BROWNIAN MOTION OF ROTATION

BY

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1. Brownian motion of rotation has been studied by F. Perrin [3] and K. Yosida [4]. However both studies are indirect and make use of the partial differential equation of diffusion. In this paper, by a direct application of the formulas for the composition of probabilities, we show how Brownian motion of rotation of the sphere may be constructed from the plane Brownian motion.

Let $(x(t), y(t))$, $0 \leq t < \infty$, be a sample path of the plane Brownian motion process. For each non-negative integer n let the interval $[0, 1]$ of real numbers be subdivided by the points $k/2^n$, $k=0, \dots, 2^n$. Let the polygonal path $\{(x, y) | x = x^{(n)}(t), y = y^{(n)}(t), 0 \leq t \leq 1\}$ be defined by:

$$x^{(n)}(t) = x\left(\frac{k-1}{2^n}\right) + 2^n\left(t - \frac{k-1}{2^n}\right)\left[x\left(\frac{k}{2^n}\right) - x\left(\frac{k-1}{2^n}\right)\right],$$

$$y^{(n)}(t) = y\left(\frac{k-1}{2^n}\right) + 2^n\left(t - \frac{k-1}{2^n}\right)\left[y\left(\frac{k}{2^n}\right) - y\left(\frac{k-1}{2^n}\right)\right],$$

for $(k-1)/2^n \leq t < k/2^n$, $k=1, \dots, 2^n$, $n=0, 1, \dots$.

Let S be a sphere of radius 1 on which has been chosen a point P and an oriented great circle through P . Let S be placed on the (x, y) -plane such that P coincides with the origin of coordinates and the great circle on S at P is tangent to the positive x -axis and is directed in the same sense. Assume now that S rolls without slipping along the polygonal path,

$$\{(x, y) | x = x^{(n)}(t), y = y^{(n)}(t), 0 \leq t \leq 1\}$$

in a way such that it has the same constant angular velocity along each of the linear portions of the path corresponding to the intervals of time $(k-1)/2^n \leq t < k/2^n$, $k=1, \dots, 2^n$. Let $R^{(n)}(t)$, $0 \leq t \leq 1$, be the rotation of S around its center as it rolls in the above manner from the point $(0, 0)$ to the point $(x^{(n)}(t), y^{(n)}(t))$. In this way, for each sample path of the plane Brownian motion process, a sequence $R^{(n)}(t)$, $0 \leq t \leq 1$, $n=0, 1, \dots$, of sample paths in the group of rotations in 3-space is obtained.

The main theorem states that $\lim_{n \rightarrow \infty} R^{(n)}(t)$ exists uniformly for t in

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the interval $[0, 1]$, for almost all sample paths of the plane Brownian motion process.

This paper is divided into sections which are numbered consecutively from 1 to 5. The first few sections contain preliminaries. §4 establishes the notation and contains a precise statement of the main theorem. In §5 the main theorem is proved.

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2. Let $(\Omega, \mathcal{E}, \text{Pr})$ be a probability space, i.e., $\Omega = \{\omega\}$ is a set of elements, $\mathcal{E} = \{E\}$ is a Borel field of subsets of called “events,” and Pr is a countably additive measure defined on $\mathcal{E}$ satisfying $\text{Pr}\{\Omega\} = 1$. $\text{Pr}\{E\}$ is called the “probability” of the event E . If $x = x(\omega)$ is a random variable, i.e., $x(\omega)$ is a real (complex) valued function defined on Ω such that $\{\omega | x(\omega) \in B\} \in \mathcal{E}$ for every Borel subset B of real (complex) numbers, then $E\{x\}$ shall be the “mean” or “expectation” of x , i.e.,

$$E\{x\} = \int_{\Omega} x(\omega) \text{Pr}(d\omega),$$

provided the integral exists.

A (mathematical) *one-dimensional Brownian motion* is a real-valued function $x(t, \omega)$ of the two variables t and ω , defined for all non-negative real numbers t , $0 \leq t < \infty$, and for all $\omega \in \Omega$, which has the following properties:

(B₁) $x(0, \omega) \equiv 0$,

(B₂) for any real numbers s, t with $0 \leq s < t < \infty$, the increment $x(t, \omega) - x(s, \omega)$ is $\mathcal{E}$ -measurable in ω and has a normal distribution with mean 0 and variance $t - s$, i.e.,

$$E(x, s, t, \alpha) = \{\omega | x(t, \omega) - x(s, \omega) < \alpha\} \in \mathcal{E}$$

and

$$\text{Pr}\{E(x, s, t, \alpha)\} = (1/(2\pi(t-s))^{1/2}) \int_{-\infty}^{\alpha} e^{-u^2/2(t-s)} du,$$

for any real number α .

(B₃) for any real numbers s_i, t_i , ($i=1, \dots, m$) with $0 \leq s_1 < t_1 \leq \dots \leq s_m < t_m < \infty$, the increments $x(t_i, \omega) - x(s_i, \omega)$, ($i=1, \dots, m$), are independent in the sense of probability theory, i.e.,

$$\text{Pr}\left\{\bigcap_{i=1}^m E(x, s_i, t_i, \alpha_i)\right\} = \prod_{i=1}^m \text{Pr}\{E(x, s_i, t_i, \alpha_i)\}$$

for any real numbers α_i ($i=1, \dots, m$).

A *two-dimensional* or *plane Brownian motion* is an ordered pair of two

mutually independent one-dimensional Brownian motions, i.e., a pair of one-dimensional Brownian motions $x(t, \omega)$ and $y(t, \omega)$ with the property that

$$\begin{aligned} & \Pr\{E(x, s, t, \alpha) \cap E(y, s', t', \alpha')\} \\ &= \Pr\{E(x, s, t, \alpha)\} \cdot \Pr\{E(y, s', t', \alpha')\} \end{aligned}$$

for any real numbers $s, t, \alpha, s', t', \alpha'$ with $0 \leq s < t < \infty$, $0 \leq s' < t' < \infty$.

