various x 's are asymptotically equal, and that

$$\lim_{n \rightarrow \infty} \Pr \left\{ \frac{V_1 + \cdots + V_n}{\left(\frac{1}{\log 2} \sum_{k=1}^n f(k) \right)^{1/2}} < \omega \right\} = \phi(\omega).$$

But

$$W_n = \sum_{k=1}^n V_k + \sum_{k=1}^n f(k) \left(1 + \frac{q_{k-1}}{q_k} \right),$$

and hence for almost all x ,

$$W_n = \sum_{k=1}^n V_k + \frac{1}{\log 2} \sum_{k=1}^n f(k) + O(1).$$

Thus

$$(11) \quad \lim_{n \rightarrow \infty} \Pr \left\{ W_n < \frac{1}{\log 2} \sum_{k=1}^n f(k) + \omega \left(\frac{1}{\log 2} \sum_{k=1}^n f(k) \right)^{1/2} \right\} = \phi(\omega),$$

which completes the proof of the lemma.

The remainder of the proof of Theorem 1 consists in transforming (11) into a statement not involving continued fractions. For this we need an estimate of q_k .

LEMMA 4. *If $\delta < 1/2$, then for almost every x there is a constant $\kappa = \kappa(x, \delta)$ such that*

$$\left| \log q_k - \frac{\pi^2}{12 \log 2} k \right| < \kappa k^{1-\delta}.$$

This results from an extension of the following theorem of Khinchin [7]:
Let F be a function of k positive integral arguments, such that for $n \geq k$,

$$\int_0^1 F^2(a_n, \dots, a_{n-k+1}) dx < C,$$

where $a_m = a_m(x)$ denotes the m th denominator in the continued fraction expansion of x . Then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=k}^n F(a_i, \dots, a_{i-k+1})$$

exists and is constant almost everywhere.

Examination of the proof shows that the theorem may be modified in two ways. The function F may be replaced by a quantity depending on a slowly increasing number of the a_m ; we write

$$(12) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n F_i(a_i, a_{i-1}, \dots, a_{i-k_i+1}),$$

and require that $i - k_i + 1$ be positive for $i \geq 1$. Secondly, the rapidity of approach of the sum in (12) to its limiting value can be estimated by replacing the ϵ occurring in Khinchin's proof by $n^{-\epsilon}$, where ϵ is now a sufficiently small

positive constant. In this way the following theorem can be proved:

Let $\{F_i(r_1, \dots, r_{k_i})\}$ be non-negative functions of the positive integral arguments $r_1, r_2, \dots$, and suppose that the integrals

$$\int_0^1 F_i^2(a_i, a_{i-1}, \dots, a_{i-k_i+1}) dx$$

are uniformly bounded. Suppose further that $\delta < 1/2$ and that

$$k_i = O(\log^\sigma i)$$

for some constant $\sigma > 0$. Then there is a constant B such that

$$\frac{1}{n} \sum_{i=1}^n F_i(a_i, \dots, a_{i-k_i+1}) = B + O(n^{-\delta})$$

for almost all x .

We put

$$\phi_i(x) = a_i + \frac{1}{a_{i-1} +} \dots \frac{1}{a_{i-k_i+1}}$$

and

$$F_i(a_i, \dots, a_{i-k_i+1}) = \log \phi_i(x).$$

Since $\phi_i(x) \leq a_i + 1$ and $\mathfrak{M}\{a_i = r\} = \mathfrak{M}\{r \leq x_i < r+1\} < 1/r^2$, we have

$$\int_0^1 F_i^2 dx \leq \int_0^1 \log^2(a_i + 1) dx \leq \sum_{r=1}^{\infty} \frac{\log^2(r+1)}{r^2}.$$

Thus for

$$k_i = 1 + [2 \log i]$$

there is a B_0 such that for almost all x ,

$$\sum_{i=1}^n \log \phi_i(x) = B_0 n + O(n^{1-\delta}).$$

On the other hand, if $\bar{\phi}_i(x) = q_i/q_{i-1}$, then by the law of the mean,

$$|\log \phi_i(x) - \log \bar{\phi}_i(x)| = \xi |\phi_i(x) - \bar{\phi}_i(x)|$$

where $\xi < 1$. Since

$$(13) \quad \bar{\phi}_i(x) = a_i + \frac{1}{a_{i-1} +} \dots \frac{1}{a_1},$$

this implies that

$$\begin{aligned} & \left| \log \phi_i(x) - \log \bar{\phi}_i(x) \right| \\ & < \left| \left(a_i + \frac{1}{a_{i-1}+} \cdots \frac{1}{a_{i-k_i+1}} \right) - \left(a_i + \frac{1}{a_{i-1}+} \cdots \frac{1}{a_{i-k_i+1}+1} \right) \right| < 1/Q_{k_i}^2, \end{aligned}$$

where P_l/Q_l is the l th convergent in the expansion (13). Since

$$Q_l \geq Q_{l-1} + Q_{l-2} > 2Q_{l-2} > \cdots > 2^{l/2},$$

we see that

$$\left| \log \phi_i(x) - \log \bar{\phi}_i(x) \right| < 2^{1-k_i} < i^{-2 \log 2}.$$

Thus for almost all x ,

$$\sum_{i=1}^n \log \bar{\phi}_i(x) = \log q_n = B_0 n + O(n^{1-\delta}).$$

Lévy [10, p. 320] showed that $B_0 = \pi^2/12 \log 2$. The proof of Lemma 4 is complete.