If we consider $B(t, \omega) = (x(t, \omega), y(t, \omega))$ as a point in a Euclidean plane then for each fixed ω , $B(t, \omega)$ may be considered as a function of t for $0 \leq t < \infty$, which assumes as values points or vectors in the plane. For fixed ω the function $B(t, \omega)$ of t , $0 \leq t < \infty$, is called a *sample path*.

The following two lemmas from probability theory shall be needed in the last section of the paper.

LEMMA 1. *If $x_1, x_2, \dots$ is a sequence of non-negative real-valued random variables and if there exists a sequence $a_1, a_2, \dots$ of non-negative constants such that*

$$(i) \quad \sum_{n=1}^{\infty} a_n < \infty,$$

$$(ii) \quad \sum_{n=1}^{\infty} \Pr\{x_n > a_n\} < \infty,$$

then $\sum_{n=1}^{\infty} x_n < \infty$, with probability 1.

The proof of Lemma 1 may be derived from the lemmas of Borel and Cantelli.

LEMMA 2. *Let $x(t, \omega)$ be one-dimensional Brownian motion. Let the interval $[0, 1]$ be subdivided by the points $k/2^n$, $k=0, 1, \dots, 2^n$, $n=0, 1, \dots$. Let $x_k^{(n)} = x_k^{(n)}(\omega) = x(k/2^n, \omega) - x((k-1)/2^n, \omega)$. Then*

$$\sum_{n=0}^{\infty} \max_{1 \leq k \leq 2^n} |x_k^{(n)}(\omega)| < \infty$$

for an ω -set having probability 1.

The proof of Lemma 2 may be derived using Lemma 1 together with well-known properties of Gaussian variables.

3. We consider all rotations of 3-dimensional space around a fixed point (the origin of coordinates). Let g_1 and g_2 be two rotations. Then their product g is the rotation obtained by first carrying out the rotation g_2 and then the rotation g_1 . This is written in the following way: $g = g_1 \cdot g_2$. The set of all rotations form a group which shall be denoted by G .

Let a fixed rectangular coordinate system be chosen in 3-space. Then points or vectors x may be represented as triples of real numbers: (x_1, x_2, x_3) .

Let $|x|$ be the Euclidean length of the vector x , i.e., $|x| = (x_1^2 + x_2^2 + x_3^2)^{1/2}$. Three by three matrices A may be considered as linear transformations on this space of vectors. The notation $|A|$ shall be used to denote the operator norm of A , i.e., $|A| = \max_{|x|=1} |A(x)|$.

A rotation g of G may be represented analytically as a 3 by 3 orthogonal matrix V with determinant 1. The group G shall be topologized by defining the distance between two rotations g_1, g_2 to be $|V_1 - V_2|$, where V_1, V_2 are the orthogonal matrices corresponding to g_1, g_2 respectively.

4. Let Δ denote the interval $0 \leq a \leq t < b < \infty$ of non-negative real numbers t . The notation $x(\Delta, \omega) = x(b, \omega) - x(a, \omega)$ shall be used. Similarly $y(\Delta, \omega) = y(b, \omega) - y(a, \omega)$ and $B(\Delta, \omega) = B(b, \omega) - B(a, \omega) = (x(\Delta, \omega), y(\Delta, \omega))$. Let $\theta(\Delta, \omega)$ be the length of the vector $B(\Delta, \omega)$, i.e., $\theta(\Delta, \omega) = |B(\Delta, \omega)|$, where $|B| = |(x, y)| = (x^2 + y^2)^{1/2}$. Let $\phi(\Delta, \omega)$ denote oriented angle the vector $B(\Delta, \omega)$ makes with the positive y -axis. $\phi(\Delta, \omega)$ shall be measured so that the values it assumes lie between $-\pi$ and π .

For ϕ_1, θ, ϕ_2 arbitrary real numbers, define the 3 by 3 real orthogonal matrix $V(\phi_1, \theta, \phi_2)$ by

$$V(\phi_1, \theta, \phi_2) = \begin{pmatrix} \cos \phi_1 & -\sin \phi_1 & 0 \\ \sin \phi_1 & \cos \phi_1 & 0 \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} 1 & 0 & 0 \\ 0 & \cos \theta & -\sin \theta \\ 0 & \sin \theta & \cos \theta \end{pmatrix} \\ \times \begin{pmatrix} \cos \phi_2 & -\sin \phi_2 & 0 \\ \sin \phi_2 & \cos \phi_2 & 0 \\ 0 & 0 & 1 \end{pmatrix}.$$

If ϕ, θ are arbitrary real numbers let

$$(1) \quad V(\phi, \theta) = V(\phi, \theta, -\phi).$$

Let the interval $[0, 1]$ of real numbers t , $0 \leq t \leq 1$, be subdivided by the points $k/2^n, k=0, 1, \dots, 2^n$ for each non-negative integer $n=0, 1, \dots$. Let the subinterval $(k-1)/2^n \leq t < k/2^n, k=1, \dots, 2^n, n=0, 1, \dots$ be denoted by $\Delta_k^{(n)}$. For an arbitrary $t \in [0, 1]$, $[2^n t]$ shall denote the largest integer k which is not greater than $2^n t$. Let $V_k^{(n)} = V_k^{(n)}(\omega)$ denote the orthogonal matrix $V(\phi(\Delta_k^{(n)}, \omega), \theta(\Delta_k^{(n)}, \omega))$, $k=1, \dots, 2^n, n=0, 1, \dots$. Let t be an arbitrary real number from $[0, 1]$, then $t \in \Delta_k^{(n)}$ where $k = [2^n t] + 1$. The orthogonal matrix $V(\phi(\Delta_k^{(n)}, \omega), 2^n(t - (k-1)/2^n)\theta(\Delta_k^{(n)}, \omega))$ shall be denoted by $V_t^{(n)} = V_t^{(n)}(\omega), n=0, 1, \dots$. Often, when it seems that there is small probability of confusion, the probability variable ω shall be omitted from expressions such as $x(t, \omega), B(t, \omega), \theta(\Delta, \omega)$, etc. and these quantities shall be written more simply as $x(t), B(t), \theta(\Delta)$, etc.

Let t be an arbitrary real number from $[0, 1]$. Let

$$\begin{aligned}
 (2) \quad P^{(n)}(t) &= P^{(n)}(t, \omega) = V_t^{(n)}(\omega) \cdot V_{[2^n t]}^{(n)}(\omega) \cdots V_1^{(n)}(\omega) \\
 &= V_t^{(n)}(\omega) \cdot \prod_{k=1}^{[2^n t]} V_k^{(n)}(\omega),
 \end{aligned}$$

$n=0, 1, \dots$. It shall be convenient to introduce further notation. The rotation of 3-space whose matrix is $V_k^{(n)}(\omega)$ shall be denoted by $g_k^{(n)}(\omega)$, $k=1, \dots, 2^n$, $n=0, 1, \dots$. The rotation of 3-space whose matrix is $V_t^{(n)}(\omega)$ shall be denoted by $g_t^{(n)}(\omega)$, $t \in [0, 1]$, $n=0, 1, \dots$. Let

$$\begin{aligned}
 (3) \quad g^{(n)}(t) &= g^{(n)}(t, \omega) = g_t^{(n)} \cdot g_{[2^n t]}^{(n)} \cdots g_1^{(n)} \\
 &= g_t^{(n)}(\omega) \prod_{k=1}^{[2^n t]} g_k^{(n)}(\omega).
 \end{aligned}$$

THEOREM. *The limit,*

$$g(t, \omega) = \lim_{n \rightarrow \infty} g^{(n)}(t, \omega)$$

exists uniformly in t , $0 \leq t \leq 1$, for $\omega(\in \Omega)$ belonging to an event which has probability 1.

We remark that it follows from the main theorem that $g(t, \omega)$ is a continuous stochastic process in the rotation group G . Furthermore, it can be shown that the characteristic matrices [2] $E\{T^l(g(t))\}$ of this process are given by

$$E\{T^l(g(t))\} = \text{diag}\{\dots e^{(l+l^2-p^2)/2 \cdot t} \dots\}$$

$p = -l, -l+1, \dots, l$, $l=0, 1/2, 1, 3/2, \dots$. Here [cf. 1] $T^l(g) = (T_{pq}^l(g))$, $l=0, 1/2, 1, 3/2, \dots$ is a complete system of mutually inequivalent continuous irreducible representations of G by unitary matrices.

5. In this section a proof of the main theorem shall be given. If $A = (a_{ij})$ is a 3 by 3 matrix of real numbers, then $\text{tr}(A)$ denotes the *trace* of A , i.e., $\text{tr}(A) = \sum_i a_{ii}$. A^* denotes the transposed matrix of A , i.e., the matrix obtained from A by interchanging its rows and columns. Further let

$$\|A\| = (\text{tr}(A \cdot A^*))^{1/2} = \left(\sum_{i,j} a_{ij}^2 \right)^{1/2}.$$

If $|A|$ is the norm of A when it is considered as an operator on the 3-dimensional Euclidean vector space, then

$$|A| \leq \|A\|.$$

Let $V(\phi, \theta)$ be the real orthogonal matrix defined by (1), then

$$(4) \quad V(\phi, \theta) = I + A(\phi, \theta),$$

for arbitrary ϕ, θ , where I =identity matrix and

$$(5) \quad A(\phi, \theta) = \begin{pmatrix} -(1 - \cos \theta) \sin^2 \phi & (1 - \cos \theta) \sin \phi \cos \phi & \sin \theta \sin \phi \\ (1 - \cos \theta) \sin \phi \cos \phi & -(1 - \cos \theta) \cos^2 \phi & -\sin \theta \cos \phi \\ -\sin \theta \sin \phi & \sin \theta \cos \phi & -(1 - \cos \theta) \end{pmatrix}.$$

Let the functions $C=C(\theta)$ and $S=S(\theta)$ be defined by the equations

$$(6) \quad 1 - \cos \theta = \frac{\theta^2}{2} - \theta^4 \cdot C(\theta),$$

$$(7) \quad \sin \theta = \theta - \theta^3 \cdot S(\theta).$$

Then C and S are bounded functions of θ , $-\infty < \theta < \infty$.

The proof of the following lemma may be demonstrated using (5).

LEMMA 3. *Let ϕ and θ be arbitrary real numbers, then*

$$\|A(\phi, \theta)\| \leq |\theta| \cdot K(|\theta|),$$

where K is a polynomial which has real non-negative coefficients.

Let the change of variables: $(\phi, \theta) \rightarrow (x, y)$ be defined by

$$(8) \quad \begin{aligned} x &= \theta \sin \phi, \\ y &= \theta \cos \phi. \end{aligned}$$

Under this change of variables the matrix $A(\phi, \theta)$ is transformed into a matrix which shall be denoted by $A(x, y)$. Using (6), (7) and (8) one may write

$$(9) \quad A(x, y) = A_1(x, y) + A_3(x, y)$$

where

$$(10) \quad A_1(x, y) = \begin{pmatrix} -x^2/2 & xy/2 & x \\ xy/2 & -y^2/2 & -y \\ -x & y & -(x^2 + y^2)/2 \end{pmatrix},$$

and

$$(11) \quad A_3(x, y) = \begin{pmatrix} (x^2 + y^2)x^2 \cdot C & -(x^2 + y^2)xy \cdot C & -(x^2 + y^2)xS \\ -(x^2 + y^2)xy \cdot C & (x^2 + y^2)y^2 \cdot C & (x^2 + y^2)y \cdot S \\ (x^2 + y^2)x \cdot S & -(x^2 + y^2)y \cdot S & (x^2 + y^2) \cdot C \end{pmatrix}$$

where

$$C = C((x^2 + y^2)^{1/2}), \quad S = S((x^2 + y^2)^{1/2}), \quad -\infty < x, y < \infty.$$

It shall be convenient to introduce the following notational convention: $B_3 = B_3(x, y, x', y', \dots; C, S, C', S', \dots)$ denotes a 3 by 3 matrix whose elements are polynomials in the variables $x, y, x', y', \dots, C, S, C', S', \dots$ any term of which is at least of third degree in the variables $x, y, x', y', \dots$. Here $C = C((x^2 + y^2)^{1/2})$, $S = S((x^2 + y^2)^{1/2})$, $C' = C((x'^2 + y'^2)^{1/2})$, $S' = S((x'^2 + y'^2)^{1/2})$, $\dots$, $-\infty < x, y, x', y', \dots < \infty$.