Now let

$$\begin{aligned} s_n &= \text{No} \left\{ k \leq n \mid \left| q_k x - p_k \right| < \frac{f(B_0^{-1} \log q_k)}{q_k} \right\}, \\ t_n(\kappa) &= \text{No} \left\{ k \leq n \mid \left| q_k x - p_k \right| < \frac{f(k - \kappa k^{1-\delta})}{q_k} \right\}. \end{aligned}$$

By (11),

$$\lim_{n \rightarrow \infty} \mathfrak{M} \left\{ t_n(\kappa) < \frac{1}{\log 2} \sum_{k=1}^n f(k - \kappa k^{1-\delta}) + \omega \left(\frac{1}{\log 2} \sum_{k=1}^n f(k - \kappa k^{1-\delta}) \right)^{1/2} \right\} = \phi(\omega).$$

Putting

$$A_n = \frac{1}{\log 2} \sum_{k=1}^n f(k),$$

it follows from (b) of Lemma 1 that for each κ ,

$$(14) \quad \lim_{n \rightarrow \infty} \mathfrak{M} \{ t_n(\kappa) < A_n + \omega A_n^{1/2} \} = \phi(\omega).$$

Let

$$\begin{aligned} F_n &= \{ x \mid s_n < A_n + \omega A_n^{1/2} \}, \\ G(\kappa) &= \{ x \mid \left| \log q_k - B_0 k \right| < \kappa k^{1-\delta} \text{ for every } k \geq 1 \}, \\ H_n(\kappa) &= \{ x \mid t_n(\kappa) < A_n + \omega A_n^{1/2} \}. \end{aligned}$$

Then by Lemma 2 and Equation (14), to each $\epsilon > 0$ there corresponds a $\kappa_0 = \kappa_0(\epsilon)$ and an $n_0 = n_0(\kappa_0, \epsilon) = n_0(\epsilon)$ such that

$$\mathfrak{N}\{G(\kappa)\} > 1 - \epsilon \quad \text{for } \kappa \geq \kappa_0$$

and

$$|\mathfrak{N}\{H_n(\pm \kappa_0)\} - \phi(\omega)| < \epsilon \quad \text{for } n \geq n_0.$$

Clearly

$$G(\kappa_0)H_n(\kappa_0) \subset F_n,$$

and since $\mathfrak{N}(AB) \geq \mathfrak{N}(A) + \mathfrak{N}(B) - 1$ if A and B are subsets of $[0, 1]$, we have that for $n \geq n_0$,

$$\mathfrak{N}\{F_n\} \geq 1 - \epsilon + \phi(\omega) - \epsilon - 1 = \phi(\omega) - 2\epsilon.$$

Similarly, since $G(\kappa_0)F_n \subset H_n(-\kappa_0)$,

$$\mathfrak{N}\{F_n\} \leq \phi(\omega) + 2\epsilon.$$

Hence

$$\lim_{n \rightarrow \infty} \mathfrak{N}\{F_n\} = \lim_{n \rightarrow \infty} \Pr \{s_n < A_n + \omega A_n^{1/2}\} = \phi(\omega).$$

By the same reasoning we can use (d) of Lemma 1 to show that if

$$r_n = \text{No} \left\{ k \leq n \mid |q_k x - p_k| < \frac{B_0 f(\log q_k)}{q_k} \right\},$$

then

$$\lim_{n \rightarrow \infty} \Pr \{r_n < A_n + \omega A_n^{1/2}\} = \phi(\omega).$$

Replacing f by f/B_0 , it follows immediately that

$$\begin{aligned} \lim_{n \rightarrow \infty} \Pr \left\{ \text{No} \left\{ k \leq n \mid |q_k x - p_k| < \frac{f(\log q_k)}{q_k} \right\} < \frac{12}{\pi^2} \sum_{k=1}^n f(k) \right. \\ \left. + \omega \left(\frac{12}{\pi^2} \sum_{k=1}^n f(k) \right)^{1/2} \right\} = \phi(\omega). \end{aligned}$$

If $|mx - l| < 1/2m$, then l/m is a convergent to x . Since $f(x) = o(1)$,

$$\begin{aligned} \text{No} \left\{ k \leq n \mid |q_k x - p_k| < \frac{f(\log q_k)}{q_k} \right\} \\ = \text{No} \left\{ m \leq q_n \mid \langle mx \rangle < \frac{f(\log m)}{m} \right\} + O(1), \end{aligned}$$

the error term being uniformly bounded for all x . Putting

$$A(n) = \frac{12}{\pi^2} \sum_{k=1}^n \frac{f(\log k)}{k}$$

and using (e) of Lemma 1 with $c=1$, it follows that

$$(15) \quad \lim_{n \rightarrow \infty} \Pr \left\{ \text{No} \left\{ m \leq q_n \mid \langle mx \rangle < \frac{f(\log m)}{m} \right\} < A(n) + \omega A(n)^{1/2} \right\} = \phi(\omega).$$

There is now a final set-theoretic argument required to eliminate q_n entirely. Put

$$\begin{aligned} F(n, \omega) &= E_x \left\{ \text{No} \left\{ m \leq q_n \mid \langle mx \rangle < \frac{f(\log m)}{m} \right\} < A(n) + \omega A(n)^{1/2} \right\}, \\ G(n, \beta, \omega) &= E_x \left\{ \text{No} \left\{ m \leq e^{\beta n} \mid \langle mx \rangle < \frac{f(\log m)}{m} \right\} < A(n) + \omega A(n)^{1/2} \right\}, \\ H_N(\epsilon) &= E_x \{ e^{B_0(1-\epsilon)\nu} < q_\nu < e^{B_0(1+\epsilon)\nu} \text{ for all } \nu \geq N \}. \end{aligned}$$

It is easily seen that

$$(16) \quad H_N(\epsilon)G(n, B_0(1+\epsilon), \omega) \subset F(n, \omega), \quad H_N(\epsilon)F(n, \omega) \subset H_N(\epsilon)G(n, B_0(1-\epsilon), \omega)$$

for $0 < \epsilon < 1$, $n \geq N$. On the other hand, we have

$$\begin{aligned} G\left(\frac{1-\epsilon}{1+\epsilon}n, B_0(1+\epsilon), \eta\right) &= E_x \left\{ \text{No} \left\{ m \leq e^{B_0(1-\epsilon)n} \mid \langle mx \rangle < \frac{f(\log m)}{m} \right\} \right. \\ &\quad \left. < A\left(\frac{1-\epsilon}{1+\epsilon}n\right) + \eta A^{1/2}\left(\frac{1-\epsilon}{1+\epsilon}n\right) \right\}, \end{aligned}$$

and hence if η is chosen so that

$$(17) \quad A\left(\frac{1-\epsilon}{1+\epsilon}n\right) + \eta A^{1/2}\left(\frac{1-\epsilon}{1+\epsilon}n\right) > A(n) + \omega A^{1/2}(n),$$

then

$$(18) \quad G\left(\frac{1-\epsilon}{1+\epsilon}n, B_0(1+\epsilon), \eta\right) \supset G(n, B_0(1-\epsilon), \omega).$$

By (c) of Lemma 1, $A(cn) = A(n) + O(1)$, so

$$A\left(\frac{1-\epsilon}{1+\epsilon}n\right) + \eta A^{1/2}\left(\frac{1-\epsilon}{1+\epsilon}n\right) = A(n) + (\eta + O(A^{-1/2}(n)))A^{1/2}(n).$$

Since $A(n) \rightarrow \infty$ as $n \rightarrow \infty$, it follows that if $\delta > 0$ is arbitrary, (17) holds with $\eta = \omega + \delta$, if $n > n_0(\epsilon, \delta)$. But then by (16) and (18),

$$\begin{array}{ll} \mathbf{r}_1 \mathbf{x} = \mathbf{a}_1 + \mathbf{x}_1, & -1/2 \leq x_1 < 1/2, \\ \mathbf{r}_2 \mathbf{x}_1 = \mathbf{a}_2 + \mathbf{x}_2, & -1/2 \leq x_2 < 1/2, \\ \cdot & \cdot \\ \mathbf{r}_n \mathbf{x}_{n-1} = \mathbf{a}_n + \mathbf{x}_n, & -1/2 \leq x_n < 1/2, \\ \cdot & \cdot \end{array}$$

Then

$$(19) \quad a_n = [r_n x_{n-1} + 1/2], \quad |x_n| = \langle r_n x_{n-1} \rangle,$$

$$(20) \quad -\left[\frac{r_n}{2}\right] \leq a_n < \left[\frac{r_n}{2}\right],$$

for $n=1, 2, \dots$, and

$$(21) \quad x = \sum_{n=1}^{\infty} \frac{a_n}{r_1 \cdots r_n}.$$

The series (21) bears an obvious relation to the expansion of x to the base r if, contrary to assumption, we take all $r_n=r$, and to the Cantor factorial expansion if $r_n=n$ for all n . In any case, the expansion is unique except for a set of measure zero.

Since x is a random variable, so is every element of $\{x_n\}$ and $\{a_n\}$, and it is easily seen that each x_n is uniformly distributed on $[-1/2, 1/2]$, and that each a_n is discretely uniformly distributed, in the sense that

$$(22) \quad \Pr \{a_n = j\} = \frac{1}{r_n} \quad \text{for} \quad -\left[\frac{r_n}{2}\right] \leq j < \left[\frac{r_n}{2}\right].$$

There is a significant difference between the two sets of variables, however, in that the a_n are statistically independent, while the x_n are not, as the Equations (19) show. Dependence makes the sequence $\{x_n\}$ difficult to analyze probabilistically, but a considerable amount of information can be gained indirectly by transferring results about $\{a_n\}$ via the relation

$$x_{n-1} = \frac{a_n}{r_n} + O\left(\frac{1}{r_n}\right).$$

THEOREM 2. *Suppose that $r_1, r_2, \dots$ is a nondecreasing sequence of positive integers such that $r_n^m > n$ for some fixed integer m . Let $R_n = r_1 r_2 \cdots r_n$, and let f be a positive function. Let S be an increasing sequence of positive integers. Then the inequality*

$$(23) \quad \langle R_n x \rangle < f(n)$$

has infinitely many solutions $n \in S$ for almost all x or almost no x , according as the series

$$(24) \quad \sum_{n \in S} f(n)$$

diverges or converges.

We note first that it suffices to consider functions f such that $f(n) \geq n^{-2}$ for all $n \in S$. For if (24) converges, then so does the series

$$\sum_{n \in S} f^*(n),$$

where

$$f^*(n) = \begin{cases} f(n) & \text{if } f(n) \geq n^{-2}, \\ n^{-2} & \text{otherwise,} \end{cases}$$

and if the inequality $\langle R_n x \rangle < f^*(n)$ has only finitely many solutions in S , the same is surely true of (23). Suppose on the other hand that (24) diverges. Then so also does

$$\sum f(n_j),$$

the summation being extended over the integers $n_j \in S$ such that $f(n_j) \geq n_j^{-2}$. These integers constitute a subsequence S' of S , and the truth of the theorem for S' implies its truth for S .