Let $V(x, y)$ denote the matrix into which $V(\phi, \theta)$ is transformed by the change of variables (8).

LEMMA 4. Let $(x, y), (x', y')$ be pairs of real numbers, then

$$(12) \quad I - V^*(x + x', y + y') \cdot V(x, y) \cdot V(x', y') \\ = W(x, y, x', y') + B_3(x, y, x', y'; C, S, C', S'),$$

where

$$(13) \quad W(x, y, x', y') = \begin{pmatrix} 0 & (yx' - xy')/2 & 0 \\ (xy' - yx')/2 & 0 & 0 \\ 0 & 0 & 0 \end{pmatrix}.$$

Proof. The left-hand side of (12) may be written

$$\begin{aligned} I - [I + A^*(x + x', y + y')][I + A(x, y)][I + A(x', y')] \\ = -A^*(x + x', y + y') - A(x, y) - A(x', y') \\ - A^*(x + x', y + y')A(x, y) - A^*(x + x', y + y')A(x', y') \\ - A(x, y)A(x', y') + B_3 \\ = -A_1^*(x + x', y + y') - A_1(x, y) - A_1(x', y') \\ - A_1^*(x + x', y + y')A_1(x, y) - A_1^*(x + x', y + y')A_1(x', y') \\ - A_1(x, y)A_1(x', y') + B_3. \end{aligned}$$

$$A_1^*(x, y) \cdot A_1(x', y') = \begin{pmatrix} xx' & xy' & 0 \\ -yx' & -yy' & 0 \\ 0 & 0 & xx' + yy' \end{pmatrix} + B_3.$$

$$A_1(x, y) \cdot A_1(x', y') = \begin{pmatrix} -xx' & -xy' & 0 \\ xy' & yy' & 0 \\ 0 & 0 & -xx' - yy' \end{pmatrix} + B_3.$$

The left-hand side of (12) may now be written:

$$\begin{aligned}
& \begin{pmatrix} (x+x')^2/2 & -(x+x')(y+y')/2 & (x+x') \\ -(x+x')(y+y') & (y+y')^2/2 & -(y+y') \\ -(x+x') & (y+y') & (x+x')^2/2 + (y+y')^2/2 \end{pmatrix} \\
& + \begin{pmatrix} x^2/2 & -xy/2 & -x \\ -xy/2 & y^2/2 & y \\ x & -y & (x^2+y^2)/2 \end{pmatrix} + \begin{pmatrix} x'^2/2 & -x'y'/2 & -x' \\ -x'y'/2 & y'^2/2 & y' \\ x' & -y' & (x'^2+y'^2)/2 \end{pmatrix} \\
& + \begin{pmatrix} -(x+x')x & (x+x')y & 0 \\ (y+y') \cdot x & -(y+y') \cdot y & 0 \\ 0 & 0 & -(x+x')x - (y+y') \cdot y \end{pmatrix} \\
& + \begin{pmatrix} -(x+x') \cdot x' & (x+x') \cdot y' & 0 \\ (y+y') \cdot x' & -(y+y') \cdot y' & 0 \\ 0 & 0 & -(x+x')x' - (y+y') \cdot y' \end{pmatrix} \\
& + \begin{pmatrix} xx' & -xy' & 0 \\ -yx' & yy' & 0 \\ 0 & 0 & xx' + yy' \end{pmatrix} + B_3 \\
& = \begin{pmatrix} 0 & (yx' - xy')/2 & 0 \\ (xy' - yx')/2 & 0 & 0 \\ 0 & 0 & 0 \end{pmatrix} + B_3 \\
& = W(x, y, x', y') + B_3.
\end{aligned}$$

This completes the proof of Lemma 4.

LEMMA 5. Let

$$x = x(\Delta_{2k-1}^{(n+1)}, \omega), \quad y = y(\Delta_{2k+1}^{(n+1)}, \omega), \quad x' = x(\Delta_{2k}^{(n+1)}, \omega), \quad y' = y(\Delta_{2k+1}^{(n+1)}, \omega)$$

for some k , $1 \leq k \leq 2^n$, and $n=0, 1, \dots$, then

$$(2^n)^{3/2} \|E\{B_3(x, y, x', y'; C, S, C', S')\}\|$$

is a bounded function of $n=0, 1, \dots$.

Proof. Let $B_3 = (B_{ij})$, $i, j=1, 2, 3$. Then

$$\begin{aligned}
(2^n)^{3/2} \|E\{B_3\}\| &= (2^n)^{3/2} \left(\sum_{i,j=1}^3 (E\{B_{ij}\})^2 \right)^{1/2} \\
&\leq (2^n)^{3/2} \sum_{i,j=1}^3 |E\{B_{ij}\}|.
\end{aligned}$$

Thus it is sufficient to show that for $i, j = 1, 2, 3$, there are non-negative constants M_{ij} such that

$$(2^n)^{3/2} |E\{B_{ij}\}| \leq M_{ij} \text{ (independently of } n\text{)}.$$

In order to prove this, recall that B_{ij} is a polynomial in the variables $x, y, x', y', C, S, C', S'$ where $C = C((x^2 + y^2)^{1/2})$, $S = S((x^2 + y^2)^{1/2})$, $S' = S'((x'^2 + y'^2)^{1/2})$: $B_{ij} = \sum_{\alpha, \dots, \delta'} A_{\alpha, \dots, \delta'} x^\alpha y^\beta x'^{\alpha'} y'^{\beta'} C^\gamma S^\delta C'^{\gamma'} S'^{\delta'}$ where $\alpha, \dots, \delta'$ take non-negative integral values, $A_{\alpha, \dots, \delta'}$ are real numbers and $\alpha + \beta + \alpha' + \beta' \geq 3$.