We suppose throughout the proof that $n \in S$. If we put

$$P_n = R_n \sum_{j=1}^n \frac{a_j}{R_j},$$

then

$$|R_n x - P_n| = |x_n| \leq 1/2,$$

so

$$|R_n x - P_n| = \langle R_n x \rangle.$$

For each n let k_n be the unique positive integer such that

$$(25) \quad [r_{n+1} \cdots r_{n+k_n-1} f(n) + 1/2] = 0, \quad [r_{n+1} \cdots r_{n+k_n} f(n) + 1/2] \neq 0;$$

in particular, if $[r_{n+1} f(n) + 1/2] \neq 0$ then $k_n = 1$. Then

$$(26) \quad \frac{1}{r_{n+1} \cdots r_{n+k_n}} \leq 2f(n).$$

Let ε_n be the event that (i.e., the set of $x \in [0, 1]$ such that)

$$a_{n+1} = \cdots = a_{n+k_n-1} = 0, \quad |a_{n+k_n}| < r_{n+1} \cdots r_{n+k_n} f(n) + 1 = \frac{R_{n+k_n}}{R_n} f(n) + 1,$$

and for $c > 0$ let $\mathfrak{F}_n(c)$ be the event that $\langle R_n x \rangle < c f(n)$.

Suppose that $x \in \mathfrak{F}_n(1)$. If $k_n = 1$, then we have

$$|x_n| < f(n),$$

$$|a_{n+1}| = |a_{n+k_n}| = |[r_{n+1} x_n + 1/2]| \leq r_{n+1} |x_n| + 1/2 < r_{n+k_n} f(n) + 1,$$

so $x \in \varepsilon_n$. If $k_n > 1$, then

$$\begin{aligned}
|a_{n+1}| &\leq [r_{n+1}f(n) + 1/2] = 0, & x_{n+1} &= r_{n+1}x, \\
|a_{n+2}| &\leq [r_{n+1}r_{n+2}f(n) + 1/2] = 0, & x_{n+2} &= r_{n+1}r_{n+2}x_n, \\
&\dots\dots\dots \\
|a_{n+k_n-1}| &\leq [r_{n+1} \cdots r_{n+k_n-1}f(n) + 1/2] = 0, & x_{n+k_n-1} &= r_{n+1} \cdots r_{n+k_n-1}x_n, \\
|a_{n+k_n}| &\leq [r_{n+1} \cdots r_{n+k_n}f(n) + 1/2] < r_{n+1} \cdots r_{n+k_n}f(n) + 1,
\end{aligned}$$

and again $x \in \mathcal{E}_n$. Hence $\mathcal{F}_n(1) \subset \mathcal{E}_n$.

On the other hand, if $x \in \mathcal{E}_n$ then

$$x = \sum_{j=1}^n \frac{a_j}{R_j} + \sum_{j=n+k_n}^{\infty} \frac{a_j}{R_j},$$

so

$$\begin{aligned}
(27) \quad \langle R_n x \rangle &= |R_n x - P_n| < \frac{(|a_{n+k_n}| + 1)R_n}{R_{n+k_n}} < \frac{\left(\frac{R_{n+k_n}}{R_n}f(n) + 2\right)R_n}{R_{n+k_n}} \\
&= f(n) + \frac{2R_n}{R_{n+k_n}}
\end{aligned}$$

and it follows from (26) that $\mathcal{E}_n \subset \mathcal{F}_n(3)$.

Thus if $\mathcal{E}_n$ occurs for only finitely many $n \in S$, the same is true of $\mathcal{F}_n(1)$; while if $\mathcal{E}_n$ occurs for infinitely many $n \in S$, the same is true of $\mathcal{F}_n(3)$. Since the convergence of (24) is unaffected by replacing $f(n)$ by $3f(n)$, there remains only the task of showing that $\mathcal{E}_n$ occurs for infinitely many $n \in S$, or only finitely many $n \in S$, for almost all x , according as (24) diverges or converges.

Since $r_n^m > n$ and $f(n) > n^{-2}$, we have $r_{n+1} \cdots r_{n+2m}f(n) > 1$. Hence $k_n \leq 2m$, and the event $\mathcal{E}_n$ depends on at most the $2m$ random variables $a_{n+1}, \dots, a_{n+2m}$. Hence for fixed l ($0 \leq l < 2m$), the events $\mathcal{E}_{2\nu m+l}$ ($\nu = 0, 1, \dots$) are independent. By (22),

$$\Pr \{ |a_n| = j \} = \begin{cases} \frac{1}{r_n} & \text{if } j = 0, \\ \frac{2}{r_n} & \text{if } 0 < j < \frac{r_n}{2}, \\ \frac{1}{r_n} & \text{if } j = \frac{r_n}{2}, \quad r_n \text{ even.} \end{cases}$$

Hence for arbitrary real $u \in [0, r_n/2)$,

$$\Pr \{ |a_n| \leq u \} = \frac{2[u] + 1}{r_n} \begin{cases} \leq (2u + 1)/r_n, \\ \geq (2u - 1)/r_n. \end{cases}$$