$$|E\{B_{ij}\}| \leq \sum_{\alpha, \dots, \delta'} M_{\alpha, \dots, \delta'} E\{|x|^\alpha \cdot |y|^\beta \cdot |x'|^{\alpha'} \cdot |y'|^{\beta'}\}$$

where $M_{\alpha, \dots, \delta'} = |A_{\alpha, \dots, \delta'}| \cdot \bar{C}^{\gamma+\gamma'} \bar{S}^{\delta+\delta'}$, here $|C(\theta)| \leq \bar{C}$, $|S(\theta)| \leq \bar{S}$ for all θ , $-\infty < \theta < \infty$.

$$\begin{aligned} E\{|x|^\alpha \cdot |y|^\beta \cdot |x'|^{\alpha'} \cdot |y'|^{\beta'}\} &= E\{|x|^\alpha\} \cdot E\{|y|^\beta\} \cdot E\{|x'|^{\alpha'}\} \cdot E\{|y'|^{\beta'}\} \\ &\leq (E\{|x|^{2\alpha}\})^{1/2} (E\{|y|^{2\beta}\})^{1/2} (E\{|x'|^{2\alpha'}\})^{1/2} (E\{|y'|^{2\beta'}\})^{1/2} \\ &= \frac{\nu(\alpha)}{(2^{n+1})^{\alpha/2}} \cdot \frac{\nu(\beta)}{(2^{n+1})^{\beta/2}} \cdot \frac{\nu(\alpha')}{(2^{n+1})^{\alpha'/2}} \cdot \frac{\nu(\beta')}{(2^{n+1})^{\beta'/2}}. \end{aligned}$$

Here $\nu(\xi)$ denotes $(2\xi-1)(2\xi-3) \cdots 1$. Thus

$$\begin{aligned} (2^n)^{3/2} |E\{B_{ij}\}| &\leq \sum_{\alpha, \dots, \delta'} M_{\alpha, \dots, \delta'} K_{\alpha, \dots, \delta'} \left(\frac{1}{2^n}\right)^{(\alpha+\beta+\alpha'+\beta'-3)/2} \\ &\leq \sum_{\alpha, \dots, \delta'} M_{\alpha, \dots, \delta'} K_{\alpha, \dots, \delta'} \\ &= M_{ij}, \end{aligned}$$

since $\alpha + \beta + \alpha' + \beta' \geq 3$. This completes the proof of Lemma 5.

Proof of the main theorem. In order to show that

$$\lim_{n \rightarrow \infty} P^{(n)}(t, \omega)$$

exists uniformly for $t \in [0, 1]$, for an ω -set having probability 1, it is sufficient to prove that the series

$$(14) \quad \sum_{n=0}^{\infty} \max_{0 \leq t \leq 1} |P^{(n)}(t, \omega) - P^{(n+1)}(t, \omega)|$$

converges with probability 1.

Let $n = 0, 1, \dots$ be fixed, let $t \in [0, 1]$ and $k = [2^n t] + 1$. Then either $t \in \Delta_{2k-1}^{(n+1)}$ or $t \in \Delta_{2k+1}^{(n+1)}$, recalling the definition (2),

$$\begin{aligned}
& |P^{(n)}(t) - P^{(n+1)}(t)| \\
&= \left| V_t^{(n)} P^{(n)}\left(\frac{k-1}{2^n}\right) - V_t^{(n+1)} P^{(n+1)}\left(\frac{2k-2}{2^{n+1}}\right) \right| \\
&= \left| V_t^{(n)} P^{(n)}\left(\frac{k-1}{2^n}\right) - V_t^{(n+1)} P^{(n)}\left(\frac{k-1}{2^n}\right) \right. \\
&\quad \left. + V_t^{(n+1)} P^{(n)}\left(\frac{k-1}{2^n}\right) - V_t^{(n+1)} P^{(n+1)}\left(\frac{2k-2}{2^{n+1}}\right) \right| \\
&\leq \left| V_t^{(n)} - V_t^{(n+1)} \right| \cdot \left| P^{(n)}\left(\frac{k-1}{2^n}\right) \right| \\
&\quad + \left| V_t^{(n+1)} \right| \cdot \left| P^{(n)}\left(\frac{k-1}{2^n}\right) - P^{(n+1)}\left(\frac{2k-2}{2^{n+1}}\right) \right|.
\end{aligned}$$

Since $V_t^{(n+1)}$ and $P^{(n)}(k-1)/2^n$ are orthogonal

$$\begin{aligned}
(15) \quad & |P^{(n)}(t) - P^{(n+1)}(t)| \leq |V_t^{(n)} - V_t^{(n+1)}| \\
& + \left| P^{(n)}\left(\frac{k-1}{2^n}\right) - P^{(n+1)}\left(\frac{2k-2}{2^{n+1}}\right) \right|, \quad t \in \Delta_{2k-1}^{(n+1)}.
\end{aligned}$$

Similarly if $t \in \Delta_{2k}^{(n+1)}$, then

$$\begin{aligned}
(15') \quad & |P^{(n)}(t) - P^{(n+1)}(t)| \leq |V_t^{(n)} - V_t^{(n+1)} V_{2k-1}^{(n+1)}| \\
& + \left| P^{(n)}\left(\frac{k-1}{2^n}\right) - P^{(n+1)}\left(\frac{2k-2}{2^{n+1}}\right) \right|.
\end{aligned}$$

Using (15) and (15') one obtains

$$\begin{aligned}
& \max_{0 \leq t \leq 1} |P^{(n)}(t) - P^{(n+1)}(t)| \\
&= \max_{1 \leq k \leq 2^n} \cdot \max_{t \in \Delta_k^{(n)}} |P^{(n)}(t) - P^{(n+1)}(t)| \\
&\leq \max_{1 \leq k \leq 2^n} \cdot \max_{t \in \Delta_{2k-1}^{(n+1)}} |P^{(n)}(t) - P^{(n+1)}(t)| \\
&\quad + \max_{1 \leq k \leq 2^n} \cdot \max_{t \in \Delta_{2k}^{(n+1)}} |P^{(n)}(t) - P^{(n+1)}(t)| \\
&\leq 2 \cdot \max_{1 \leq k \leq 2^n} \left| P^{(n)}\left(\frac{k}{2^n}\right) - P^{(n+1)}\left(\frac{2k}{2^{n+1}}\right) \right| \\
&\quad + \max_{1 \leq k \leq 2^n} \cdot \max_{t \in \Delta_{2k-1}^{(n+1)}} |V_t^{(n)} - V_t^{(n+1)}| \\
&\quad + \max_{1 \leq k \leq 2^n} \cdot \max_{t \in \Delta_{2k}^{(n+1)}} |V_t^{(n)} - V_t^{(n+1)} \cdot V_{2k-1}^{(n+1)}|.
\end{aligned}$$

Introduce the notation $P_k^{(n)} = P^{(n)}(k/2^n)$, $k = 1, \dots, 2^n$, $n = 0, 1, \dots$. In order to prove that (14) converges with probability 1 it is thus sufficient to prove that

$$(16) \quad \sum_{n=0}^{\infty} \max_{1 \leq k \leq 2^n} \cdot \max_{t \in \Delta_{2k-1}^{(n+1)}} |V_t^{(n)}(\omega) - V_t^{(n+1)}(\omega)|,$$

$$(17) \quad \sum_{n=0}^{\infty} \max_{1 \leq k \leq 2^n} \cdot \max_{t \in \Delta_{2k}^{(n+1)}} |V_t^{(n)}(\omega) - V_t^{(n+1)}(\omega) \cdot V_{2k-1}^{(n+1)}(\omega)|$$

and

$$(18) \quad \sum_{n=0}^{\infty} \max_{1 \leq k \leq 2^n} |P_k^{(n)}(\omega) - P_{2k}^{(n+1)}(\omega)|,$$

all converge with probability 1.

The convergence of (16) with probability 1 shall be demonstrated first. Let $t \in \Delta_{2k-1}^{(n+1)}$, then

$$V_t^{(n)} = V(\phi(\Delta_k^{(n)}), \theta_t(\Delta_k^{(n)})) = I + A(\phi(\Delta_k^{(n)}), \theta_t(\Delta_k^{(n)}))$$

and

$$\begin{aligned} V_t^{(n+1)} &= V(\phi(\Delta_{2k-1}^{(n+1)}), \theta_t(\Delta_{2k-1}^{(n+1)})) \\ &= I + A(\phi(\Delta_{2k-1}^{(n+1)}), \theta_t(\Delta_{2k-1}^{(n+1)})). \end{aligned}$$

Here $\theta_t(\Delta_k^{(n)})$ is an abbreviation of $2^n(t - (k-1)/2^n)\theta(\Delta_k^{(n)})$ and $\theta_t(\Delta_{2k-1}^{(n+1)})$ is an abbreviation of $2^{n+1}(t - (2k-2)/2^{n+1})\theta(\Delta_{2k-1}^{(n+1)})$.

$$\begin{aligned} |V_t^{(n)} - V_t^{(n+1)}| &\leq |A(\phi(\Delta_k^{(n)}), \theta_t(\Delta_k^{(n)}))| + |A(\phi(\Delta_{2k-1}^{(n+1)}), \theta_t(\Delta_{2k-1}^{(n+1)}))| \\ &\leq \theta_t(\Delta_k^{(n)})K(\theta_t(\Delta_k^{(n)})) + \theta_t(\Delta_{2k-1}^{(n+1)})K(\theta_t(\Delta_{2k-1}^{(n+1)})), \end{aligned}$$

by Lemma 3. Here K is a polynomial which has real non-negative coefficients. From the above inequality one obtains

$$\begin{aligned} &\max_{1 \leq k \leq 2^n} \max_{t \in \Delta_{2k-1}^{(n+1)}} |V_t^{(n)} - V_t^{(n+1)}| \\ &\leq \max_{1 \leq k \leq 2^n} [\theta(\Delta_k^{(n)})K(\theta(\Delta_k^{(n)})) + \theta(\Delta_{2k-1}^{(n+1)})K(\theta(\Delta_{2k-1}^{(n+1)}))] \\ &\leq \max_{1 \leq k \leq 2^n} \theta(\Delta_k^{(n)}) \cdot K\left(\max_{1 \leq k \leq 2^n} \theta(\Delta_k^{(n)})\right) \\ &\quad + \max_{1 \leq k \leq 2^n} \theta(\Delta_{2k-1}^{(n+1)}) \cdot K\left(\max_{1 \leq k \leq 2^n} \theta(\Delta_{2k-1}^{(n+1)})\right). \end{aligned}$$

Since $\theta(\Delta_k^{(n)}) \leq |x(\Delta_k^{(n)})| + |y(\Delta_k^{(n)})|$, by Lemma 2

$$\sum_{n=0}^{\infty} \max_{1 \leq k \leq 2^n} \theta(\Delta_k^{(n)}) < \infty \text{ with probability 1.}$$

Similarly

$$\sum_{n=0}^{\infty} \max_{1 \leq k \leq 2^n} \theta(\Delta_{2k-1}^{(n+1)}) < \infty, \text{ with probability 1.}$$

Let $\max_{1 \leq k \leq 2^n} \theta(\Delta_k^{(n)}) = z_n$, $\sum_0^\infty z_n = z < \infty$ with probability 1. Then $\sum_0^\infty z_n K(z_n) \leq \sum_0^\infty z_n K(z) = zK(z)$, with probability 1. It follows that (15) converges with probability 1. The demonstration that (17) converges with probability 1 is not essentially different from the proof that (15) converges with probability 1. It will be omitted.

Consider the series (18). In order to show that this series converges with probability 1, use shall be made of Lemma 1. Thus the probability

$$(19) \quad \Pr \left\{ \max_{1 \leq k \leq 2^n} |P_k^{(n)} - P_{2k}^{(n+1)}| > a \right\}$$

must be estimated. Here a is a positive constant.