Thus, because of the independence of the a_n , we have

$$(28) \quad \Pr \{ \varepsilon_n \} \leq \frac{1}{r_{n+1}} \cdots \frac{1}{r_{n+k_n-1}} \cdot \frac{2 \frac{R_{n+k_n}}{R_n} f(n) + 3}{r_{n+k_n}} = 2f(n) + \frac{3R_n}{R_{n+k_n}},$$

and by (26),

$$\Pr \{ \varepsilon_n \} < 8f(n).$$

Also

$$(29) \quad \Pr \{ \varepsilon_n \} \geq \frac{2 \frac{R_{n+k_n}}{R_n} f(n) + 1}{r_{n+1} \cdots r_{n+k_n}} > 2f(n).$$

Hence for each l the series⁽¹⁾

$$\sum_{\nu: 2\nu m + l \in S} \Pr \{ \varepsilon_{2\nu m + l} \}$$

converges or diverges with the series

$$(30) \quad \sum_{\nu: 2\nu m + l \in S} f(2\nu m + l).$$

But if the series

$$(31) \quad \sum_{n \in S} f(n)$$

diverges, at least one of the series (30), for $0 \leq l < 2m$, must diverge, while if (31) converges, all the series (30) converge. The theorem therefore follows from the Borel-Cantelli lemmas.

5. We now consider the case in which (23) has infinitely many solutions for almost all x , and investigate the number of such solutions with $n \leq N$. For simplicity we suppose that S is the full set of positive integers.

THEOREM 3. *Let $\{r_n\}$ and $\{R_n\}$ be as described in Theorem 2. Let f be a positive function such that*

$$\sum_{n=1}^{\infty} f(n) = \infty, \quad f(n) = O(n^{-1/2-\epsilon}).$$

Let k_n be the positive integer defined in (25), and suppose that

$$(32) \quad \sum_{n=1}^{\infty} (r_{n+1} \cdots r_{n+k_n})^{-1} < \infty.$$

(1) The symbol $\sum_{\nu: \dots}$ means summation over those ν such that $\dots$.

Then

$$(33) \quad \lim_{N \rightarrow \infty} \Pr \left\{ \text{No } \left\{ n \leq N \mid \langle R_n x \rangle < f(n) \right\} \right. \\ \left. < 2 \sum_{n=1}^N f(n) + \omega \left(2 \sum_{n=1}^N f(n) \right)^{1/2} \right\} = \phi(\omega).$$

According to Theorem 2, the n for which $f(n) < n^{-2}$ contribute only a bounded number of solutions of the inequality (23), so we may suppose that $f(n) \geq n^{-2}$. Put

$$X_n = \begin{cases} 1 & \text{if } \langle R_n x \rangle < f(n), \\ 0 & \text{otherwise,} \end{cases}$$

and

$$S_N = \sum_{n=1}^N X_n.$$

Similarly, put

$$Y_n = \begin{cases} 1 & \text{if } \varepsilon_n \text{ occurs,} \\ 0 & \text{otherwise} \end{cases}$$

and

$$T_N = \sum_{n=1}^N Y_n,$$

where ε_n has the same meaning as before. Since $\mathfrak{F}_n(1) \subset \varepsilon_n$, we have

$$(34) \quad S_N < T_N.$$

On the other hand, if $Y_n = 1$ then either $X_n = 1$ or

$$(35) \quad \langle R_n x \rangle \in \left[f(n), f(n) + \frac{2R_n}{R_{n+k_n}} \right],$$

by (27). Because of the uniform distribution of the x_n , the probability of the event (35) is $2R_n/R_{n+k_n}$, and by (32) and the first Borel-Cantelli lemma, the event (35) occurs only finitely many times, for almost all x . Thus given $\epsilon > 0$, there is a constant M so large that

$$(36) \quad T_N < S_N + M$$

for all N and all x not in a set of measure at most ϵ . Combining (34) and (36), we see that (33) will follow if it can be shown that

$$(37) \quad \lim_{N \rightarrow \infty} \Pr \left\{ T_N < 2 \sum_{n=1}^N f(n) + \omega \left(2 \sum_{n=1}^N f(n) \right)^{1/2} \right\} = \phi(\omega).$$

To this end we first prove a general lemma, suggested by work of Hoeffding and Robbins [5]. A set of random variables $Z_1, Z_2, \dots$ is said to be m -dependent if for every r, s and n for which $n > s > r + m$, the sets $Z_1, \dots, Z_r$ and $Z_s, \dots, Z_n$ are independent. (The variables Y_n above are $2m$ -dependent.)

THEOREM 4. *Let $Z_1, Z_2, \dots$ be a sequence of m -dependent random variables such that*

$$Z_n = \begin{cases} 1 & \text{with probability } p_n, \\ 0 & \text{with probability } 1 - p_n. \end{cases}$$

Suppose that

$$(38) \quad \sum_{n=1}^{\infty} p_n = \infty,$$

$$(39) \quad p_n = O(n^{-1/2-\epsilon}), \quad \epsilon > 0,$$

$$(40) \quad \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} |\text{Cov}(Z_i, Z_{i+j})| < \infty.$$

Then

$$\lim_{n \rightarrow \infty} \Pr \left\{ Z_1 + \dots + Z_n < \sum_{k=1}^n p_k + \omega \left(\sum_{k=1}^n p_k \right)^{1/2} \right\} = \phi(\omega).$$