Let $Q_k^{(n)} = \|P_k^{(n)} - P_{2k}^{(n+1)}\|^2$, then

$$\begin{aligned} & \Pr \left\{ \max_{1 \leq k \leq 2^n} |P_k^{(n)} - P_{2k}^{(n+1)}| > a \right\} \\ & \leq \Pr \left\{ \max_{1 \leq k \leq 2^n} \|P_k^{(n)} - P_{2k}^{(n+1)}\|^2 > a^2 \right\} \\ & = \Pr \left\{ \max_{1 \leq k \leq 2^n} Q_k^{(n)} > a^2 \right\}. \end{aligned}$$

Introduce the ω -sets $E_k = E_k^{(n)}$ defined by

$$E_k = \{Q_1^{(n)} \leq a^2, \dots, Q_{k-1}^{(n)} \leq a^2, Q_k^{(n)} > a^2\},$$

$k=1, \dots, 2^n, n=0, 1, \dots$. Then

$$(20) \quad \Pr \left\{ \max_{1 \leq k \leq 2^n} Q_k^{(n)} > a^2 \right\} = \sum_{k=1}^{2^n} \Pr\{E_k\}.$$

Let the real valued random variables $x_k = x_k^{(n)}(\omega)$ be defined by

$$x_k = \begin{cases} 1, & \omega \in E_k \\ 0, & \omega \in \Omega - E_k, \end{cases} \quad k = 1, \dots, 2^n.$$

Let $\sigma_k = x_1 + \dots + x_k$, $k=1, \dots, 2^n$. Let $E_0 = \emptyset$, $\sigma_0 \equiv 0$, then

$$\begin{aligned}
E\{Q_{2^n}^{(n)}\} &\geq E\{\sigma_{2^n} Q_{2^n}^{(n)}\} \\
&= \sum_{k=1}^{2^n} E\{\sigma_k Q_k^{(n)} - \sigma_{k-1} Q_{k-1}^{(n)}\} \\
&= \sum_{k=1}^{2^n} E\{x_k Q_k^{(n)} + \sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\} \\
&= \sum_{k=1}^{2^n} E\{x_k \cdot Q_k^{(n)}\} + \sum_{k=1}^{2^n} E\{\sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\} \\
&\geq a^2 \sum_{k=1}^{2^n} \Pr\{E_k\} + \sum_{k=1}^{2^n} E\{\sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\}.
\end{aligned}$$

Combining this inequality with (20) one obtains

$$(21) \quad \Pr\left\{\max_{1 \leq k \leq 2^n} Q_k^{(n)} > a^2\right\} \leq \frac{1}{a^2} [E\{Q_{2^n}^{(n)}\} - \sum_{k=1}^{2^n} E\{\sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\}].$$

$$\begin{aligned}
&E\{\sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\} \\
&= E\{\sigma_{k-1}(\|P_k^{(n)} - P_{2k}^{(n+1)}\|^2 - \|P_{k-1}^{(n)} - P_{2k-2}^{(n+1)}\|^2)\} \\
&= E\{\sigma_{k-1} \operatorname{tr}[(P_k^{(n)} - P_{2k}^{(n+1)}) \cdot (P_k^{(n)*} - P_{2k}^{(n+1)*}) \\
&\quad - (P_{k-1}^{(n)} - P_{2k-2}^{(n+1)}) \cdot (P_{k-1}^{(n)*} - P_{2k-2}^{(n+1)*})]\} \\
&= 2E\{\sigma_{k-1} \cdot \operatorname{tr}[P_{2k-2}^{(n+1)} P_{k-1}^{(n)*} - P_{2k}^{(n+1)} P_k^{(n)*}]\}.
\end{aligned}$$

By definition of $P_k^{(n)}$ and $P_{2k}^{(n+1)}$:

$$P_k^{(n)} = V_k^{(n)} P_{k-1}^{(n)}, \quad P_{2k}^{(n+1)} = V_{2k}^{(n+1)} P_{2k-1}^{(n+1)} P_{2k-2}^{(n+1)}.$$

Thus

$$\begin{aligned}
&\operatorname{tr}[P_{2k-2}^{(n+1)} P_{k-1}^{(n)*} - P_{2k}^{(n+1)} P_k^{(n)*}] \\
&= \operatorname{tr}(P_{2k-2}^{(n+1)} P_{k-1}^{(n)*}) - \operatorname{tr}(V_{2k}^{(n+1)} V_{2k-1}^{(n+1)} P_{2k-2}^{(n+1)} P_{k-1}^{(n)*} V_k^{(n)*}) \\
&= \operatorname{tr}(P_{2k-2}^{(n+1)} P_{k-1}^{(n)*}) - \operatorname{tr}(V_k^{(n)*} V_{2k}^{(n+1)} V_{2k-1}^{(n+1)} P_{2k-2}^{(n+1)} P_{k-1}^{(n)*}) \\
&= \operatorname{tr}[(I - V_k^{(n)*} V_{2k}^{(n+1)} V_{2k-1}^{(n+1)}) \cdot (P_{2k-2}^{(n+1)} P_{k-1}^{(n)*})]. \\
&2E\{\sigma_{k-1} \operatorname{tr}[P_{2k-2}^{(n+1)} P_{k-1}^{(n)*} - P_{2k}^{(n+1)} P_k^{(n)*}]\} \\
&= 2 \operatorname{tr}[E\{(I - V_k^{(n)*} V_{2k}^{(n+1)} V_{2k-1}^{(n+1)}) \cdot (\sigma_{k-1} P_{2k-2}^{(n+1)} P_{k-1}^{(n)*})\}] \\
&= 2 \operatorname{tr}[E\{(I - V_k^{(n)*} V_{2k}^{(n+1)} V_{2k-1}^{(n+1)})\} \cdot E\{\sigma_{k-1} P_{2k-2}^{(n+1)} P_{k-1}^{(n)*}\}].
\end{aligned}$$