We decompose the finite sequence $1, 2, \dots, n$ into blocks, in the following way. Choose η smaller than ϵ , and find an integer l_0 such that

$$(41) \quad (l_0 + 1)^{2+\eta} - l_0^{2+\eta} > 2m.$$

For $q \geq 1$ put

$$l_q = [(l_0 + q)^{2+\eta}],$$

and define $\kappa = \kappa(n)$ by the inequality

$$l_\kappa \leq n < l_{\kappa+1}.$$

For $1 \leq q < \kappa - 1$, let I_{q+1} be the set of integers j such that $l_q < j \leq l_{q+1} - m$, and let J_{q+1} be the set of integers j such that $l_{q+1} - m < j \leq l_{q+1}$. Finally, put

$$U_q = \sum_{\nu \in I_q} Z_\nu = \sum_{I_q} Z_\nu,$$

$$V_q = \sum_{J_q} Z_\nu,$$

for $q=2, \dots, \kappa$, so that

$$Q_n = \sum_{\nu=1}^n Z_\nu = \sum_{\nu=1}^{l_1} Z_\nu + \sum_{q=2}^{\kappa} U_q + \sum_{q=2}^{\kappa} V_q + \sum_{\nu=l_{\kappa}+1}^n Z_\nu.$$

By the definitions of l_0 and m -dependence, the variables $U_2, \dots, U_{\kappa}$ are independent, as are $V_2, \dots, V_{\kappa}$. We shall show that the limiting behavior of Q_n is determined by that of $\sum U_q$, and then apply a standard version of the central limit theorem.

Since l_1 is fixed and the Z 's are bounded, the sum

$$\sum_{\nu=1}^{l_1} Z_\nu$$

is clearly negligible in the limit, if $\text{Var}(S_n) \rightarrow \infty$. By (40), (39), and (38),

$$\begin{aligned} \text{Var} \left(\sum_{q=2}^{\kappa} V_q \right) &= \sum_{q=2}^{\kappa} \sum_{J_q} \text{Var}(Z_\nu) + 2 \sum_{q=2}^{\kappa} \sum_{J_q} \text{Cov}(Z_\mu, Z_\nu) \\ &= \sum_{q=2}^{\kappa} \sum_{J_q} (p_\nu - p_\nu^2) + O(1) \\ &= \sum_{q=2}^{\kappa} \sum_{J_q} p_\nu + O(1) \\ &= \sum_{q=2}^{\kappa} \sum_{\nu=1}^m O(l_q^{-1/2-\epsilon}) + O(1) \\ &= O \left(\sum_{q=2}^{\kappa} q^{-1-2\epsilon-\eta/2-\epsilon\eta} \right) + O(1), \end{aligned}$$

so that

$$(42) \quad \text{Var} \left(\sum_{q=2}^{\kappa} V_q \right) = O(1).$$

Turning to U_q , we see that

$$(43) \quad E(U_q) = \sum_{I_q} p_\nu = e_q,$$

and

$$\text{Var}(U_q) = \sum_I \text{Var}(Z_\nu) + 2 \sum_{\mu, \nu \in I_q; \mu < \nu < \mu+m} \text{Cov}(Z_\mu, Z_\nu),$$

so that

$$(44) \quad \sigma_{\kappa}^2 = \text{Var}(U_2 + \dots + U_{\kappa}) = \sum_{q=2}^{\kappa} e_q + O(1).$$

Now

$$\begin{aligned} e_q &< c \sum_{l_q} \frac{1}{p^{1/2+\epsilon}} < c(l_q^{1/2-\epsilon} - l_{q-1}^{1/2-\epsilon}) \\ &< cl_{q-1}^{1/2-\epsilon} \left\{ \left(1 + \frac{1}{q} \right)^{(1/2-\epsilon)(2+\eta)} - 1 \right\} \\ &= O\left(q^{(2+\eta)(1/2-\epsilon)} \cdot \frac{1}{q} \right) \end{aligned}$$

and so

$$(45) \quad e_q = O(1).$$

This implies in particular that

$$(46) \quad \text{Var} \left(\sum_{\nu=l_k+1}^n Z_\nu \right) = O(1),$$

and hence, since

$$\sum_{q=2}^{\infty} \sum_{J_q} p_\nu < \infty,$$

that

$$(47) \quad \sigma_\kappa^2 = \sum_{\nu=1}^n p_\nu + O(1), \quad E(U_2 + \cdots + U_\kappa) = \sum_{\nu=1}^n p_\nu + O(1).$$

If we put

$$\pi_n = \sum_{\nu=1}^n p_\nu,$$

then (42) shows that

$$\text{Var} \left(\pi_n^{-1/2} \sum_{q=2}^{\kappa} V_q \right) = O(1),$$

and it follows from Chebyshev's inequality that the random variable $\pi_n^{-1/2} \sum_2^\kappa V_q$ approaches zero in probability. By the same reasoning this is true also of $\pi_n^{-1/2} \sum_1^{l_1} Z_\nu$. Combining these facts with (46), we see [1, p. 254] that the limiting distribution of $(Q_n - \pi_n)/\pi_n^{1/2}$ is identical with that of

$$(48) \quad (U_2 + \cdots + U_\kappa - \pi_n)/\pi_n^{1/2}.$$

We now wish to apply Lyapunov's criterion [1, p. 213], according to which the normalized sum (48) is asymptotically normal, with mean zero and variance 1, if

$$(49) \quad \left(\sum_{q=2}^{\kappa} \rho_q^3 \right)^{1/3} = O(\sigma_{\kappa}),$$

where

$$\rho_q^3 = E(|U_q - E(U_q)|^3).$$

This will complete the proof of Theorem 4. We have

$$\begin{aligned} \rho_q^3 &\leq E \left\{ \left(\sum_{I_q} |Z_v - p_v| \right)^3 \right\} \\ &< 6E \left\{ \sum_{v \in I_q} |Z_v - p_v|^3 + \sum_{\mu, v \in I_q} |Z_{\mu} - p_{\mu}| \cdot |Z_v - p_v|^2 \right. \\ &\quad \left. + \sum_{\mu, v, \lambda \in I_q} |Z_{\mu} - p_{\mu}| \cdot |Z_v - p_v| \cdot |Z_{\lambda} - p_{\lambda}| \right\}. \end{aligned}$$