Since $P_{2k-2}^{(n+1)} P_{k-1}^{(n)*}$ is orthogonal and σ_{k-1} may take only the values 0 and 1, any element of the matrix $E\{\sigma_{k-1} P_{2k-2}^{(n+1)} P_{k-1}^{(n)*}\}$ is in absolute value no greater than 1. Thus:

$$\begin{aligned}
& |E\{\sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\}| \\
&= |2 \cdot \text{tr}[E\{I - V_k^{(n)*} V_{2k}^{(n+1)} \cdot V_{2k-1}^{(n+1)}\} \cdot E\{\sigma_{k-1} P_{2k-2}^{(n+1)} P_{k-1}^{(n)*}\}]]| \\
&\leq 2 \cdot 3 \cdot \|E\{I - V_k^{(n)*} V_{2k}^{(n+1)} V_{2k-1}^{(n+1)}\}\|.
\end{aligned}$$

Hence

$$(22) \quad |E\{\sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\}| \leq 6 \cdot \|E\{I - V_k^{(n)*} \cdot V_{2k}^{(n+1)} V_{2k-1}^{(n+1)}\}\|.$$

Similarly one may compute

$$(23) \quad |E\{Q_k^{(n)} - Q_{k-1}^{(n)}\}| \leq 6 \cdot \|E\{I - V_k^{(n)*} \cdot V_{2k}^{(n+1)} \cdot V_{2k-1}^{(n+1)}\}\|.$$

Turning to the evaluation of

$$\begin{aligned}
& E\{I - V_k^{(n)*} V_{2k}^{(n+1)} V_{2k-1}^{(n+1)}\}: \\
V_k^{(n)*} &= V^*(\phi(\Delta_k^{(n)}), \theta(\Delta_k^{(n)})) = V^*(x(\Delta_k^{(n)}), y(\Delta_k^{(n)})), \\
V_{2k}^{(n+1)} &= V(\phi(\Delta_{2k}^{(n+1)}), \theta(\Delta_{2k}^{(n+1)})) = V(x(\Delta_{2k}^{(n+1)}), y(\Delta_{2k}^{(n+1)})), \\
V_{2k-1}^{(n+1)} &= V(\phi(\Delta_{2k-1}^{(n+1)}), \theta(\Delta_{2k-1}^{(n+1)})) = V(x(\Delta_{2k-1}^{(n+1)}), y(\Delta_{2k-1}^{(n+1)})).
\end{aligned}$$

Let $x = x(\Delta_{2k}^{(n+1)})$, $y = y(\Delta_{2k}^{(n+1)})$, $x' = x(\Delta_{2k-1}^{(n+1)})$, $y' = y(\Delta_{2k-1}^{(n+1)})$, then $x + x' = x(\Delta_k^{(n)})$, $y + y' = y(\Delta_k^{(n)})$. By Lemma 4:

$$\begin{aligned}
& \|E\{I - V_k^{(n)*} V_{2k}^{(n+1)} V_{2k-1}^{(n+1)}\}\| \\
&= \|E\{I - V^*(x + x', y + y') V(x, y) V(x', y')\}\| \\
&= \|E\{W(x, y, x', y')\} + E\{B_3\}\| \\
&= \|E\{B_3\}\|,
\end{aligned}$$

$E\{W\}$ being the zero matrix because the gaussian variables x, y, x', y' are independent and have mean zero. Now

$$\|E\{B_3\}\| = \left(\frac{1}{2^n}\right)^{3/2} b_n, \quad n = 0, 1, \dots,$$

where by Lemma 5, b_n is a bounded sequence of non-negative real numbers. Thus

$$\begin{aligned}
|E\{\sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\}| &\leq \left(\frac{1}{2^n}\right)^{3/2} \cdot 6b_n, \\
|E\{Q_k^{(n)} - Q_{k-1}^{(n)}\}| &\leq \left(\frac{1}{2^n}\right)^{3/2} \cdot 6 \cdot b_n,
\end{aligned}$$

$k = 1, \dots, 2^n$, $n = 0, 1, \dots$. From these inequalities one obtains:

$$\begin{aligned}
 E\{Q_{2^n}^{(n)}\} &= E\left\{\sum_{k=1}^{2^n} (Q_k^{(n)} - Q_{k-1}^{(n)})\right\} \\
 (24) \qquad &\leq \sum_{k=1}^{2^n} |E\{Q_k^{(n)} - Q_{k-1}^{(n)}\}| \leq \left(\frac{1}{2^n}\right)^{1/2} \cdot 6b_n.
 \end{aligned}$$

Similarly

$$(25) \qquad \left| \sum_{k=1}^{2^n} E\{\sigma_{k-1}(Q_k^{(n)} - Q_{k-1}^{(n)})\} \right| \leq \left(\frac{1}{2^n}\right)^{1/2} \cdot 6 \cdot b_n.$$

Let the estimates (23) and (24) be applied to (21). One obtains

$$\Pr\left\{\max_{1 \leq k \leq 2^n} Q_k^{(n)} > a^2\right\} \leq \frac{1}{a^2} \left(\frac{1}{2^n}\right)^{1/2} \cdot (12b_n).$$

Finally

$$(26) \qquad \Pr\left\{\max_{1 \leq k \leq 2^n} |P_k^{(n)} - P_{2k}^{(n+1)}| > a\right\} \leq \frac{1}{a^2} \cdot \left(\frac{1}{2^n}\right)^{1/2} \cdot c_n$$

$n=0, 1, 2, \dots$ where c_n is a bounded sequence of non-negative real numbers. In (26) let $a=a_n=2^{-n/8}$ then $\sum_{n=0}^{\infty} a_n < \infty$ and

$$\Pr\left\{\max_{1 \leq k \leq 2^n} |P_k^{(n)} - P_{2k}^{(n+1)}| > a_n\right\} \leq \left(\frac{1}{2^n}\right)^{1/4} \cdot c_n$$

so that $\sum_{n=0}^{\infty} \Pr\{\max_{1 \leq k \leq 2^n} |P_k^{(n)} - P_{2k}^{(n+1)}| > a_n\} < \infty$. From Lemma 1, one may now conclude that

$$\sum_{n=0}^{\infty} \max_{1 \leq k \leq 2^n} |P_k^{(n)}(\omega) - P_{2k}^{(n+1)}(\omega)| < \infty,$$

for an ω -set of probability 1. This proves that (18) converges with probability 1 and completes the proof of the main Theorem.

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