Now

$$\begin{aligned} \sum_{I_q} E(|Z_v - p_v|^3) &= \sum_{I_q} (1 - p_v)^3 p_v + \sum_{I_q} p_v^3 (1 - p_v) \\ &= e_q + O \left(\sum_{I_q} p_v^2 \right). \end{aligned}$$

Since $|Z_v - p_v| < 1$, we have, by the generalized Hölder inequality [4, p. 140],

$$\begin{aligned} \sum_{\mu, v \in I_q} E(|Z_{\mu} - p_{\mu}| \cdot |Z_v - p_v|^2) &\leq \sum_{\mu, v \in I_q} E(|Z_{\mu} - p_{\mu}| \cdot |Z_v - p_v|) \\ &\leq \left(\sum_{\mu, v \in I_q} \text{Var}(Z_{\mu}) \text{Var}(Z_v) \right)^{1/2} \leq \sum_{\mu \in I_q} \text{Var}(Z_{\mu}) \\ &= \sum_{\mu \in I_q} (p_{\mu} - p_{\mu}^2) = e_q + O \left(\sum_{I_q} p_{\mu}^2 \right). \end{aligned}$$

Similarly,

$$\begin{aligned} \sum_{\mu, v, \lambda \in I_q} E(|Z_{\mu} - p_{\mu}| \cdot |Z_v - p_v| \cdot |Z_{\lambda} - p_{\lambda}|) \\ \leq \left\{ \sum_{I_q} E(|Z_{\mu} - p_{\mu}|^3) E(|Z_v - p_v|^3) E(|Z_{\lambda} - p_{\lambda}|^3) \right\}^{1/3} \\ \leq \sum_{I_q} E(|Z_{\mu} - p_{\mu}|^3) = e_q + O \left(\sum_{I_q} p_{\mu}^2 \right). \end{aligned}$$

Thus (49) reduces to the triviality

$$\sum_{q=2}^{\kappa} e_q + O(1) = o \left\{ \left(\sum_{q=2}^{\kappa} e_q \right)^{3/2} \right\}.$$

To complete the proof of Theorem 3, we must show that the hypotheses of Theorem 4 are satisfied when $Z_n = Y_n$, $p_n = \Pr \{ \varepsilon_n \}$. We know that

$$2f(n) \leq p_n \leq 8f(n),$$

and hence, from the hypotheses of Theorem 3, we obtain (38) and (39). Since the Y_n are $2m$ -dependent, we can rewrite (40) in the form

$$(50) \quad \sum_{i=1}^{\infty} \sum_{j=1}^{2m} | \text{Cov} (Y_i, Y_{i+j}) | < \infty.$$

Now if $j > k_n$, then Y_i and Y_{i+j} are independent, and their covariance is 0. If $i \leq j \leq k_n$, then

$$\begin{aligned} | \text{Cov} (Y_i, Y_{i+j}) | &= | E(Y_i Y_{i+j}) - E(Y_i)E(Y_{i+j}) | \\ &= | \Pr \{ Y_i = Y_{i+j} = 1 \} - \Pr \{ Y_i = 1 \} \cdot \Pr \{ Y_{i+j} = 1 \} | \\ &\leq (r_{n+1} \cdots r_{n+k_n})^{-1} + 8f(i)f(i+j), \end{aligned}$$

and the convergence of (50) follows from (32).

6. A strong theorem.

THEOREM 5. *Let $\{R_n\}$ and $f(n)$ satisfy the hypotheses of Theorem 3. Then for almost all x , the number of integers $m \leq n$, for which $\langle R_m x \rangle < f(m)$, is asymptotic to*

$$2 \sum_{k=1}^n f(k).$$

As in the proof of Theorem 3, it suffices to prove the theorem with S_n replaced by $T_n = \sum_{k=1}^n Y_k$, and to suppose that $f(n) > n^{-2}$, so that the Y_k are $2m$ -dependent. We write

$$\begin{aligned} T_n &= \sum^* Y_{2m\nu+1} + \sum^* Y_{2m\nu+2} + \cdots + \sum^* Y_{2m\nu+2m} \\ &= T_n^{(1)} + T_n^{(2)} + \cdots + T_n^{(2m)}, \end{aligned}$$

where each summation extends over those ν for which the subscripts are not larger than n . The terms in $T_n^{(j)}$ are independent and uniformly bounded, and

$$E(T_n^{(j)}) = 2 \sum^* f(2m\nu + j), \quad \text{Var} (T_n^{(j)}) = 2 \sum^* f(2m\nu + j) + O(1).$$

Hence Kolmogorov's version of the law of the iterated logarithm [8] implies that for $1 \leq j \leq 2m$,

$$\Pr \left\{ \limsup_{n \rightarrow \infty} \frac{| T_n^{(j)} - 2 \sum^* f(2m\nu + j) |}{2(\sum^* f(2m\nu + j) \cdot \log \log \sum^* f(2m\nu + j))^{1/2}} = 1 \right\} = 1;$$

and it follows from these equations that

$$\Pr \left\{ \left| T_n - 2 \sum_{k=1}^n f(k) \right| = O \left(\sum_{j=1}^{2m} \left(\sum^* f(2m\nu + j) \cdot \log_2 \sum^* f(2m\nu + j) \right)^{1/2} \right) \right\} = 1,$$

and the theorem is a weak consequence of this result.

Note added in proof.

I. There is a strong version of Theorem 1:

Under the hypotheses of Theorem 1, the number of solutions $m \leq n$ of the inequality $\langle mx \rangle < g(m)$ is asymptotic to

$$\frac{12}{\pi^2} \sum_{k=1}^n g(k),$$

for almost all x .

The proof depends on a strong law of large numbers for dependent variables, due to Lévy [10, p. 253]: *Under the hypotheses of Lemma 3,*

$$\Pr \left\{ \lim_{t \rightarrow \infty} \frac{S(t)}{t^{1/2+\epsilon}} = 0 \right\} = 1$$

for every positive constant ϵ . Using this in place of Lemma 3, we obtain a strong analogue of Lemma 2, to the effect that for $\epsilon > 0$,

$$\Pr \left\{ W_n - (\log 2)^{-1} \sum_1^n f(k) = o \left(\left(\sum_1^n f(k) \right)^{1/2+\epsilon} \right) \right\} = 1,$$

and thereafter the proof parallels that of Theorem 1.

II. It has been pointed out to me that Lemma 3 is not immediately applicable in the proof of Lemma 2, since $E_k(V_k)$, in the equation preceding (8), means $E(V_k, \text{ given } a_0, \dots, a_k)$ and not $E(V_k, \text{ given } V_0, \dots, V_{k-1})$, and it is possible that V_{k-1} , for example, is not uniquely determined by $a_0, \dots, a_k$. But in order for this to be the case it is necessary, since $|q_{k-1}x - p_{k-1}| = (q_{k-1}x + q_{k-2})^{-1}$ and $a_k = [x_k]$, that

$$\frac{1}{q_{k-1}(a_k + 1) + q_{k-2}} < \frac{f(k-1)}{q_{k-1}} < \frac{1}{q_{k-1}a_k + q_{k-2}}.$$

This happens only if

$$a_k = \left[\frac{1}{f(k-1)} - \frac{q_{k-2}}{q_{k-1}} \right].$$

The difficulty vanishes, therefore, if we prove the following theorem, and exclude from the beginning the exceptional set mentioned in it (taking $b=1$ and $h(k)=1/f(k-1)$):

Let h be a real-valued function on the positive integers, with $h(k) > ck$ for some positive constant c . Then for every positive constant b , the set of x , for which the inequality $|a_k - h(k)| < b$ has infinitely many solutions, has measure zero.

Put $F_k(t) = \Pr \{x_k < t\}$; then Lévy's form of the Gauss-Kuzmin theorem [10, pp. 298–306] asserts that for some g with $0 < g < 1$,

$$\left| F_k(t) - \frac{1}{\log 2} \log \frac{2t}{t+1} \right| < g^{k-1}$$

for all $t > 1$ and all positive integers k . Now the inequality $|a_k - h(k)| < b$ is equivalent to

$$h(k) - b < x_k < h(k) + b + 1,$$

and we have

$$\begin{aligned} \lim_{k \rightarrow \infty} \Pr \{h(k) - b < x_k < h(k) + b + 1\} \\ &< \frac{1}{\log 2} \log \left(\frac{2(h(k) + b + 1)}{h(k) + b + 2} \cdot \frac{h(k) - b + 1}{2(h(k) - b)} \right) + 2g^{k-1} \\ &= \frac{1}{\log 2} \log \frac{2h^2(k) + 4h(k) - 2(b^2 - 1)}{2h^2(k) + 4h(k) - 2(b^2 + b)} + 2g^{k-1} \\ &= \frac{1}{\log 2} \log (1 + O(h^{-2}(k))) + 2g^{k-1} = O(h^{-2}(k)) + 2g^{k-1}. \end{aligned}$$

Hence the probabilities of the inequalities in question form the terms of a convergent series, and the required result follows from the Borel-Cantelli lemma.

REFERENCES

1. H. Cramér, *Methods of mathematical statistics*, Princeton University Press, 1946.
2. J. L. Doob, *Stochastic processes*, New York, J. Wiley and Sons, 1953.
3. W. Feller, *Probability theory and its applications*, New York, J. Wiley and Sons, 1950.
4. G. H. Hardy, J. E. Littlewood and G. Pólya, *Inequalities*, Cambridge University Press, 1934.
5. W. Hoeffding and H. Robbins, *The central limit theorem for dependent random variables*, Trans. Amer. Math. Soc. vol. 15 (1948) pp. 773–780.
6. A. Ya. Khinchin, *Einige Sätze über Kettenbrüche, mit Anwendungen auf die Theorie der diophantischen Approximationen*, Math. Ann. vol. 92 (1924) pp. 115–125.
7. ———, *Zur metrischen Kettenbruchtheorie*, Comp. Math. vol. 3 (1936) pp. 276–285.
8. A. Kolmogorov, *Über das Gesetz des iterierten Logarithmus*, Math. Ann. vol. 101 (1929) pp. 126–135.
9. P. Lévy, *Sur le développement en fraction continue d'un nombre choisie au hasard*, Comp. Math. vol. 3 (1936) pp. 286–303.
10. ———, *Théorie de l'addition des variables aléatoires*, Paris, Gauthier-Villars, 1937.

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